

## Problem Set 3: Solutions

## 1. [Cover and Thomas 7.1]

(a) Define the following notation,

$$\begin{aligned} C &= I_{p^*(x)}(X; Y) \\ &= \max_{p(x)} I_{p(x)}(X; Y) \\ \tilde{C} &= I_{\tilde{p}^*(x)}(X; \tilde{Y}) \\ &= \max_{p(x)} I_{p(x)}(X; \tilde{Y}) \end{aligned}$$

We would like to show that  $\tilde{C} = I_{\tilde{p}^*(x)}(X; \tilde{Y}) \leq I_{p^*(x)}(X; Y) = C$ .

Notice that  $X, Y$ , and  $\tilde{Y}$  form a Markov chain such that  $X \rightarrow Y \rightarrow \tilde{Y}$ . Using the data-processing inequality (Theorem 2.8.1), we know that,

$$I_{\tilde{p}^*(x)}(X; \tilde{Y}) \leq I_{\tilde{p}^*(x)}(X; Y) \quad (3.1)$$

$$\leq I_{p^*(x)}(X; Y) \quad (3.2)$$

(b) We would like to determine under what conditions the following equality holds. Given our result in Equation 3.2, it is sufficient to show,

$$I_{\tilde{p}^*(x)}(X; \tilde{Y}) \geq I_{p^*(x)}(X; Y)$$

We know that the following equality is true for Markov chains (see proof of Theorem 2.8.1),

$$I_{p^*(x)}(X; \tilde{Y}) = I_{p^*(x)}(X; Y) - I_{p^*(x)}(X; Y|\tilde{Y})$$

However,  $\tilde{p}^*(x)$  and  $p^*(x)$  may not be the same distribution, so

$$I_{\tilde{p}^*(x)}(X; \tilde{Y}) \geq I_{p^*(x)}(X; \tilde{Y}) \quad (3.3)$$

$$= I_{p^*(x)}(X; Y) - I_{p^*(x)}(X; Y|\tilde{Y}) \quad (3.4)$$

We can show our objective inequality if  $I_{p^*(x)}(X; Y|\tilde{Y}) = 0$ . This occurs if  $\tilde{Y} = g(Y)$  is an injective function.

## 2. [Cover and Thomas 7.2]

Consider the behavior of this channel as depicted in Figure 3.1.

When  $|a| \neq 1$ , this is a Noisy Channel with Nonoverlapping Outputs. We would like to compute the capacity of the channel in this situation,

$$\begin{aligned} C &= \max_{p(x)} I(X; Y) \\ &= \max_{p(x)} H(X) - H(X|Y) \end{aligned}$$

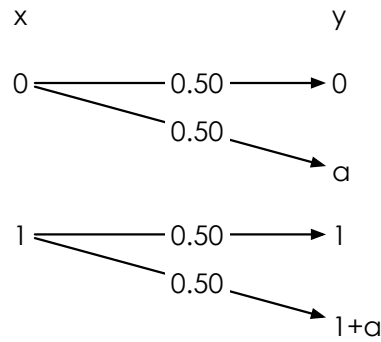


Figure 3.1: Noisy channel model for question 7.2.

Because  $X$  can be determined by  $Y$ ,  $H(X|Y) = 0$ . Therefore,

$$\begin{aligned} C &= \max_{p(x)} H(X) \\ &= 1 \text{ bit} \end{aligned}$$

When  $|a| = 1$ , this is a Binary Erasure Channel. We will compute the capacity for  $a = 1$ . We begin by defining the conditional entropy,

$$\begin{aligned} H(X|Y) &= \sum_{y \in Y} P(Y = y) H(X|Y = y) \\ &= \frac{1}{4} H(X|Y = 0) + \frac{1}{2} H(X|Y = 1) + \frac{1}{4} H(X|Y = 2) \\ &= \frac{1}{2} H(X) \end{aligned}$$

We can now compute the capacity,

$$\begin{aligned} C &= \max_{p(x)} I(X; Y) \\ &= \max_{p(x)} H(X) - H(X|Y) \\ &= \max_{p(x)} H(X) - \frac{1}{2} H(X) \\ &= \max_{p(x)} \frac{1}{2} H(X) \\ &= \frac{1}{2} \max_{p(x)} H(X) \\ &= \frac{1}{2} \text{ bit} \end{aligned}$$

