# COMPSCI 514: Algorithms for Data Science

Cameron Musco University of Massachusetts Amherst. Fall 2022. Lecture 8

#### Last Class:

- Finish up Bloom filter analysis and optimization of parameters.
- Start on streaming algorithms and distinct elements estimation via hashing.

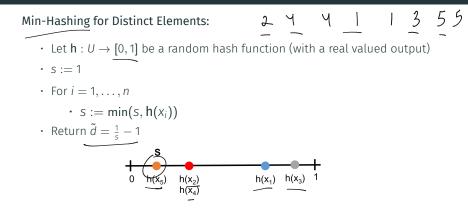
#### Last Class:

- Finish up Bloom filter analysis and optimization of parameters.
- Start on streaming algorithms and distinct elements estimation via hashing.

#### This Class:

- Analysis of the distinct elements algorithm.
- The median trick for boosting success probability.
- Sketch of the ideas behind practical algorithms for distinct elements estimation.

# Hashing for Distinct Elements



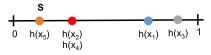
# Hashing for Distinct Elements

Min-Hashing for Distinct Elements:

- + Let  $\mathbf{h}: U \rightarrow [0,1]$  be a random hash function (with a real valued output)
- s := 1
- For  $i = 1, \ldots, n$

• 
$$s := \min(s, h(x_i))$$

• Return  $\tilde{d} = \frac{1}{s} - 1$ 



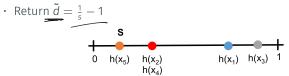
 After all items are processed, s is the minimum of d points chosen uniformly at random on [0, 1]. Where d = # distinct elements.

# Hashing for Distinct Elements

Min-Hashing for Distinct Elements:

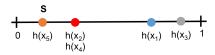
- + Let  $\mathbf{h}: U \rightarrow [0,1]$  be a random hash function (with a real valued output)
- s := 1
- For  $i = 1, \ldots, n$

• 
$$s := \min(s, h(x_i))$$

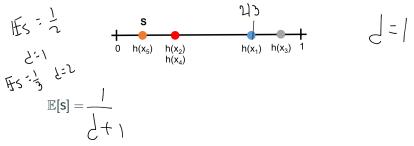


- After all items are processed, s is the minimum of d points chosen uniformly at random on [0, 1]. Where d = # distinct elements.
- Intuition: The larger *d* is, the smaller we expect s to be.

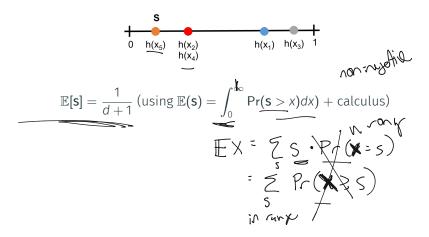
**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



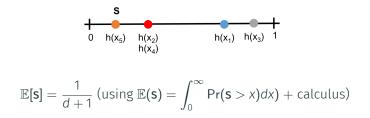
**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



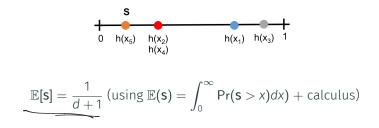
**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



- So our estimate  $\widehat{d} = \frac{1}{s} - 1$  is correct if **s** exactly equals its expectation.

$$\frac{1}{\frac{d+1}{d+1}} - 1 = \frac{1}{\frac{d}{d+1}}$$

**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



• So our estimate  $\hat{\mathbf{d}} = \frac{1}{s} - 1$  is correct if  $\mathbf{s}$  exactly equals its expectation. Does this mean  $\mathbb{E}[\hat{\mathbf{d}}] = d$ ?  $\mathbb{N}^{\mathcal{D}}$ .

