COMPSCI 514: Algorithms for Data Science

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Lecture 12

- Problem Set 2 due tomorrow at 11:59pm.
- Solutions will be released shortly after submission so they can be used to study.
- No quiz this week.
- Midterm review office hours: Tuesday 10/15 1:30-3pm Herter Hall 205. Wednesday 10/16 10am-11:30am Goessmann 151. Thursday in class.
- The midterm will be Thursday 7pm-9pm in ILC N151.

Last Class:

- Intro to dimensionality reduction.
- Intro to low-distortion embeddings and the Johnson-Lindenstrauss Lemma.

This Class:

- Reduction of JL Lemma to the Distributional JL Lemma.
- Proof the Distributional JL Lemma.
- Example application of JL to clustering.

Johnson-Lindenstrauss Lemma: For any set of points $\vec{x}_1, \ldots, \vec{x}_n \in \mathbb{R}^d$ and $\epsilon > 0$ there exists a linear map $\boldsymbol{\Pi}: \mathbb{R}^d \to \mathbb{R}^m$ $\frac{1}{2}$ such that $m = O\left(\frac{\log n}{\epsilon^2}\right)$ and letting $\tilde{x}_i = \mathbf{n}\vec{x}_i$:

For all $i, j : (1 - \epsilon) \|\vec{x}_i - \vec{x}_i\|_2 \leq \|\tilde{x}_i - \tilde{x}_i\|_2 \leq (1 + \epsilon) \|\vec{x}_i - \vec{x}_i\|_2.$

Further, if **Π** *∈* R *^m×^d* has each entry chosen i.i.d. from *N* (0*,* 1*/m*), it satisfies the guarantee with high probability.

Distributional JL

The Johnson-Lindenstrauss Lemma is a direct consequence of a closely related lemma:

Distributional JL Lemma: Let **Π** *∈* R *^m×^d* have each entry chosen i.i.d. as $\mathcal{N}(0,1/m)$. If we set $m = O\left(\frac{\log(1/\delta)}{\epsilon^2}\right)$ $\frac{(1/\delta)}{\epsilon^2}$, then for any $\vec{y} \in \mathbb{R}^d$, with probability $\geq 1 - \delta$

$$
(1 - \epsilon) \|\vec{y}\|_2 \le \|\mathbf{\Pi}\vec{y}\|_2 \le (1 + \epsilon) \|\vec{y}\|_2
$$

Applying a random matrix **Π** to any vector *⃗y* preserves *⃗y*'s norm with high probability.

- Like a low-distortion embedding, but for the length of a compressed vector rather than distances between vectors.
- Can be proven from first principles.

^Π *[∈]* ^R*m×^d* : random projection matrix. *d*: original dimension. *m*: compressed dimension, *ϵ*: embedding error, *δ*: embedding failure prob. 5

Distributional JL =*⇒* JL

Distributional JL Lemma =*⇒* JL Lemma: Distributional JL show that a random projection **Π** preserves the norm of any *y*. The main JL Lemma says that **Π** preserves distances between vectors.

Since **Π** is linear these are the same thing!

Proof: Given $\vec{x}_1, \ldots, \vec{x}_n$, define $\binom{n}{2}$ vectors \vec{y}_{ij} where $\vec{y}_{ij} = \vec{x}_i - \vec{x}_j$.

• If we choose **Π** with $m = O\left(\frac{\log 1/\delta}{\epsilon^2}\right)$ $\frac{\zeta^{1/\delta}}{\epsilon^2}$), for each $\vec{\mathcal{Y}}_{ij}$ with probability *≥* 1 *− δ* we have:

 $(1 - \sqrt{16})\vec{v}$ \vec{v} $\vec{$

Distributional JL =*⇒* JL

Claim: If we choose **Π** with i.i.d. *N* (0*,* 1*/m*) entries and $m = O\left(\frac{\log(1/\delta')}{\epsilon^2}\right)$ $\left(\frac{1/\delta'}{\epsilon^2}\right)$, letting $\tilde{\mathbf{x}}_i = \boldsymbol{\Pi} \vec{x}_i$, for each pair \vec{x}_i, \vec{x}_j with probability *≥* 1 *− δ ′* we have:

 $(1-\epsilon)\|\vec{x}_i-\vec{x}_i\|_2 \leq \|\tilde{\mathbf{x}}_i-\tilde{\mathbf{x}}_i\|_2 \leq (1+\epsilon)\|\vec{x}_i-\vec{x}_i\|_2.$

With what probability are all pairwise distances preserved?

Union bound: With probability $\geq 1 - {n \choose 2} \cdot \delta'$ all pairwise distances are preserved.

