COMPSCI 514: ALGORITHMS FOR DATA SCIENCE

Cameron Musco University of Massachusetts Amherst. Spring 2020. Lecture 3

LOGISTICS

By Thursday:

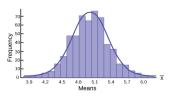
- · Sign up for Piazza.
- Sign up for Gradescope (code on class website) and fill out the Gradescope consent poll on Piazza. Contact me via email if you don't consent to use Gradescope.
- First problem set will be available in the next day or two, due 2/14.

Last Class We Covered:

- Markov's inequality: the most fundamental concentration bound.
- Random hash functions, collision free hashing, and two-level hashing (analysis with linearity of expectation and Markov's inequality.)
- · 2-universal and pairwise independent hash functions.
- Chebyshev's inequality and an elementary proof of the law of large numbers.

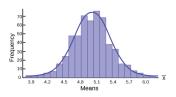
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First: We'll show learn about the union bound and apply it to randomized load balancing.

WORKING APPLICATION

Randomized Load Balancing:



- *n* requests randomly assigned to *k* servers using a random hash function.
- Letting \mathbf{R}_i be the number of requests assigned to server i, $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$ and we provision each server with the capacity to serve twice its expected load: $\frac{2n}{k}$ requests.
- What is the probability that a server exceeds its capacity?
- To apply Chebyshev's inequality, need to bound $\underline{\text{Var}[R_{\text{i}}]}$

Recall that we can write \mathbf{R}_i as:

$$R_i = \sum_{j=1}^n R_{i,j}$$

where $\mathbf{R}_{i,j}$ is 1 if request j is assigned to server i and 0 o.w.

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 (linearity of variance)

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$$Var[\mathbf{R}_{i,j}] = \mathbb{E}\left[\left(\mathbf{R}_{i,j} - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2}\right]$$

$$= Pr(\mathbf{R}_{i,j} = 1) \cdot \left(1 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2} + Pr(\mathbf{R}_{i,j} = 0) \cdot \left(0 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2}$$

$$\frac{1}{K} \left(1 - \frac{1}{K}\right) + \left(1 - \frac{1}{K}\right) \left(0 - \frac{1}{K}\right)^{2}$$

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$$= \frac{1}{k} \cdot \left(1 - \frac{1}{k}\right)^{2} + \left(1 - \frac{1}{k}\right) \cdot \left(0 - \frac{1}{k}\right)^{2}$$

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= \frac{1}{k} \cdot \left(1 - \frac{1}{k}\right)^{2} + \left(1 - \frac{1}{k}\right) \cdot \left(0 - \frac{1}{k}\right)^{2} \\
= \frac{1}{k} - \frac{1}{k^{2}} \leq \frac{1}{k}$$

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$$\begin{aligned} \text{Var}[\mathbf{R}_{i,j}] &= \mathbb{E}\left[\left(\mathbf{R}_{i,j} - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2}\right] \\ &= \text{Pr}(\mathbf{R}_{i,j} = 1) \cdot \left(1 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2} + \text{Pr}(\mathbf{R}_{i,j} = 0) \cdot \left(0 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2} \\ &= \frac{1}{k} \cdot \left(1 - \frac{1}{k}\right)^{2} + \left(1 - \frac{1}{k}\right) \cdot \left(0 - \frac{1}{k}\right)^{2} \\ &= \frac{1}{k} - \frac{1}{k^{2}} \le \frac{1}{k} \implies \text{Var}[\mathbf{R}_{i}] \le \frac{n}{k}. \end{aligned}$$

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$$\Pr\left(\mathbf{R}_{i} \geq \frac{2n}{k}\right) \leq \Pr\left(|\mathbf{R}_{i} - \mathbb{E}[\mathbf{R}_{i}]| \geq \frac{n}{k}\right)$$

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· Overload probability is extremely small when $k \ll n!$

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- · Overload probability is extremely small when $k \ll n!$
- · Might seem counterintuitive bound gets worse as *k* grows.
- When k is large, the number of requests each server sees in expectation is very small so the law of large numbers doesn't 'kick in'.