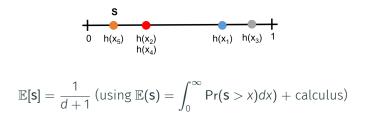
$$ECJ = \frac{1}{2}$$

$$ECJ = \frac{1}{2}$$

$$ECJ = \frac{1}{2}$$

$$ECJ = \frac{1}{2}$$

**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements.



• So our estimate  $\hat{\mathbf{d}} = \frac{1}{s} - 1$  is correct if **s** exactly equals its expectation. Does this mean  $\mathbb{E}[\hat{\mathbf{d}}] = d$ ? No, but:

**s** is the minimum of *d* points chosen uniformly at random on [0, 1]. Where d = # distinct elements. 3  $h(x_1) \quad h(x_3) \quad 1$  $h(x_5)$  $h(x_2)$ h(x₄) 0 11 S  $\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$  (using  $\mathbb{E}(\mathbf{s}) = \int_{0}^{\infty} \Pr(\mathbf{s} > x) dx$ ) + calculus)  $\begin{array}{c} \text{ulus} \\ & S \leq \frac{1}{6} (1 + \varepsilon) \\ & \zeta \leq \frac{1}{5} - 1 \geq \frac{1}{2} - 1 \\ & S \text{ its} \quad (1 + \varepsilon) \end{array}$ • So our estimate  $\hat{\mathbf{d}} = \frac{1}{s} - 1$  is correct if **s** exactly equals its 2 (1-26) 2 expectation. Does this mean  $\mathbb{E}[\hat{\mathbf{d}}] = d$ ? No, but: • Approximation is robust: if  $|\mathbf{s} - \mathbb{E}[\mathbf{s}]| \le \epsilon \cdot \mathbb{E}[\mathbf{s}]$  for any  $\epsilon \in (0, 1/2)$  and a small constant  $c \leq 4$ : 1J-214<Ed  $(1-c\epsilon)d \leq \widehat{\mathbf{d}} \leq (1+c\epsilon)d$ ۱ - ب<u>با</u> بر بر

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$

s: minimum of d distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{s} - 1$ : estimate of # distinct elements d.

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$
 and  $\operatorname{Var}[\mathbf{s}] \leq \frac{1}{(d+1)^2}$  (also via calculus).

**s**: minimum of *d* distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{s} - 1$ : estimate of *#* distinct elements *d*.

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$
 and  $\operatorname{Var}[\mathbf{s}] \leq \frac{1}{(d+1)^2}$  (also via calculus).

Chebyshev's Inequality:

$$\Pr[|\mathbf{s} - \mathbb{E}[\mathbf{s}]| \ge \epsilon \mathbb{E}[\mathbf{s}]] \le \frac{\operatorname{Var}[\mathbf{s}]}{(\epsilon \mathbb{E}[\mathbf{s}])^2}$$

s: minimum of d distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{5} - 1$ : estimate of # distinct elements d.

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$
 and  $\operatorname{Var}[\mathbf{s}] \le \frac{1}{(d+1)^2}$  (also via calculus).

Chebyshev's Inequality:

$$\Pr\left[\underline{|s} - \mathbb{E}[s]\right] \ge \epsilon \mathbb{E}[s] \le \frac{\operatorname{Var}[s]}{(\epsilon \mathbb{E}[s])^2} = \frac{1}{\epsilon^2}. \ge 1$$

s: minimum of d distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{5} - 1$ : estimate of # distinct elements d.

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$
 and  $\operatorname{Var}[\mathbf{s}] \leq \frac{1}{(d+1)^2}$  (also via calculus).

Chebyshev's Inequality:

$$\Pr\left[|\mathbf{S} - \mathbb{E}[\mathbf{S}]| \ge \epsilon \mathbb{E}[\mathbf{S}]\right] \le \frac{\operatorname{Var}[\mathbf{S}]}{(\epsilon \mathbb{E}[\mathbf{S}])^2} = \frac{1}{\epsilon^2}.$$

Bound is vacuous for any  $\epsilon < 1$ .

**s**: minimum of *d* distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{s} - 1$ : estimate of *#* distinct elements *d*.

$$\mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$$
 and  $\operatorname{Var}[\mathbf{s}] \leq \frac{1}{(d+1)^2}$  (also via calculus).

Chebyshev's Inequality:

$$\Pr\left[|\mathbf{S} - \mathbb{E}[\mathbf{S}]| \ge \epsilon \mathbb{E}[\mathbf{S}]\right] \le \frac{\operatorname{Var}[\mathbf{S}]}{(\epsilon \mathbb{E}[\mathbf{S}])^2} = \frac{1}{\epsilon^2}.$$

Bound is vacuous for any  $\epsilon$  < 1. How can we improve accuracy?