The computation for  $a = -1$  is similar

### 3. [Cover and Thomas 7.3]

$$I(\vec{X}; \vec{Y}) = H(\vec{X}) - H(\vec{X}|\vec{Y})$$

However, because this is a binary symmetric channel, the uncertainty about  $X_i$  and  $Z_i$  is equivalent given  $Y_i$ . We can replace  $\vec{X}$  with  $\vec{Z}$ ,

$$I(\vec{X}; \vec{Y}) = H(\vec{X}) - H(\vec{Z}|\vec{Y}) \tag{3.5}$$

Now we will derive a bound for  $H(\vec{Z}|\vec{Y})$  using properties of conditional entropy (Theorems 2.6.5 and 2.6.6).

$$\begin{aligned} H(\vec{Z}|\vec{Y}) &\leq H(\vec{Z}) \\ &\leq \sum_{i=1}^n H(Z_i) \\ &= nH(p) \end{aligned}$$

Replacing  $H(\vec{Z}|\vec{Y})$  with  $nH(p)$  in Equation 3.5 will reduce the right hand side. This gives us the following inequality.

$$I(\vec{X}; \vec{Y}) \geq H(\vec{X}) - nH(p) \tag{3.6}$$

Define the following notation,

$$H_{\tilde{p}^*(x)}(\vec{X}) - nH(p) = \max_{p(x)} H_{p(x)}(\vec{X}) - nH(p)$$

This defines the maximum value of the right hand side of Equation 3.6. Assuming that  $H(\vec{X}) = \sum H(X_i)$ , the maximizing distribution,  $\tilde{p}^*(x)$ , is uniform. This means that

$$\begin{aligned} H_{\tilde{p}^*(x)}(\vec{X}) - nH(p) &= n - nH(p) \\ &= n(1 - H(p)) \\ &= nC \end{aligned}$$

We are interested in the capacity of the channel with memory.

$$\max_{p(\vec{x})} I_{p(\vec{x})}(\vec{X}; \vec{Y}) \geq nC$$

4. [Cover and Thomas 7.8]

We define our set of distributions as,

$$\begin{aligned} p(x) &= \begin{bmatrix} 1 - \lambda \\ \lambda \end{bmatrix} \\ p(y|x) &= \begin{bmatrix} 1 & 0 \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \\ p(y) &= \begin{bmatrix} 1 - \lambda \\ \lambda \end{bmatrix}^T \times \begin{bmatrix} 1 & 0 \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \\ &= \begin{bmatrix} 1 - \frac{\lambda}{2} & \frac{\lambda}{2} \end{bmatrix} \end{aligned}$$

First, we compute the entropy,  $H(Y)$ ,

$$\begin{aligned} H(Y) &= -1 \times \begin{bmatrix} 1 - \frac{\lambda}{2} & \frac{\lambda}{2} \end{bmatrix} \times \log \begin{bmatrix} 1 - \frac{\lambda}{2} \\ \frac{\lambda}{2} \end{bmatrix} \\ &= -1 \times \left( \left(1 - \frac{\lambda}{2}\right) \log \left(1 - \frac{\lambda}{2}\right) + \left(\frac{\lambda}{2}\right) \log \left(\frac{\lambda}{2}\right) \right) \end{aligned}$$

Next, we compute the conditional entropy,  $H(Y|X)$ ,

$$\begin{aligned} H(Y|X) &= p(X=0) H(Y|X=0) + p(X=1) H(Y|X=1) \\ &= (1-\lambda) H\left(\begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix}\right) + \lambda H\left(\begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ 1 & 1 \end{bmatrix}\right) \\ &= (1-\lambda) \cdot 0 + \lambda \cdot 1 \\ &= \lambda \end{aligned}$$

We can use  $H(Y)$  and  $H(Y|X)$  to compute the capacity of the channel,

$$\begin{aligned} C &= \max_{p(x)} I(X; Y) \\ &= \max_{p(x)} H(Y) - H(Y|X) \\ &= \max_{p(x)} -1 \times \left( \left(1 - \frac{\lambda}{2}\right) \log\left(1 - \frac{\lambda}{2}\right) + \left(\frac{\lambda}{2}\right) \log\left(\frac{\lambda}{2}\right) \right) - \lambda \end{aligned}$$

