

# Graphical Models

## Lecture 5:

## Parameter Estimation & Lagrange Multipliers

Andrew McCallum  
mccallum@cs.umass.edu

Thanks to Noah Smith and Carlos Guestrin for some slide materials.

# Administration

- HW#2 due date

# Learning

- Bayesian Networks can be built by hand.
  - Experts' time is expensive.
  - There may not be any experts.
  - Large models are unwieldy.
  - Knowledge doesn't always transfer across domains.
- Data is often cheap (now).
  - Remember that this was not always the case!

# Notation

- $P^*$  is the true distribution from which our samples were drawn.
- $\mathbf{x}^{(1)}, \mathbf{x}^{(2)}, \dots, \mathbf{x}^{(M)}$  drawn IID from  $P^*$ .

# Goal of Learning?

- **Density estimation:**

Return a model  $M$  that precisely captures  $P^*$

- **Prediction:**


Optimize quality of answers to specific queries,

*e.g.*  $P(x_i | x_j, x_k)$

- **Knowledge discovery:**

Reveal facts about the domain.

# Learning Bayesian Networks

	Known structure	Unknown structure
Fully observed data	 (today)	hard (later)
Missing data	hard (later)	very hard

# MLE Basics

- Likelihood function
- Sufficient statistic: vector representation of the data that summarizes everything you need to compute likelihood
  - If  $\tau(\text{dataset}_1) = \tau(\text{dataset}_2)$  then the likelihood functions are the same.
- For distributions over one random variable, this is usually not hard.
- What about Bayesian networks?

# Key Idea

- For known structure and fully observed data, MLE for a Bayesian network whose CPDs have disjoint parameters

equates to

MLE for each of its CPDs.

- That's it!
- Why?



# Decomposability

$$\begin{aligned}\theta_{\text{MLE}} &= \arg \max_{\theta} \prod_t P(\mathbf{X} = \mathbf{x}^{(t)} \mid \theta) \\ &= \arg \max_{\theta} \prod_t \prod_i P(X_i = x_i^{(t)} \mid \text{Parents}(X_i) = \text{Parents}(x_i), \theta) \\ &= \arg \max_{\theta} \sum_t \sum_i \log P(X_i = x_i^{(t)} \mid \text{Parents}(X_i) = \text{Parents}(x_i), \theta)\end{aligned}$$

If the parameters  $\theta$  are partitioned by CPT ...

$$= \arg \max_{\theta} \sum_i \sum_t \log P(X_i = x_i^{(t)} \mid \text{Parents}(X_i) = \text{Parents}(x_i), \theta_i)$$

(swap order of sums)

# Decomposability

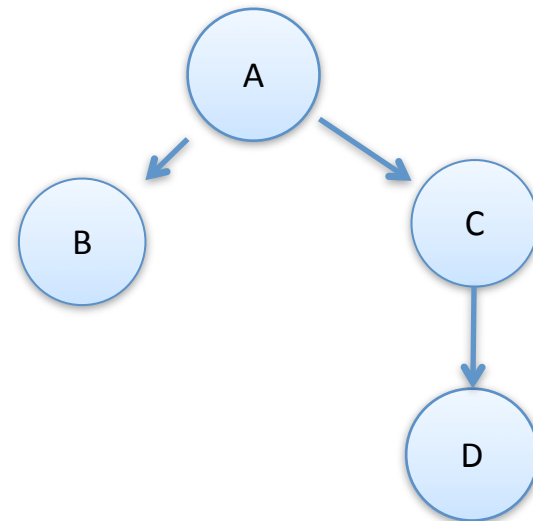
Example

$$\langle a^{(1)}, b^{(1)}, c^{(1)}, d^{(1)} \rangle$$

$$\langle a^{(2)}, b^{(2)}, c^{(2)}, d^{(2)} \rangle$$

⋮

$$\langle a^{(M)}, b^{(M)}, c^{(M)}, d^{(M)} \rangle$$



$$\boldsymbol{\theta} = \langle \boldsymbol{\theta}_A, \boldsymbol{\theta}_{B|A}, \boldsymbol{\theta}_{C|A}, \boldsymbol{\theta}_{D|C} \rangle$$