**s**: minimum of *d* distinct hashes chosen randomly over [0, 1], computed by hashing algorithm.  $\hat{\mathbf{d}} = \frac{1}{5} - 1$ : estimate of *#* distinct elements *d*.

Leverage the law of large numbers: improve accuracy via repeated independent trials.

Leverage the law of large numbers: improve accuracy via repeated independent trials.

#### Hashing for Distinct Elements (Improved):

- + Let  $\underline{h}: U \to [0, 1]$  be a random hash function
- s := 1
- For  $i = 1, \ldots, n$ 
  - $s := \min(s, h(x_i))$
- Return  $\widehat{d} = \frac{1}{s} 1$

Leverage the law of large numbers: improve accuracy via repeated independent trials.

#### Hashing for Distinct Elements (Improved):

- + Let  $\mathbf{h}_1, \mathbf{h}_2, \dots, \mathbf{h}_k : U \to [0, 1]$  be random hash functions
- s := 1
- For  $i = 1, \ldots, n$ 
  - $s := \min(s, h(x_i))$
- Return  $\widehat{d} = \frac{1}{s} 1$

Leverage the law of large numbers: improve accuracy via repeated independent trials.

#### Hashing for Distinct Elements (Improved):

- + Let  $h_1,h_2,\ldots,h_k: \mathcal{U} \to [0,1]$  be random hash functions
- $\cdot \ \mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$
- For i = 1, ..., n
  - $s := \min(s, h(x_i))$
- Return  $\widehat{d} = \frac{1}{s} 1$

Leverage the law of large numbers: improve accuracy via repeated independent trials.

#### Hashing for Distinct Elements (Improved):

+ Let  $h_1,h_2,\ldots,h_k: \mathcal{U} \to [0,1]$  be random hash functions

• 
$$s_1, s_2, ..., s_k := 1$$

• For  $i = 1, \ldots, n$ 

• For j=1,...,k, 
$$\mathbf{s}_j := \min(\mathbf{s}_j, \mathbf{h}_j(x_i))$$

• Return  $\hat{\mathbf{d}} = \frac{1}{s} - 1$ 

Leverage the law of large numbers: improve accuracy via repeated independent trials.

#### Hashing for Distinct Elements (Improved):

• Let  $\mathbf{h}_1, \mathbf{h}_2, \dots, \mathbf{h}_k : U \to [0, 1]$  be random hash functions

• 
$$\mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$$

- For i = 1, ..., n
  - For j=1,...,k,  $s_i := \min(s_i, h_i(x_i))$
- $\mathbf{S} := \frac{1}{k} \sum_{j=1}^{k} \mathbf{S}_{j}$ • Return  $\hat{\mathbf{d}} = \frac{1}{5} - 1$

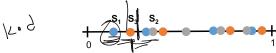
Leverage the law of large numbers: improve accuracy via repeated independent trials.

Hashing for Distinct Elements (Improved):

+ Let  $h_1,h_2,\ldots,h_k: \mathcal{U} \to [0,1]$  be random hash functions

• 
$$\mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$$

- For i = 1, ..., n
  - For j=1,...,k,  $\mathbf{s}_j := \min(\mathbf{s}_j, \mathbf{h}_j(\mathbf{x}_i))$
- $\mathbf{s} := \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}$  Return  $\hat{\mathbf{d}} = \frac{1}{s} 1$   $\frac{1}{2^{j}} = \mathbb{E}s$



$$\underline{\mathbf{s}} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}.$$
 Have already shown that for  $j = 1, ..., k$ :  
$$\underbrace{\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1}}_{\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}}}$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}. \text{ Have already shown that for } j = 1, \dots, k:$$
$$\underbrace{\mathbb{E}[\mathbf{s}_{j}]}_{\text{Var}[\mathbf{s}_{j}]} = \frac{1}{d+1} \implies \underbrace{\mathbb{E}[\mathbf{s}]}_{\text{Tabular}} \underbrace{\mathbb{E}[\mathbf{s}]}_{\text{Tabular}} = \frac{1}{d+1}$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}$$
. Have already shown that for  $j = 1, ..., k$ :  
 $\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1}$  (linearity of expectation)  
 $\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}}$ 