Notice that we are maximizing a function of  $\lambda$ ,

$$f(\lambda) = -1 \times \left( \left(1 - \frac{\lambda}{2}\right) \log\left(1 - \frac{\lambda}{2}\right) + \left(\frac{\lambda}{2}\right) \log\left(\frac{\lambda}{2}\right) \right) - \lambda$$

To find the maximum of this function, we differentiate with respect to *lambda*,

$$\begin{aligned} \frac{df}{d\lambda} &= - \left(1 - \frac{\lambda}{2}\right) \times \frac{1}{1 - \frac{\lambda}{2}} \times \left(-\frac{1}{2}\right) \\ &\quad - \left(-\frac{1}{2}\right) \times \log\left(1 - \frac{\lambda}{2}\right) \\ &\quad - \frac{\lambda}{2} \times \frac{1}{\frac{\lambda}{2}} \times \frac{1}{2} \\ &\quad - \frac{1}{2} \times \log\left(\frac{\lambda}{2}\right) \\ &\quad - 1 \\ &= \frac{1}{2} \log\left(1 - \frac{\lambda}{2}\right) - \frac{1}{2} \log\left(\frac{\lambda}{2}\right) - 1 \end{aligned}$$

Setting this to zero, we can derive the maximum,

$$\begin{aligned} \frac{1}{2} \log\left(1 - \frac{\lambda}{2}\right) - \frac{1}{2} \log\left(\frac{\lambda}{2}\right) - 1 &= 0 \\ \lambda &= \frac{2}{5} \\ f(\lambda) &= \log 5 - 2 \text{ bits} \\ &\approx 0.3219 \text{ bits} \end{aligned}$$

##### 5. [Cover and Thomas 7.13]

(a) Given the following distributions,

$$\begin{aligned} p(x) &= \begin{bmatrix} 1 - \lambda \\ \lambda \end{bmatrix} \\ p(y|x) &= \begin{bmatrix} 1 - \alpha - \epsilon & \epsilon & \alpha \\ \epsilon & 1 - \alpha - \epsilon & \alpha \end{bmatrix} \\ p(y) &= \begin{bmatrix} 1 - \lambda \\ \lambda \end{bmatrix}^T \begin{bmatrix} 1 - \alpha - \epsilon & \epsilon & \alpha \\ \epsilon & 1 - \alpha - \epsilon & \alpha \end{bmatrix} \\ &= \begin{bmatrix} 1 - \alpha - \epsilon + \alpha\lambda + 2\epsilon\lambda - \lambda \\ \epsilon - 2\epsilon\lambda + \lambda - \alpha\lambda \\ \alpha \end{bmatrix}^T \end{aligned}$$

We would like to compute the capacity of this channel,

$$\begin{aligned} C &= \max_{\lambda} I(X; Y) \\ &= \max_{\lambda} H(Y) - H(Y|X) \end{aligned}$$

However, we can show that  $H(Y|X)$  does not depend on  $\lambda$ ,

$$\begin{aligned} H(Y|X) &= p(X=0) H(Y|X=0) + p(X=1) H(Y|X=1) \\ &= (1-\lambda) H(\begin{bmatrix} 1 - \alpha - \epsilon & \epsilon \\ \epsilon & 1 - \alpha - \epsilon \end{bmatrix}) + \lambda H(\begin{bmatrix} \epsilon & 1 - \alpha - \epsilon \\ \epsilon & 1 - \alpha - \epsilon \end{bmatrix}) \\ &= H(\begin{bmatrix} 1 - \alpha - \epsilon & \epsilon \\ \epsilon & 1 - \alpha - \epsilon \end{bmatrix}) \end{aligned}$$

This means we only need to find the  $\lambda$  maximizing  $H(Y)$ . We could differentiate using our calculation of  $p(y)$ . Instead, we use the method from Section 7.1.5, defining  $E$  be the event that  $\{Y = e\}$ . We can use this to derive  $H(Y)$ .

