

Generation of Non-Uniform Random Numbers

Refs: Chapter 8 in Law and book by Devroye (watch for typos)

Peter J. Haas

CS 590M: Simulation
Spring Semester 2020

Generation of Non-Uniform Random Numbers

- Acceptance-Rejection
- Convolution Method
- Composition Method
- Alias Method
- Random Permutations and Samples
- Non-Homogeneous Poisson Processes

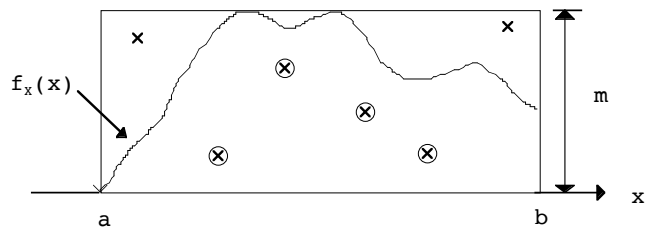
Acceptance-Rejection

Goal: Generate a random variate X having pdf f_X

- ▶ Avoids computation of $F^{-1}(u)$ as in inversion method
- ▶ Assumes f_X is easy to calculate

Special case: $f_X(x) > 0$ only on $[a, b]$ (finite support)

- ▶ Throw down points uniformly in enclosing rectangle R , reject points above f_X curve



- ▶ Return x -coordinate of accepted point

Acceptance-Rejection, Continued

Claim:

x -coordinate of an accepted point has pdf f_X

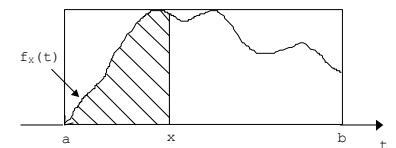
Proof

1. Let (Z_1, Z_2) be the (x, y) -coordinates of a random point distributed uniformly in R and fix $x \in [a, b]$
2. Then $P(Z_1 \leq x, \text{acceptance}) = P(Z_1 \leq x, Z_2 \leq f_X(Z_1))$
3. But $P(Z_1 \leq x, Z_2 \leq f_X(Z_1)) = \text{prob that } (Z_1, Z_2) \text{ falls in shaded region:}$

$$P(Z_1 \leq x, \text{acceptance}) =$$

$$P(\text{acceptance}) =$$

$$P(Z_1 \leq x \mid \text{acceptance}) =$$



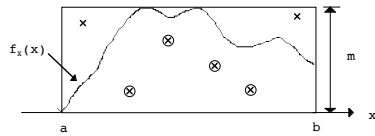
Acceptance-Rejection, Continued

Acceptance-Rejection Algorithm (Finite Support)

1. Generate $U_1, U_2 \stackrel{D}{\sim} \text{Uniform}(0, 1)$ (U_1 and U_2 are independent)
2. Set $Z_1 = a + (b - a)U_1$ and $Z_2 = mU_2$ (inversion method)
3. if $Z_2 \leq f_X(Z_1)$, return $X = Z_1$, else go to step 1

How many (U_1, U_2) pairs must we generate?

- ▶ N (= number pairs generated) has geometric dist'n:
 $P(N = k) = p(1 - p)^{k-1}$ where $p = 1/(\text{area of } R)$
- ▶ So $E[N] = 1/p = (\text{area of } R) = (b - a)m$
- ▶ So make m as small as possible

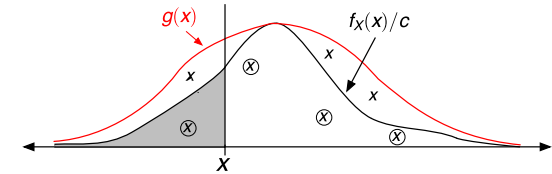


5 / 21

Generalized Acceptance-Rejection: Infinite Support

Find density g that majorizes f_X

- ▶ There exists a constant c such that $f_X(x)/c \leq g(x)$ for all x
- ▶ Smallest such constant is $c = \sup_x (f_X(x)/g(x))$



Generalized Acceptance-Rejection Algorithm

1. Generate $Z \stackrel{D}{\sim} g$ and $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$ (Z, U independent)
2. if $Ug(Z) \leq f_X(Z)/c$, return $X = Z$, else go to step 1

Expected number of (Z, U) pairs generated: $E[N] = c$

6 / 21

Convolution Method

Goal: Generate X where $X = Y_1 + \dots + Y_m$ [Y_1, \dots, Y_m i.i.d.]