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}.$$
 Have already shown that for  $j = 1, ..., k$ :  

$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$

$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \cong \frac{1}{k(d+1)^{2}}$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}.$$
 Have already shown that for  $j = 1, ..., k$ :  

$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$

$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)}$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}. \text{ Have already shown that for } j = 1, \dots, k:$$
$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$
$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)}$$

Chebyshev Inequality:

$$\Pr\left[\underline{|\mathbf{s} - \underline{\mathbb{E}[\mathbf{s}]}|} \ge \underline{\epsilon \mathbb{E}[\mathbf{s}]}\right] \le \frac{V_{ar}[\mathbf{s}]}{(\epsilon \mathbb{E}[\mathbf{s}])^2}$$

$$\begin{split} \mathbf{s} &= \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}. \text{ Have already shown that for } j = 1, \dots, k: \\ \mathbf{v}_{\mathbf{s}'}(\mathbf{s}') \stackrel{!}{\xrightarrow{}}_{\mathbf{k}'} \cdot \sum \mathbf{v}_{\mathbf{s}'}(\mathbf{s}') \stackrel{!}{\xrightarrow{}}_{\mathbf{k}'} \underbrace{\mathbf{k}} \cdot \mathbf{v}_{\mathbf{s}'}(\mathbf{s}') \\ & \mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)} \\ & \text{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \text{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)} \end{split}$$

Chebyshev Inequality:

$$\Pr\left[|\mathbf{s} - \mathbb{E}[\mathbf{s}]| \ge \epsilon \mathbb{E}[\mathbf{s}]\right] \le \frac{\operatorname{Var}[\mathbf{s}]}{(\epsilon \mathbb{E}[\mathbf{s}])^2} = \frac{\mathbb{E}[\mathbf{s}]^2/k}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{k \cdot \epsilon^2}$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}.$$
 Have already shown that for  $j = 1, ..., k$ :  

$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$

$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)}$$

Chebyshev Inequality:  

$$\Pr[|\mathbf{s} - \mathbb{E}[\mathbf{s}]| \ge \epsilon \mathbb{E}[\mathbf{s}]] \le \frac{\operatorname{Var}[\mathbf{s}]}{(\epsilon \mathbb{E}[\mathbf{s}])^2} = \frac{\mathbb{E}[\mathbf{s}]^2/k}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{\underline{k} \cdot \epsilon^2} \le \delta$$
How should we set k if we want an error with probability at most  $\delta$ ?  

$$\ker \mathbb{E}[\mathbf{s}] = \frac{1}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{\frac{k}{\epsilon^2} \cdot \epsilon^2} = \frac{1}{\frac{k}{\epsilon^2} \cdot \epsilon^2} = \frac{1}{\frac{k}{\epsilon^2} \cdot \epsilon^2} = \frac{1}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{\frac{k}{\epsilon^2} \cdot \epsilon^2} = \frac{1}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{\frac{k}{\epsilon^2} \cdot \epsilon^2} = \frac{1}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1$$

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}. \text{ Have already shown that for } j = 1, \dots, k:$$
$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$
$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)}$$

Chebyshev Inequality:

$$\Pr\left[|\mathbf{S} - \mathbb{E}[\mathbf{S}]| \ge \epsilon \mathbb{E}[\mathbf{S}]\right] \le \frac{\operatorname{Var}[\mathbf{S}]}{(\epsilon \mathbb{E}[\mathbf{S}])^2} = \frac{\mathbb{E}[\mathbf{S}]^2/k}{\epsilon^2 \mathbb{E}[\mathbf{S}]^2} = \frac{1}{k \cdot \epsilon^2}$$