What is the probability that the maximum server load exceeds $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{k}$. I.e., that some server is overloaded if we give each $\frac{2n}{k}$ capacity?

n: total number of requests, k: number of servers randomly assigned requests, R_i : number of requests assigned to server i. $\mathbb{E}[R_i] = \frac{n}{k}$. $\text{Var}[R_i] = \frac{n}{k}$.

What is the probability that the maximum server load exceeds $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{k}$. I.e., that some server is overloaded if we give each $\frac{2n}{k}$ capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right)$$

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What is the probability that the maximum server load exceeds $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{R}$. I.e., that some server is overloaded if we give each $\frac{2n}{R}$ capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right) = \Pr\left(\left[\mathbf{R}_{1} \geq \frac{2n}{k}\right] \cup \left[\mathbf{R}_{2} \geq \frac{2n}{k}\right] \cup \ldots \cup \left[\mathbf{R}_{k} \geq \frac{2n}{k}\right]\right)$$

n: total number of requests, k: number of servers randomly assigned requests, \mathbf{R}_i : number of requests assigned to server i. $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$. $\mathrm{Var}[\mathbf{R}_i] = \frac{n}{k}$.

What is the probability that the maximum server load exceeds $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{h}$. I.e., that some server is overloaded if we give each $\frac{2n}{h}$ capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right) = \Pr\left(\left[\mathbf{R}_{1} \geq \frac{2n}{k}\right] \text{ or } \left[\mathbf{R}_{2} \geq \frac{2n}{k}\right] \text{ or } \dots \text{ or } \left[\mathbf{R}_{k} \geq \frac{2n}{k}\right]\right)$$

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We want to show that $\Pr\left(\bigcup_{i=1}^{k}\left[\mathbf{R}_{i}\geq\frac{2n}{k}\right]\right)$ is small.

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We want to show that $\Pr\left(\bigcup_{i=1}^{k}\left[\mathbf{R}_{i}\geq\frac{2n}{k}\right]\right)$ is small.

How do we do this? Note that $\mathbf{R}_1, \dots, \mathbf{R}_k$ are correlated in a somewhat complex way.

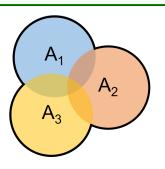
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Union Bound: For any random events $A_1, A_2, ..., A_k$,

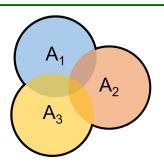
$$\Pr\left(A_1 \cup A_2 \cup \ldots \cup A_k\right) \leq \Pr(A_1) + \Pr(A_2) + \ldots + \Pr(A_k).$$

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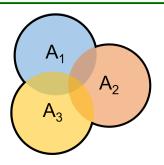
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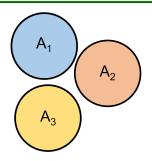
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When is the union bound tight? When $A_1, ..., A_k$ are all disjoint.

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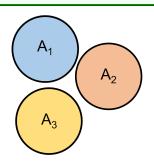
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When is the union bound tight? When $A_1, ..., A_k$ are all disjoint.

On the first problem set, you will prove the union bound, as a consequence of Markov's inquality.

What is the probability that the maximum server load exceeds $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{k}$. I.e., that some server is overloaded if we give each capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right) = \Pr\left(\bigcup_{i=1}^{k} \left[\mathbf{R}_{i} \geq \frac{2n}{k}\right]\right)$$

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APPLYING THE UNION BOUND

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$$\leq \sum_{i=1}^{k} \frac{k}{n} = \frac{k^{2}}{n} \qquad \text{(Bound from Chebyshev's)}$$

As long as $k \le O(\sqrt{n})$, with good probability, the maximum server load will be small (compared to the expected load).

n: total number of requests, k: number of servers randomly assigned requests, \mathbf{R}_i : number of requests assigned to server i. $\mathbb{E}[\mathbf{R}_i] = \frac{n}{b}$. $\mathrm{Var}[\mathbf{R}_i] = \frac{n}{b}$.

ANOTHER VIEW ON THIS PROBLEM

The number of servers must be small compared to the number of requests $(k = O(\sqrt{n}))$ for the maximum load to be bounded in comparison to the expected load with good probability.

n: total number of requests, k: number of servers randomly assigned requests.

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The number of servers must be small compared to the number of requests $(k = O(\sqrt{n}))$ for the maximum load to be bounded in comparison to the expected load with good probability.