# Decomposability

$$\theta_{\text{MLE}} = \arg \max_{\theta} \sum_t \log P(A = a^{(t)}, B = b^{(t)}, C = c^{(t)}, D = d^{(t)})$$

$$\begin{aligned} &= \arg \max_{\theta} \sum_t \log P(A = a^{(t)}) + \log P(B = b^{(t)} \mid A = a^{(t)}) \\ &\quad + \log P(C = c^{(t)} \mid A = a^{(t)}) + \log P(D = d^{(t)} \mid C = c^{(t)}) \\ &= \arg \max_{\theta} \sum_t \log P(A = a^{(t)}) + \sum_t \log P(B = b^{(t)} \mid A = a^{(t)}) \\ &\quad + \sum_t \log P(C = c^{(t)} \mid A = a^{(t)}) + \sum_t \log P(D = d^{(t)} \mid C = c^{(t)}) \\ &= \left\langle \arg \max_{\theta_A} \sum_t \log P(A = a^{(t)}), \arg \max_{\theta_{B|A}} \sum_t \log P(B = b^{(t)} \mid A = a^{(t)}), \right. \\ &\quad \left. \arg \max_{\theta_{C|A}} \sum_t \log P(C = c^{(t)} \mid A = a^{(t)}), \arg \max_{\theta_{D|C}} \sum_t \log P(D = d^{(t)} \mid C = c^{(t)}) \right\rangle \end{aligned}$$

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# Deriving the MLE

- Many distributions have a closed form for the MLE.
- Solve (analytically, and with constraints),  $\forall j$ :

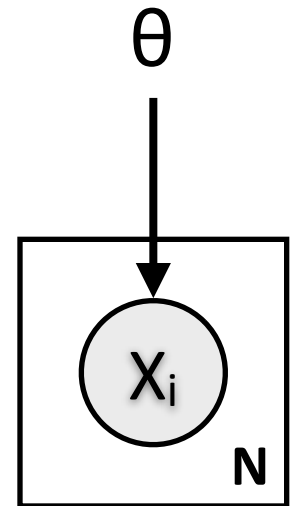
$$\frac{\partial}{\partial \theta_j} \sum_t \log P(X_i = x_i^{(t)} \mid \text{Parents}(X_i) = \text{Parents}(x_i^{(t)})) = 0$$

- Typically *convex*.
- Eg: Gaussian, binomial, multinomial.
- Today: Binomial and multinomial, with Lagrange Multipliers.

# Binomial Distribution

- $P(Y = \text{heads}) = \theta$ ,  $P(Y = \text{tails}) = 1 - \theta$
- “IID” assumption
  - Each flip is independent of the others.
  - All flips are distributed identically.

$$P(\mathbf{Y} \mid \theta, N) = \theta^{\#\text{heads}(\mathbf{Y})} \times (1 - \theta)^{\#\text{tails}(\mathbf{Y})}$$





# Maximum Likelihood Estimation

- Data: sequence  $\mathbf{Y}$  of flip outcomes
- Assumption: binomial distribution; flips are IID
- Goal: select  $\theta$
- Maximum likelihood estimation: treat this as an optimization problem over  $\theta$

$$\begin{aligned}\theta_{\text{MLE}} &= \arg \max_{\theta} P(\mathbf{Y} \mid \theta) \\ &= \arg \max_{\theta} \log P(\mathbf{Y} \mid \theta)\end{aligned}$$