How should we set k if we want an error with probability at most  $\delta$ ?  $k = \frac{1}{\epsilon^2 \cdot \delta}.$ 

$$\mathbf{s} = \frac{1}{k} \sum_{j=1}^{k} \mathbf{s}_{j}. \text{ Have already shown that for } j = 1, \dots, k:$$
$$\mathbb{E}[\mathbf{s}_{j}] = \frac{1}{d+1} \implies \mathbb{E}[\mathbf{s}] = \frac{1}{d+1} \text{ (linearity of expectation)}$$
$$\operatorname{Var}[\mathbf{s}_{j}] \leq \frac{1}{(d+1)^{2}} \implies \operatorname{Var}[\mathbf{s}] \leq \frac{1}{k \cdot (d+1)^{2}} \text{ (linearity of variance)}$$

Chebyshev Inequality:

$$\Pr\left[|\mathbf{s} - \mathbb{E}[\mathbf{s}]| \ge \epsilon \mathbb{E}[\mathbf{s}]\right] \le \frac{\operatorname{Var}[\mathbf{s}]}{(\epsilon \mathbb{E}[\mathbf{s}])^2} = \frac{\mathbb{E}[\mathbf{s}]^2/k}{\epsilon^2 \mathbb{E}[\mathbf{s}]^2} = \frac{1}{k \cdot \epsilon^2} = \frac{\epsilon^2 \cdot \delta}{\epsilon^2} = \delta.$$

How should we set k if we want an error with probability at most  $\delta$ ?  $k = \frac{1}{\epsilon^2 \cdot \delta}.$ 

## Space Complexity

Hashing for Distinct Elements:

- + Let  $h_1,h_2,\ldots,h_{\it k}: {\it U} \rightarrow [0,1]$  be random hash functions
- $\cdot \ \mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$
- For  $i = 1, \ldots, n$ 
  - For j=1,..., k,  $\mathbf{s}_j := \min(\mathbf{s}_j, \mathbf{h}_j(x_i))$
- $\mathbf{S} := \frac{1}{k} \sum_{j=1}^{k} \mathbf{S}_{j}$
- Return  $\widehat{\mathbf{d}} = \frac{1}{s} 1$





• Setting  $k = \frac{1}{\epsilon^{2} \cdot \delta}$ , algorithm returns  $\hat{\mathbf{d}}$  with  $|d - \hat{\mathbf{d}}| \le 4\epsilon \cdot d$  with probability at least  $1 - \delta$ .

## Space Complexity

Hashing for Distinct Elements:

- + Let  $h_1,h_2,\ldots,h_{\it k}:U\rightarrow [0,1]$  be random hash functions
- $\cdot \ \mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$
- For  $i = 1, \ldots, n$ 
  - For j=1,..., k,  $\mathbf{s}_j := \min(\mathbf{s}_j, \mathbf{h}_j(x_i))$
- $\mathbf{S} := \frac{1}{k} \sum_{j=1}^{k} \mathbf{S}_{j}$
- Return  $\widehat{\mathbf{d}} = \frac{1}{s} 1$



• Setting  $k = \frac{1}{\epsilon^2 \cdot \delta}$ , algorithm returns  $\widehat{\mathbf{d}}$  with  $|d - \widehat{\mathbf{d}}| \le 4\epsilon \cdot d$  with probability at least  $1 - \delta$ .

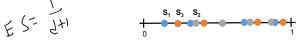
• Space complexity is  $k = \frac{1}{\epsilon^2 \cdot \delta}$  real numbers  $s_1, \ldots, s_k$ .

$$\varepsilon = d = 01$$
  $\frac{1}{101^3} = 1001$   $\frac{1}{101^2} = 5 = 50000$ 

# Space Complexity

Hashing for Distinct Elements:

- + Let  $h_1,h_2,\ldots,h_{\it k}: {\it U} \rightarrow [0,1]$  be random hash functions
- $\cdot \ \mathbf{s}_1, \mathbf{s}_2, \dots, \mathbf{s}_k := 1$
- For  $i = 1, \ldots, n$ 
  - For j=1,..., k,  $s_j := \min(s_j, h_j(x_i))$
- $\mathbf{S} := \frac{1}{k} \sum_{j=1}^{k} \mathbf{S}_{j}$
- Return  $\widehat{\mathbf{d}} = \frac{1}{s} 1$



- Setting  $k = \frac{1}{\epsilon^2 \cdot \delta}$ , algorithm returns  $\widehat{\mathbf{d}}$  with  $|d \widehat{\mathbf{d}}| \le 4\epsilon \cdot d$  with probability at least  $1 \delta$ .
- Space complexity is  $k = \frac{1}{\epsilon^2 \cdot \delta}$  real numbers  $s_1, \ldots, s_k$ .
- $\delta = 5\%$  failure rate gives a factor 20 overhead in space complexity.