Apply the claim with $\delta' = \delta / {n \choose 2}$. \implies for $m = O\left(\frac{\log(1/\delta')}{\epsilon^2}\right)$ $\frac{(1/\delta')}{\epsilon^2}\bigg)$, all pairwise distances are preserved with probability *≥* 1 *− δ*.

$$
m = O\left(\frac{\log(1/\delta')}{\epsilon^2}\right) = O\left(\frac{\log(\binom{n}{2}/\delta)}{\epsilon^2}\right) = O\left(\frac{\log(n^2/\delta)}{\epsilon^2}\right) = O\left(\frac{\log(n/\delta)}{\epsilon^2}\right)
$$

Yields the JL lemma.

Distributional JL Lemma: Let **Π** *∈* R *^m×^d* have each entry chosen i.i.d. as $\mathcal{N}(0,1/m)$. If we set $m = O\left(\frac{\log(1/\delta)}{\epsilon^2}\right)$ $\frac{(1/\delta)}{\epsilon^2}$), then for any $\vec{y} \in \mathbb{R}^d$, with probability $\geq 1 - \delta$

$$
(1-\epsilon)\|\vec{y}\|_2 \leq \|\mathbf{\Pi}\vec{y}\|_2 \leq (1+\epsilon)\|\vec{y}\|_2
$$

- \cdot Let $\tilde{\mathsf{y}}$ denote $\boldsymbol{\mathsf{\Pi}}\vec{\mathsf{y}}$ and let $\boldsymbol{\mathsf{\Pi}}(j)$ denote the j^{th} row of $\boldsymbol{\mathsf{\Pi}}.$
- \cdot For any *j*, $\tilde{y}(j) = \langle \mathbf{\Pi}(j),\vec{y} \rangle = \sum_{i=1}^d \mathbf{g}_i \cdot \vec{y}(i)$ where $\mathbf{g}_i \sim \mathcal{N}(0,1/m)$.

⃗^y [∈] ^R*^d* : arbitrary vector, ỹ *[∈]* ^R*m*: compressed vector, **^Π** *[∈]* ^R*m×^d* : random projection. *d*: original dim. *m*: compressed dim, *ϵ*: error, *δ*: failure prob. 8

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- \cdot For any *j*, $\tilde{y}(j) = \langle \mathbf{\Pi}(j),\vec{y} \rangle = \sum_{i=1}^d \mathbf{g}_i \cdot \vec{y}(i)$ where $\mathbf{g}_i \sim \mathcal{N}(0,1/m)$.
- \cdot **g**_{*i*} · $\vec{y}(i) \sim \mathcal{N}(0, \frac{\vec{y}(i)^2}{m})$ $\frac{\delta(\vec{r})^2}{m}$): normally distributed with variance $\frac{\vec{y}(\vec{r})^2}{m}$ $\frac{(l)}{m}$.

What is the distribution of ỹ(*j*)? Also Gaussian!

⃗^y [∈] ^R*^d* : arbitrary vector, ỹ *[∈]* ^R*m*: compressed vector, **^Π** *[∈]* ^R*m×^d* : random projection mapping *⃗y →* ỹ. **Π**(*j*): *j th* row of **Π**, *d*: original dimension. *m*: compressed dimension, **g**_i: normally distributed random variable.

Letting
$$
\tilde{\mathbf{y}} = \mathbf{\Pi} \vec{y}
$$
, we have $\tilde{\mathbf{y}}(j) = \langle \mathbf{\Pi}(j), \vec{y} \rangle$ and:
\n
$$
\tilde{\mathbf{y}}(j) = \sum_{i=1}^{d} \mathbf{g}_i \cdot \vec{y}(i) \text{ where } \mathbf{g}_i \cdot \vec{y}(i) \sim \mathcal{N}\left(0, \frac{\vec{y}(i)^2}{m}\right).
$$

Stability of Gaussian Random Variables. For independent *a ∼* $\mathcal{N}(\mu_1, \sigma_1^2)$ and $b \sim \mathcal{N}(\mu_2, \sigma_2^2)$ we have:

$$
a+b \sim \mathcal{N}(\mu_1+\mu_2, \sigma_1^2+\sigma_2^2)
$$

$$
\bigwedge + \bigwedge = \bigwedge
$$

Thus, $\tilde{y}(j) \sim \mathcal{N}(0, \frac{\vec{y}(1)^2}{m} + \frac{\vec{y}(2)^2}{m} + \ldots + \frac{\vec{y}(d)^2}{m}$ $\frac{d}{m}$ ² $\frac{\|\vec{y}\|_2^2}{m}$) I.e., \tilde{y} itself is a random Gaussian vector. Rotational invariance of the Gaussian distribution.