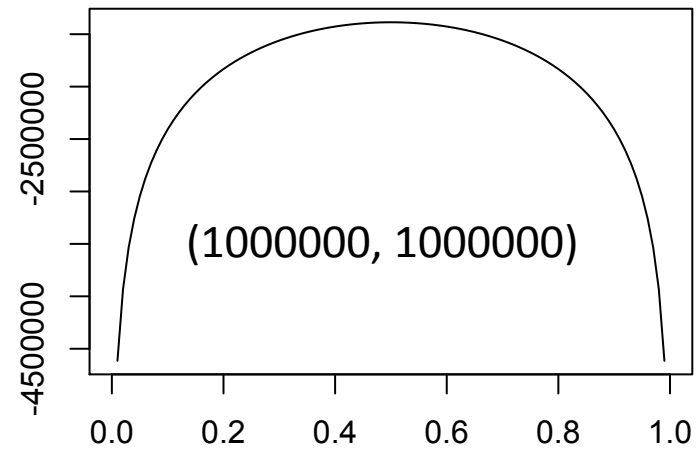
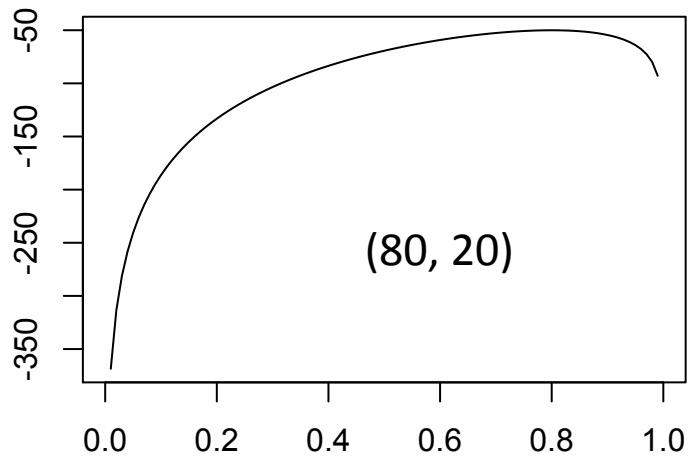
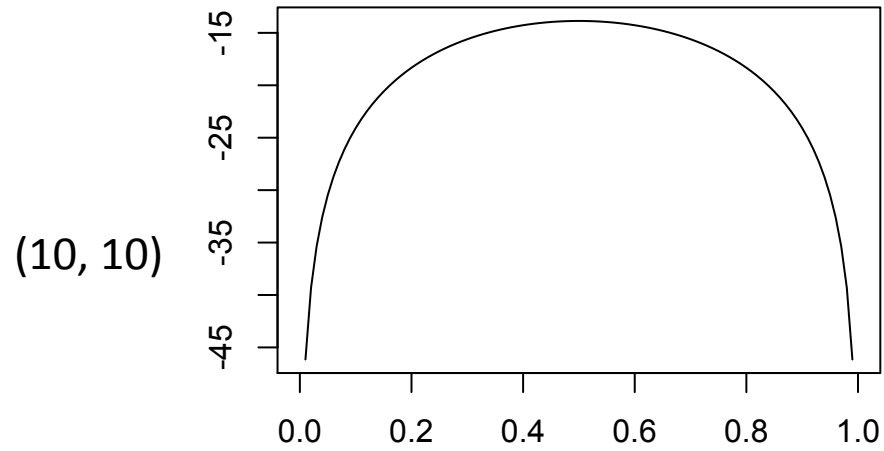
# MLE for the Binomial

$$\begin{aligned}\theta_{\text{MLE}} &= \arg \max_{\theta} P(\mathbf{Y} \mid \theta) \\ &= \arg \max_{\theta} \log P(\mathbf{Y} \mid \theta)\end{aligned}$$

$$P(\mathbf{Y} \mid \theta, N) = \theta^{\#\text{heads}(\mathbf{Y})} \times (1 - \theta)^{\#\text{tails}(\mathbf{Y})}$$

$$\arg \max_{\theta} \#\text{heads}(\mathbf{Y}) \log \theta + \#\text{tails}(\mathbf{Y}) \log(1 - \theta)$$

# MLE for the Binomial



# Deriving the Binomial MLE

- Board work
- Use a little calculus...

# Deriving the Multinomial MLE

- Board work
- Introduce Lagrange Multipliers
- Use them to solve for MLE of a multinomial.

# Deriving Functional Form for Maximum Entropy Classifiers

- Board work
- Lagrange again...

# Generalized Linear Model

- Score is defined as a *linear* function of  $\mathbf{X}$ :

$$f(\mathbf{X}) = w_0 + \underbrace{\sum_i w_i X_i}_Z$$

$Z = f(\mathbf{X})$  is a  
random variable

- Probability distribution over binary value  $Y$  is defined by:

$$P(Y = 1) = \text{sigmoid}(f(\mathbf{X}))$$

- Sample  $Y$ .

From lecture 3!

$$\text{sigmoid}(z) = \frac{e^z}{1 + e^z}$$

# Markov Networks as a Generalized Linear Model

- Sigmoid equates to *binary* output log-linear model.
- More generally, *multinomial* logit: take a linear score ( $Z$  in lecture 3), exponentiate, and normalize ( $Z$  in Gibbs dist.)
  - Don't confuse the  $Z$ s.
- The generalized linear model we used for CPDs is a log-linear distribution.