- ▶ $f_X = f_{Y_1} * f_{Y_2} * \dots * f_{Y_m}$
- ▶ Where the convolution $f * g$ is defined by
 $(f * g)(x) = \int_{-\infty}^{\infty} f(x - y)g(y) dy$

Convolution Algorithm

- ▶ Generate Y_1, \dots, Y_m
- ▶ Return $X = Y_1 + \dots + Y_m$

Example: Binomial Distribution

- ▶ Suppose $X \stackrel{D}{\sim} \text{Binom}(m, p)$
- ▶ Then $X = Y_1 + \dots + Y_m$ where $Y_i \stackrel{D}{\sim} \text{Bernoulli}(p)$
- ▶ Often part of a more complex algorithm (e.g., do something else if m large)

7 / 21

Composition Method

Suppose that we can write

- ▶ $F_X(x) = p_1 F_{Y_1}(x) + p_2 F_{Y_2}(x) + \dots + p_m F_{Y_m}(x)$ or
- ▶ $f_X(x) = p_1 f_{Y_1}(x) + p_2 f_{Y_2}(x) + \dots + p_m f_{Y_m}(x)$

where p_i 's are nonnegative and $\sum_{i=1}^m p_i = 1$

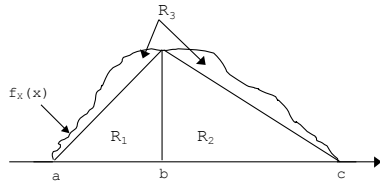
Composition Method

1. Generate a discrete RV J where $P(J = j) = p_j$ for $1 \leq j \leq m$
2. Generate Y_J from F_{Y_j} or f_{Y_j}
3. Return $X = Y_J$

Can lead to very fast generation algorithms

8 / 21

Composition Method: Example



- ▶ Let $p_i = \text{area of } R_i$ for $1 \leq i \leq 3$
- ▶ Then $f_X = p_1 f_{Y_1} + p_2 f_{Y_2} + p_3 f_{Y_3}$, where

$$f_{Y_1}(x) = \begin{cases} \frac{2(x-a)}{(b-a)^2} & \text{if } a \leq x \leq b; \\ 0 & \text{otherwise} \end{cases}$$

$$f_{Y_2}(x) = \begin{cases} \frac{2(c-x)}{(c-b)^2} & \text{if } b \leq x \leq c; \\ 0 & \text{otherwise} \end{cases}$$

$$f_{Y_3}(x) = (1/p_3)(f_X(x) - p_1 f_{Y_1}(x) - p_2 f_{Y_2}(x))$$

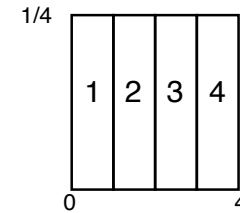
- ▶ Easy to generate Y_1 and Y_2 (the “usual case”):
 - ▶ $Y_1 = \max(U_1, U_2)$ and $Y_2 = \min(U_1, U_2)$ or use inversion

9 / 21

Alias Method for Discrete Random Variables

Goal: Generate X with $P(X = x_i) = p_i$ for $1 \leq i \leq n$

Easy case: $p_1 = p_2 = \dots = p_n$



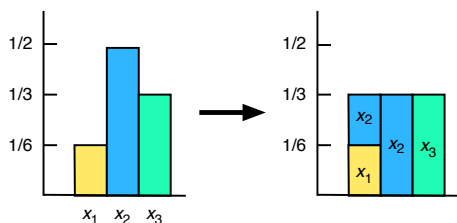
Algorithm

1. Generate $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$
2. Return x_J , where $J = \lceil nU \rceil$

$\lceil x \rceil = \text{smallest integer } \geq x$ (ceiling function)