$$\begin{aligned} H(Y) &= H(Y, E) \\ &= H(E) + H(Y|E) \\ &= H(E) + (1 - \alpha)H(Y|E = 0) \end{aligned}$$

where the last line follows from the fact that  $H(Y|E = 1) = 0$ . Because  $H(E)$  is not a function of  $\lambda$ , we can leave it here. Therefore, we want to maximize  $H(Y|E = 0)$ . So we need to compute  $p(y|E = 0)$ ,

$$\begin{aligned} p(Y = 0|E = 0) &= \frac{P(E = 0|Y = 0)P(Y = 0)}{P(E = 0)} \\ &= \frac{1 + \alpha - \epsilon + \alpha\lambda + 2\epsilon\lambda - \lambda}{1 - \alpha} \\ p(Y = 1|E = 0) &= \frac{P(E = 0|Y = 1)P(Y = 1)}{P(E = 0)} \\ &= \frac{\epsilon - 2\epsilon\lambda + \lambda - \alpha\lambda}{1 - \alpha} \\ &= 1 - P(Y = 0|E = 0) \end{aligned}$$

Again, we could differentiate  $H(Y|E = 0)$  with respect to  $\lambda$  but that's hairy. Instead, we'll recall that  $H(Y|E = 0) \leq 1$  with equality when  $p(Y = 0|E = 0) = p(Y = 1|E = 0)$ .

$$\begin{aligned} p(Y = 0|E = 0) &= p(Y = 1|E = 0) \\ \frac{1 - \alpha - \epsilon + \alpha\lambda^* + 2\epsilon\lambda^* - \lambda^*}{1 - \alpha} &= \frac{\epsilon - 2\epsilon\lambda^* + \lambda^* - \alpha\lambda^*}{1 - \alpha} \\ \lambda^* &= \frac{1}{2} \end{aligned}$$

The channel capacity is,

$$\begin{aligned} C &= \max_{\lambda} I(X; Y) \\ &= H(E) + (1 - \alpha)H_{\lambda^*}(Y|E=0) - H(Y|X) \\ &= H([\alpha \ 1 - \alpha]) + \frac{1 - \alpha}{1 - \alpha} - H([1 - \alpha - \epsilon \ \epsilon \ \alpha]) \end{aligned}$$

(b) In the situation where  $\alpha = 0$ ,

$$\begin{aligned} C &= H([0 \ 1]) + \frac{1 - 0}{1 - 0} - H([1 - \epsilon \ \epsilon \ 0]) \\ &= \frac{1}{1 - \epsilon} - H([1 - \epsilon \ \epsilon]) \end{aligned}$$

(c) In the situation where  $\epsilon = 0$ ,

$$\begin{aligned} C &= H([\alpha \ 1 - \alpha]) + \frac{1 - \alpha}{1 - \alpha} - H([1 - \alpha \ 0 \ \alpha]) \\ &= \frac{1 - \alpha}{1 - \alpha} - H([1 - \alpha \ 0 \ \alpha]) \end{aligned}$$

### 6. [Cover and Thomas 7.15]

Given the following distributions,

$$\begin{aligned} p(x) &= \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix} \\ p(y) &= \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \end{bmatrix} \\ p(x, y) &= \begin{bmatrix} 0.45 & 0.05 \\ 0.05 & 0.45 \end{bmatrix} \\ p(y|x) &= \begin{bmatrix} 0.90 & 0.10 \\ 0.10 & 0.90 \end{bmatrix} \end{aligned}$$

(a)

$$\begin{aligned} H(X) &= 1 \text{ bit} \\ H(Y) &= 1 \text{ bit} \\ H(X, Y) &\approx 1.469 \text{ bits} \\ I(X; Y) &= H(X) + H(Y) - H(X, Y) \\ &\approx 0.531 \text{ bits} \end{aligned}$$

(b) For  $X^n$ ,

$$\begin{aligned} \left| -\frac{1}{n} \log p(x^n) - H(X) \right| &< \epsilon \\ \left| -\frac{1}{n} \log \left( \frac{1}{2} \right)^2 - H(X) \right| &< \epsilon \\ 0 &< \epsilon \end{aligned}$$

Therefore, all  $X^n$  are typical. The proof for  $Y^n$  is similar.