⃗^y [∈] ^R*^d* : arbitrary vector, ỹ *[∈]* ^R*m*: compressed vector, **^Π** *[∈]* ^R*m×^d* : random projection mapping *⃗y →* ỹ. **Π**(*j*): *j th* row of **Π**, *d*: original dimension. *m*: com-

So far: Letting **Π** *∈* R *^d×^m* have each entry chosen i.i.d. as *^N* (0*,* ¹*/m*), \vec{y} *c* \vec{y} *c* \vec{y} *c* \vec{y} *c* \vec{y} *i* $\vec{y$

 $\tilde{y}(j) \sim \mathcal{N}(0, ||\vec{y}||_2^2/m).$

What is $\mathbb{E}[\|\tilde{\mathsf{y}}\|_2^2]$?

$$
\mathbb{E}[\|\tilde{\mathbf{y}}\|_{2}^{2}] = \mathbb{E}\left[\sum_{j=1}^{m} \tilde{\mathbf{y}}(j)^{2}\right] = \sum_{j=1}^{m} \mathbb{E}[\tilde{\mathbf{y}}(j)^{2}] = \sum_{j=1}^{m} \frac{\|\vec{y}\|_{2}^{2}}{m} = \|\vec{y}\|_{2}^{2}
$$

So \tilde{v} has the right norm in expectation.

How is *∥*ỹ*∥* 2 ² distributed? Does it concentrate?

⃗^y [∈] ^R*^d* : arbitrary vector, ỹ *[∈]* ^R*m*: compressed vector, **^Π** *[∈]* ^R*m×^d* : random projection mapping *⃗y →* ỹ. **Π**(*j*): *j th* row of **Π**, *d*: original dimension. *m*: compressed dimension, g*ⁱ* : normally distributed random variable

So far: Letting **Π** *∈* R *^d×^m* have each entry chosen i.i.d. as *^N* (0*,* ¹*/m*), \vec{y} *c* \vec{y} *c* \vec{y} *c* \vec{y} *c* \vec{y} *i* $\vec{y$

 $\widetilde{\mathbf{y}}(j) \sim \mathcal{N}(0, \|\vec{y}\|_2^2/m)$ and $\mathbb{E}[\|\widetilde{\mathbf{y}}\|_2^2] = \|\vec{y}\|_2^2$

∣ $|\mathbf{\tilde{y}}|_2^2 = \sum_{i=1}^m \mathbf{\tilde{y}}(j)^2$ a Chi-Squared random variable with *m* degrees of freedom (a sum of *m* squared independent Gaussians)

Lemma: (Chi-Squared Concentration) Letting Z be a Chi-Squared random variable with *m* degrees of freedom,

$$
Pr[|Z - \mathbb{E}Z| > \epsilon \mathbb{E}Z] < 2e^{-m\epsilon^2/8}.
$$

Example Application: *k*-means clustering

Goal: Separate *n* points in *d* dimensional space into *k* groups.

Write in terms of distances: $Cost(\mathcal{C}_1, \ldots, \mathcal{C}_k) = \min_{\mathcal{C}_1, \ldots, \mathcal{C}_k}$ ∑ *k j*=1 ∑ $\vec{x}_1, \vec{x}_2 \in C_k$ $||\vec{x}_1 - \vec{x}_2||_2^2$

Example Application: *k*-means clustering

k-means Objective:
$$
Cost(C_1, ..., C_k) = \min_{C_1, ..., C_k} \sum_{j=1}^k \sum_{\vec{x}_1, \vec{x}_2 \in C_k} ||\vec{x}_1 - \vec{x}_2||_2^2
$$

If we randomly project to $m=O\left(\frac{\log n}{\epsilon^2}\right)$ dimensions, for all pairs \vec{x}_1,\vec{x}_2 ,

$$
(1-\epsilon)\|\vec{x}_1-\vec{x}_2\|_2^2\leq \|\tilde{\mathbf{x}}_1-\tilde{\mathbf{x}}_2\|_2^2\leq (1+\epsilon)\|\vec{x}_1-\vec{x}_2\|_2^2\implies
$$

Letting
$$
\overline{Cost}(\mathcal{C}_1, ..., \mathcal{C}_k) = \min_{\mathcal{C}_1, ..., \mathcal{C}_k} \sum_{j=1}^k \sum_{\tilde{\mathbf{x}}_1, \tilde{\mathbf{x}}_2 \in \mathcal{C}_k} ||\tilde{\mathbf{x}}_1 - \tilde{\mathbf{x}}_2||_2^2
$$

 $(1-\epsilon)$ Cost $(\mathcal{C}_1,\ldots,\mathcal{C}_k) \leq \text{Cost}(\mathcal{C}_1,\ldots,\mathcal{C}_k) \leq (1+\epsilon) \text{Cost}(\mathcal{C}_1,\ldots,\mathcal{C}_k).$

Upshot: Can cluster in *m* dimensional space (much more efficiently) and minimize *Cost*(*C*¹ *, . . . , Ck*). The optimal set of clusters will have true cost within $1 + c\epsilon$ times the true optimal. Good exercise to prove this.