10 / 21

Alias Method: General Case



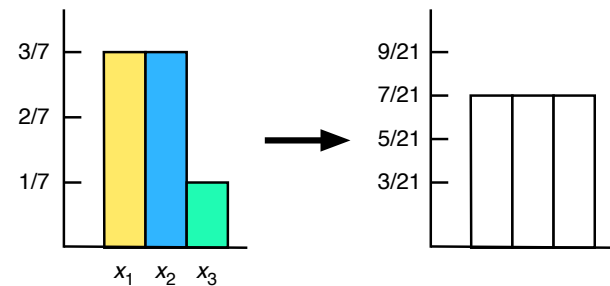
i	x_i	p_i	a_i	r_i
1	x_1	1/6	x_2	0.5
2	x_2	1/2	x_2	1.0
3	x_3	1/3	x_3	1.0

Alias Algorithm

1. Generate $U_1, U_2 \stackrel{D}{\sim} \text{Uniform}(0, 1)$
2. Set $I = \lceil nU_1 \rceil$
3. If $U_2 \leq r_I$ return $X = x_I$ else return $X = a_I$

11 / 21

Alias Method: Another Example



i	x_i	p_i	a_i	r_i
1	x_1	3/7		
2	x_2	3/7		
3	x_3	1/7		

12 / 21

Generation of Non-Uniform Random Numbers

Acceptance-Rejection

Convolution Method

Composition Method

Alias Method

Random Permutations and Samples

Non-Homogeneous Poisson Processes

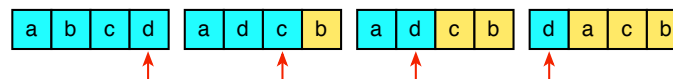
13 / 21

Random Permutations: Fisher-Yates Shuffle

Goal: Create random permutation of array x of length n

Fisher-Yates Algorithm

1. Set $i \leftarrow n$
2. Generate random integer N between 1 and i (e.g., as $\lceil iU \rceil$)
3. Swap $x[N]$ and $x[i]$
4. Set $i \leftarrow i - 1$
5. If $i = 1$ then exit



Q: What about this algorithm:

For $i = 1$ to n : swap $x[i]$ with random entry

Other random objects: graphs, matrices, random vectors, ...

14 / 21

Sequential Bernoulli Sampling

- ▶ Stream of items x_1, x_2, \dots ,
- ▶ Insert each item into sample with probability p
- ▶ Expected sample size after n th item = np
- ▶ Fast implementation: generate skips directly (geometrically distributed)

Bernoulli Sampling:

Generate $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$

Set $\Delta \leftarrow 1 + \lfloor \ln U / \ln(1 - p) \rfloor$ [Geometric on $\{1, 2, \dots\}$, HW #4]

Set $m \leftarrow \Delta$

Upon arrival of x_i :

- ▶ if $i = m$, then
 - ▶ Include x_i in sample
 - ▶ Generate $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$
 - ▶ Set $\Delta \leftarrow 1 + \lfloor \ln U / \ln(1 - p) \rfloor$
 - ▶ set $m \leftarrow m + \Delta$

15 / 21

Reservoir Sampling

- ▶ Stream of items x_1, x_2, \dots ,
- ▶ Maintain a uniform random sample of size N w.o. replacement

Reservoir Sampling:

Upon arrival of x_i :

- ▶ if $i \leq N$, then include x_i in sample
 - ▶ if $i > N$, then
 - ▶ Generate $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$
 - ▶ If $U \leq N/i$, then include x_i in sample, replacing randomly chosen victim
- ▶ Can generate skips directly using acceptance-rejection [JS Vitter, ACM Trans. Math. Softw., 11(1): 37-57, 1985]