How can we improve our dependence on the failure rate  $\delta$ ?

How can we improve our dependence on the failure rate  $\delta$ ? The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5$  – each using  $k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.

How can we improve our dependence on the failure rate  $\delta$ ?

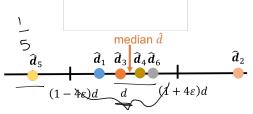
The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5$  – each using  $k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.

• Letting  $\hat{\underline{d}}_1, \dots, \hat{\underline{d}}_t$  be the outcomes of the *t* trials, return  $\hat{d} = median(\hat{d}_1, \dots, \hat{d}_t)$ .

How can we improve our dependence on the failure rate  $\delta$ ?

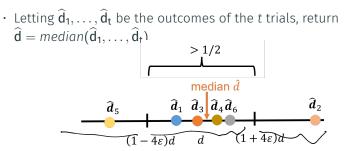
The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5 - \text{each using } k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.

· Letting  $\hat{d}_1, \ldots, \hat{d}_t$  be the outcomes of the *t* trials, return  $\hat{d} = median(\hat{d}_1, \ldots, \hat{d}_t)$ 



How can we improve our dependence on the failure rate  $\delta$ ?

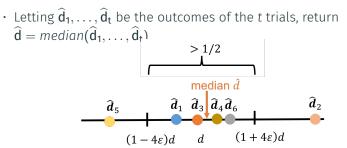
The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5$  – each using  $k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.



• If > 1/2 of trials fall in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ , then the median will.

How can we improve our dependence on the failure rate  $\delta$ ?

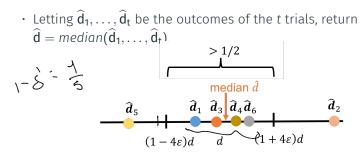
The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5$  – each using  $k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.



- If > 1/2 of trials fall in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$ , then the median will.
- Have < 1/2 of trials on both the left and right.

How can we improve our dependence on the failure rate  $\delta$ ?

The median trick: Run  $t = O(\log 1/\delta)$  trials each with failure probability  $\delta' = 1/5$  – each using  $k = \frac{1}{\delta'\epsilon^2} = \frac{5}{\epsilon^2}$  hash functions.



- If  $\geq 2/3$  of trials fall in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$ , then the median will.
- Have < 1/3 of trials on both the left and right.

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{d} = \textit{median}(\widehat{d}_1, \dots, \widehat{d}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{d} = \textit{median}(\widehat{d}_1, \dots, \widehat{d}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]?$ 

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{d} = \textit{median}(\widehat{d}_1, \dots, \widehat{d}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let **X** be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .

$$\Pr\left(\widehat{\mathbf{d}} \notin [(1-4\epsilon)d, (1+4\epsilon)d]\right) \leq \Pr\left(\mathbf{X} < \frac{2}{3} \cdot t\right)$$

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{d} = \textit{median}(\widehat{d}_1, \dots, \widehat{d}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[X] =$ 

$$\Pr\left(\widehat{\mathbf{d}} \notin [(1-4\epsilon)d, (1+4\epsilon)d]\right) \leq \Pr\left(\mathbf{X} < \frac{2}{3} \cdot t\right)$$

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{\mathbf{d}} = median(\widehat{\mathbf{d}}_1, \dots, \widehat{\mathbf{d}}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\begin{aligned}
\mathbb{E}[X] &= \frac{4}{5} \cdot t. \\
\Pr\left(\widehat{d} \notin [(1 - 4\epsilon)d, (1 + 4\epsilon)d]\right) \leq \Pr\left(X < \frac{2}{3} \cdot t\right)
\end{aligned}$ 