(c)

$$\begin{aligned} H(X, Y, Z) &= H(X, Y) + H(Z|X, Y) & H(X, Y, Z) &= H(X, Z) + H(Y|X, Z) \\ &= H(X, Y) & &= H(X, Z) \end{aligned}$$

$$\begin{aligned}
 H(X, Z) &= H(X) + H(Z|X) \\
 &= H(X) + H(Z) && \text{since } Z \text{ is independent of } X \\
 &= H(X, Y)
 \end{aligned} \tag{3.7}$$

We also know that  $z^n$  is typical. Therefore,

$$\begin{aligned}
 \epsilon &> \left| -\frac{1}{n} \log p(z^n) - H(Z) \right| && z^n \text{ is typical} \\
 &= \left| -\frac{1}{n} \log p(z^n) - H(Z) + \left( -\frac{1}{n} \log p(x^n) - H(X) \right) \right| && \text{shown in part b} \\
 &= \left| -\frac{1}{n} (\log p(z^n) + \log p(x^n)) - (H(Z) + H(X)) \right| \\
 &= \left| -\frac{1}{n} (\log p(z^n)p(x^n)) - H(X, Y) \right| && \text{Equation 3.7} \\
 &= \left| -\frac{1}{n} (\log p(x^n, y^n)) - H(X, Y) \right| && \text{Equation 7.161 in the text}
 \end{aligned}$$

Therefore,  $(x^n, y^n)$  is jointly typical.

(d) By inspecting  $p(x, y)$ , above, we know that,

$$p(x) = [0.90 \quad 0.10]$$

By the definition of typicality, we know that if  $z^n$  is in the set, then

$$\begin{aligned}
 H(Z) - \epsilon &< -\frac{1}{n} \log p(z^n) < H(Z) + \epsilon \\
 H([0.90 \quad 0.10]) - 0.20 &< -\frac{1}{n} \log p(z^n) < H([0.90 \quad 0.10]) + 0.20 \\
 0.269 &< -\frac{1}{n} \log p(z^n) < 0.669
 \end{aligned}$$

This corresponds to  $k = 1, 2, 3, 4$ . Therefore  $|A_{0.10}^{25}(Z)| = 15275$ .

(e)

$$\begin{aligned}
 Pr((x^n(i), Y^n) \in A_\epsilon^n(X, Y)) &= Pr(Y^n - x^n(i) \in A_\epsilon^n(Z)) \\
 &= Pr(x^n(i) + Z^n - x^n(i) \in A_\epsilon^n(Z)) \\
 &= Pr(Z^n \in A_\epsilon^n(Z)) \\
 &= \sum_{z^n \in A_\epsilon^n(Z)} p(z^n) \\
 &= \sum_{z^n \in A_\epsilon^n(Z)} p^k (1-p)^{n-k} \\
 &= \sum_{k=1}^4 \binom{n}{k} p^k (1-p)^{n-k} \\
 &\approx 0.8302
 \end{aligned}$$

(f)

$$\begin{aligned}
Pr((X^n, y^n) \in A_\epsilon^n(X, Y)) &= Pr(y^n - X^n \in A_\epsilon^n(Z)) \\
&= \sum_{x^n} Pr(y^n - x^n \in A_\epsilon^n(Z)) \\
&= \sum_{z^n \in A_\epsilon^n(Z)} p(x^n) \\
&= \sum_{z^n \in A_\epsilon^n(Z)} \frac{1}{2^n} \\
&= \frac{|A_\epsilon^n(Z)|}{2^n}
\end{aligned}$$

(g)

(h)

## 7. [Cover and Thomas 7.20]

(a)

$$\begin{aligned}
I(X; Y_1, Y_2) &= H(Y_1, Y_2) - H(Y_1, Y_2|X) \\
&= H(Y_1) + H(Y_2|Y_1) - H(Y_1|X) \\
&\quad - H(Y_2|Y_1, X) \\
&= H(Y_1) + H(Y_2) - I(Y_1, Y_2) - H(Y_1|X) \\
&\quad - H(Y_2|Y_1, X) \\
&= H(Y_1) - H(Y_1|X) + H(Y_2) - H(Y_2|X) \\
&\quad + H(Y_2|X) - H(Y_2|Y_1, X) - I(Y_1, Y_2) \\
&= I(Y_1; X) + I(Y_2; X) + I(Y_2, Y_1|X) - I(Y_1, Y_2) \\
&= I(Y_1; X) + I(Y_2; X) - I(Y_1, Y_2) \\
&= 2I(Y_1; X) - I(Y_1, Y_2)
\end{aligned}$$