 There are many requests routed to a relatively small number of servers so the load seen on each server is close to what is expected via law of large numbers.

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ANOTHER VIEW ON THIS PROBLEM

The number of servers must be small compared to the number of requests $(k = O(\sqrt{n}))$ for the maximum load to be bounded in comparison to the expected load with good probability.

- There are many requests routed to a relatively small number of servers so the load seen on each server is close to what is expected via law of large numbers.
- A Useful Exercise: Given n requests, and assuming all servers have fixed capacity C, how many servers should you provision so that with probability ≥ 99/100 no server is assigned more than C requests?

n: total number of requests, k: number of servers randomly assigned requests.



Questions on union bound, Chebyshev's inequality, random hashing?

$$\mathbb{E}[\mathbf{H}] = \frac{n}{2} = 50 \text{ and } Var[\mathbf{H}] =$$

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We flip n=100 independent coins, each are heads with probability 1/2 and tails with probability 1/2. Let **H** be the number of heads.

$$\mathbb{E}[H] = \frac{n}{2} = 50 \text{ and } Var[H] = \frac{n}{4} = 25$$

Markov's:

$$Pr(H \ge 60) \le .833$$

$$Pr(H \ge 70) \le .714$$

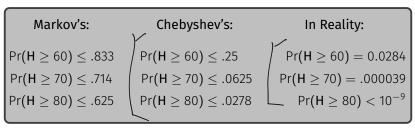
$$Pr(H \ge 80) \le .625$$

$$\mathbb{E}[\mathbf{H}] = \frac{n}{2} = 50 \text{ and } Var[\mathbf{H}] = \frac{n}{4} = 25 \rightarrow s.d. = 5$$

Markov's:	Chebyshev's:	
$Pr(H \ge 60) \le .833$	$Pr(H \ge 60) \le .25$	Ži
$Pr(H \ge 70) \le .714$	$Pr(H \ge 70) \le .0625$	1/47
$Pr(H \ge 80) \le .625$	$Pr(H \ge 80) \le .0278$	162

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$$\mathbb{E}[\mathbf{H}] = \frac{n}{2} = 50 \text{ and } Var[\mathbf{H}] = \frac{n}{4} = 25 \rightarrow s.d. = 5$$



H has a simple Binomial distribution, so can compute these probabilities exactly.

To be fair.... Markov and Chebyshev's inequalities apply much more generally than to Binomial random variables like coin flips.

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• Markov's: $Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$. First Moment.

To be fair.... Markov and Chebyshev's inequalities apply much more generally than to Binomial random variables like coin flips.

Can we obtain tighter concentration bounds that still apply to very general distributions?

- Markov's: $\Pr(\mathbf{X} \geq t) \leq \frac{\mathbb{E}[\mathbf{X}]}{t}$ First Moment. Chebyshev's: $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| \geq t) = \Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]|^2 \geq t^2)$

To be fair.... Markov and Chebyshev's inequalities apply much more generally than to Binomial random variables like coin flips.

Can we obtain tighter concentration bounds that still apply to very general distributions?

- · Markov's: $\Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$. First Moment.
- Chebyshev's: $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| \ge t) = \Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]|^2 \ge t^2) \le \frac{\text{Var}[\mathbf{X}]}{t^2}$. Second Moment.
- · What if we just apply Markov's inequality to even higher moments?

Consider any random variable X:

$$\Pr(|\mathbf{X} - \mathbb{E}[\mathbf{X}]| \ge t) = \Pr((\mathbf{X} - \mathbb{E}[\mathbf{X}])^4 \ge t^4)$$

Consider any random variable X:

From sider any random variable X:
$$\sqrt{t} \quad \text{with} \quad \text{with} \quad \text{with} \quad \text{for some problem}$$

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• Apply Fourth Moment Bound: $\Pr(|\mathbf{H} - \mathbb{E}[\mathbf{H}]| \ge t) \le \frac{1862.5}{t^4}$.

Chebyshev's:

 $Pr(H \ge 60) \le .25$

 $Pr(H \ge 70) \le .0625$

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In Reality:

 $Pr(H \ge 60) = 0.0284$

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- · Why monotonic? $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| > t) = \Pr(f(|\mathbf{X} \mathbb{E}[\mathbf{X}]|) > f(t)).$

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Chernoff bound, Bernstein inequalities, Hoeffding's inequality, Azuma's inequality, Berry-Esseen theorem. etc. **Moment Generating Function:** Consider for any t > 0:

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- · We will not cover the proofs in the this class.