16 / 21

Reservoir Sampling: Simple Example

- ▶ Sample size = 1
- ▶ S_i = sample state after processing j th item (called i_j)
- ▶ accept item i_1 into S_1 with probability 1

$$P(i_1 \in S_1) = 1$$

- ▶ accept item i_2 into S_2 with probability $1/2$

$$P(i_1 \in S_2) = P(i_1 \in S_1) \times P(i_2 \notin S_2) = (1)(1/2) = 1/2$$

$$P(i_2 \in S_2) = 1/2$$

- ▶ accept item i_3 into S_3 with probability $1/3$

$$P(i_1 \in S_3) = P(i_1 \in S_2) \times P(i_3 \notin S_3) = (1/2)(2/3) = 1/3$$

$$P(i_2 \in S_3) = P(i_2 \in S_2) \times P(i_3 \notin S_3) = (1/2)(2/3) = 1/3$$

$$P(i_3 \in S_3) = 1/3$$

17 / 21

Generation of Non-Uniform Random Numbers

- Acceptance-Rejection
- Convolution Method
- Composition Method
- Alias Method
- Random Permutations and Samples
- Non-Homogeneous Poisson Processes

18 / 21

Non-Homogeneous Poisson Processes: Thinning

Ordinary (Homogeneous) Poisson process

- ▶ Times between successive events are i.i.d. $\text{Exp}(\lambda)$
 - ▶ Event probabilities for disjoint time intervals are independent (Markov property)
 - ▶ $P(\text{event in } t + \Delta t) \approx \lambda \Delta t$ for Δt very small
 - ▶ $P(T_n > y + z \mid T_{n-1} = y) = e^{-\lambda z}$
- $$P(n \text{ events in } [t, t + \Delta t]) = (\lambda \Delta t)^n e^{-\lambda \Delta t} / n!$$

Non-Homogeneous Poisson process

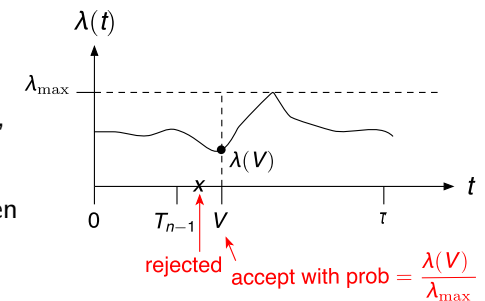
- ▶ Event probabilities for disjoint time intervals are independent
- ▶ $P(\text{event in } t + \Delta t) \approx \lambda(t) \Delta t$ for Δ very small
[$\lambda(t)$ is sometimes called a **hazard function**]
- ▶ $P(T_n > y + z \mid T_{n-1} = y) = e^{-\int_y^{y+z} \lambda(u) du}$
- ▶ Can capture, e.g., time-of-day effects

19 / 21

Thinning, Continued

Suppose that $\lambda(t) \leq \lambda_{\max}$
for $t \in [0, \tau]$

- ▶ Idea: Generate “too many” events according to a $\text{Poisson}(\lambda_{\max})$ process, then reject some of the events

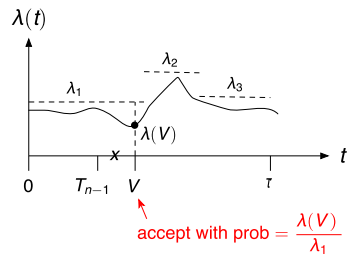


Thinning Algorithm:

1. Set $T_0 = 0$, $V = 0$, and $n = 0$
2. Set $n \leftarrow n + 1$ [Generate T_n]
3. Generate $E \stackrel{D}{\sim} \text{Exp}(\lambda_{\max})$ and $U \stackrel{D}{\sim} \text{Uniform}(0, 1)$
4. Set $V \leftarrow V + E$ [Proposed event time]
5. If $U \leq \lambda(V) / \lambda_{\max}$ then set $T_n = V$ else go to Step 3
6. If $T_n < \tau$ then go to Step 2 else exit

20 / 21

Thinning, Continued



Ross's Improvement

- ▶ Piecewise-constant upper-bounding to reduce rejections
- ▶ Correction for event times that span segments

Many other approaches

- ▶ Inversion (see HW #4)
- ▶ Idiosyncratic methods:
Exploit special properties of Poisson process