 $Y_1$  and  $Y_2$  conditionally independent given  $X$  $Y_1$  and  $Y_2$  identically distributed given  $X$ 

(b)

$$\begin{aligned}
C_{X \rightarrow (Y_1, Y_2)} &= \max_{p(x)} I(X; Y_1, Y_2) \\
&= \max_{p(x)} (2I(X; Y_1) - I(Y_1, Y_2)) \\
&\leq \max_{p(x)} 2I(X; Y_1) && \text{since } I(Y_1; Y_2) \geq 0 \\
&= 2 \max_{p(x)} I(X; Y_1) \\
&= 2C_{X \rightarrow Y_1}
\end{aligned}$$

## 8. [Cover and Thomas 7.30]

(a)

$$\begin{aligned}
C &= \max_{p(x)} I(X; Y) \\
&= \max_{p(x)} H(X) - H(X|Y)
\end{aligned}$$



We cleverly select  $\mathcal{Z}$  so that  $H(X|Y) = 0$ . This occurs when  $\mathcal{Z}$  results in a channel with nonoverlapping outputs. One such set of values is  $\mathcal{Z} = \{4, 8, 12\}$ .

We pick the uniform distribution over  $\mathcal{X}$  to maximize  $H(X)$ . The entropy for this distribution is  $\log |\mathcal{X}| = 2$  bits. This is also our maximum channel capacity.

(b)

$$\begin{aligned} H(X, Y, Z) &= H(X, Y, Z) \\ H(X, Y|Z) + H(Z) &= H(X, Z|Y) + H(Y) \\ H(X|Z) + H(Y|X, Z) + H(Z) &= H(X|Y) + H(Z|X, Y) + H(Y) \\ H(X|Z) + H(Z) &= H(X|Y) + H(Y) \\ H(X) + H(Z) &= H(X|Y) + H(Y) \\ H(X) - H(X|Y) &= H(Y) - H(Z) \\ I(X; Y) &= H(Y) - H(Z) \\ I(X; Y) &= H(Y) - \log 3 \end{aligned}$$

Therefore  $\min I(X; Y) = \min H(Y)$ . The minimum entropy for  $Y$  occurs when  $|\mathcal{Y}|$  is small. This occurs when  $\mathcal{Z}$  is a set of 3 consecutive integers. In this case,  $\mathcal{Y}$  is a set of six consecutive integers. In the case of  $\mathcal{Z} = \{0, 1, 2\}$ , we have

$$\begin{aligned} P(Y = 0) &= \frac{1}{3}\lambda_0 \\ P(Y = 1) &= \frac{1}{3}(\lambda_0 + \lambda_1) \\ P(Y = 2) &= \frac{1}{3}(\lambda_0 + \lambda_1 + \lambda_2) \\ P(Y = 3) &= \frac{1}{3}(1 - \lambda_0) \\ P(Y = 4) &= \frac{1}{3}(1 - (\lambda_0 + \lambda_1)) \\ P(Y = 5) &= \frac{1}{3}(1 - (\lambda_0 + \lambda_1 + \lambda_2)) \end{aligned}$$

where  $p(i) = \lambda_i$  and  $\lambda_3 = 1 - \sum_{i=0}^2 \lambda_i$ . We need to find the values of  $\lambda_i$  which maximize  $H(Y)$ . This occurs when  $Y$  is uniformly distributed or,

$$\begin{bmatrix} \lambda_0 \\ \lambda_1 \\ \lambda_2 \end{bmatrix} = \begin{bmatrix} \frac{1}{2} \\ 0 \\ 0 \end{bmatrix}$$

In this case,  $H(Y) = \log 6$  and  $C = \log 6 - \log 3 = 1$ .