Bernstein Inequality: Consider independent random variables

$$X_1, \ldots, X_n$$
 all falling in $[-M, M]$. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$ and $\sigma^2 = \text{Var}[\sum_{i=1}^n X_i] = \sum_{i=1}^n \text{Var}[X_i]$. For any $t \ge 0$:

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \geq t\right) \leq 2 \exp\left(-\frac{t^{2}}{2\sigma^{2} + \frac{4}{3}Mt}\right).$$

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· An exponentially stronger dependence on s!

COMPARISION TO CHEBYSHEV'S

Consider again bounding the number of heads \mathbf{H} in n=100 independent coin flips.

Chebyshev's:	Bernstein:	In Reality:
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Getting much closer to the true probability.

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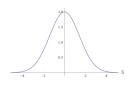
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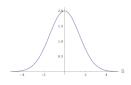
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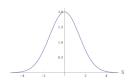
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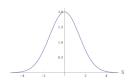
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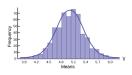
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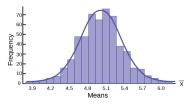
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Central Limit Theorem Interpretation: Bernstein's inequality gives a quantitative version of the CLT. The distribution of the sum of *bounded* independent random variables can be upper bounded with a Gaussian (normal) distribution.



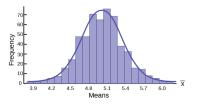
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Stronger Central Limit Theorem: The distribution of the sum of *n bounded* independent random variables converges to a Gaussian (normal) distribution as *n* goes to infinity.



CENTRAL LIMIT THEOREM

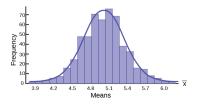
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Stronger Central Limit Theorem: The distribution of the sum of *n bounded* independent random variables converges to a Gaussian (normal) distribution as *n* goes to infinity.



- Why is the Gaussian distribution is so important in statistics, science, ML, etc.?
- Many random variables can be approximated as the sum of a large number of small and roughly independent random effects. Thus, their distribution looks Gaussian by CLT.

THE CHERNOFF BOUND

A useful variation of the Bernstein inequality for binary (indicator) random variables is:

Chernoff Bound (simplified version): Consider independent random variables $\mathbf{X}_1,\ldots,\mathbf{X}_n$ taking values in $\{0,1\}$. Let $\mu=\mathbb{E}[\sum_{i=1}^n\mathbf{X}_i]$. For any $\delta\geq 0$

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \geq \delta \mu\right) \leq 2 \exp\left(-\frac{\delta^{2} \mu}{2 + \delta}\right).$$

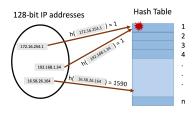
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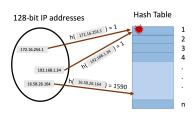
As δ gets larger and larger, the bound falls of exponentially fast.

RETURN TO RANDOM HASHING



We hash m values x_1, \ldots, x_m using a random hash function into a table with n = m entries.

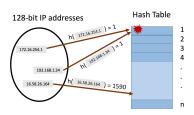
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By the Chernoff Bound: for any $\delta \geq 0$,

$$\Pr(\underline{\mathbf{S}_i \ge 1 + \delta}) \le \Pr\left(\left|\sum_{i=1}^n \mathbf{S}_{i,j} - 1\right| \ge \delta\right) \le 2 \exp\left(-\frac{\delta^2}{2 + \delta}\right)$$

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Apply Union Bound:

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$$\leq \sum_{i=1}^{m} \Pr(\mathbf{S}_{i} \geq 20 \log m + 1) \leq m \cdot \frac{2}{m^{18}} = \frac{2}{m^{17}}.$$

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- · Using Chebyshev's inequality could only show the maximum load is bounded by $O(\sqrt{m})$ with good probability.
- The Chebyshev bound holds even with a pairwise independent hash function. The stronger Chernoff-based bound can be shown to hold with a k-wise independent hash function for $k = O(\log m)$.

Questions?

This concludes probability review/concentration bounds.