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{\mathbf{d}} = median(\widehat{\mathbf{d}}_1, \dots, \widehat{\mathbf{d}}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[X] = \frac{4}{5} \cdot t. \qquad \underbrace{5}_{\mathcal{L}} \cdot \underbrace{\frac{4}{5}}_{\mathcal{L}}, \quad \underbrace{\frac{2}{5}}_{\mathcal{L}}$   $\underbrace{\Pr\left(\widehat{d} \notin [(1 - 4\epsilon)d, (1 + 4\epsilon)d]\right)}_{\mathcal{L}} \ge \Pr\left(X < \frac{5}{6} \cdot \mathbb{E}[X]\right)$ 

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{d} = \textit{median}(\widehat{d}_1, \dots, \widehat{d}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[\overline{X]} = \frac{4}{5} \cdot t.$ 

$$\Pr\left(\widehat{\mathsf{d}} \notin [(1-4\epsilon)d, (1+4\epsilon)d]\right) \leq \Pr\left(\mathsf{X} < \frac{5}{6} \cdot \mathbb{E}[\mathsf{X}]\right) \leq \Pr\left(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \geq \underbrace{\frac{1}{6}\mathbb{E}[\mathsf{X}]}\right)$$

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{\mathbf{d}} = median(\widehat{\mathbf{d}}_1, \dots, \widehat{\mathbf{d}}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[X] = \frac{4}{5} \cdot t.$ 

$$\Pr\left(\widehat{\mathsf{d}} \notin [(1-4\epsilon)d, (1+4\epsilon)d]\right) \leq \Pr\left(\mathsf{X} < \frac{5}{6} \cdot \mathbb{E}[\mathsf{X}]\right) \leq \Pr\left(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \geq \frac{1}{6}\mathbb{E}[\mathsf{X}]\right)$$

Apply Chernoff bound:

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least 4/5.
- $\cdot \ \widehat{\mathbf{d}} = median(\widehat{\mathbf{d}}_1, \dots, \widehat{\mathbf{d}}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]$ ?

• Let X be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[X] = \frac{4}{5} \cdot t. \qquad Pr\left(X < \frac{2}{5} + \frac{1}{5}\right)$   $\Pr\left(\widehat{d \notin}[(1 - 4\epsilon)d, (1 + 4\epsilon)d]\right) \leq \Pr\left(X < \frac{5}{6} \cdot \mathbb{E}[X]\right) \leq \Pr\left(|X - \mathbb{E}[X]| \geq \frac{1}{6}\mathbb{E}[X]\right)$ 

Apply Chernoff bound:

$$\Pr\left(|\mathbf{X} - \mathbb{E}[\mathbf{X}]| \ge \frac{1}{6}\mathbb{E}[\mathbf{X}]\right) \le \underbrace{2\exp\left(-\frac{1}{6}^2 \cdot \frac{4}{5}t\right)}_{2} = O\left(e^{-ct}\right).$$

- $\hat{\mathbf{d}}_1, \dots, \hat{\mathbf{d}}_t$  are the outcomes of the *t* trials, each falling in  $[(1-4\epsilon)d, (1+4\epsilon)d]$  with probability at least 4/5.
- $\cdot \hat{\mathbf{d}} = median(\hat{\mathbf{d}}_1, \ldots, \hat{\mathbf{d}}_t).$

What is the probability that the median  $\hat{\mathbf{d}}$  falls in  $[(1-4\epsilon)d, (1+4\epsilon)d]?$ 

• Let **X** be the # of trials falling in  $[(1 - 4\epsilon)d, (1 + 4\epsilon)d]$ .  $\mathbb{E}[\mathbf{X}] = \frac{4}{r} \cdot t.$ 

 $\Pr\left(\widehat{\mathsf{d}} \notin [(1-4\epsilon)d, (1+4\epsilon)d]\right) \leq \Pr\left(\mathsf{X} < \frac{5}{6} \cdot \mathbb{E}[\mathsf{X}]\right) \\ \leq \Pr\left(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \geq \frac{1}{6}\mathbb{E}[\mathsf{X}]\right)$ Sol = n-ct

Apply Chernoff bound:

$$\Pr\left(|\mathbf{X} - \mathbb{E}[\mathbf{X}]| \ge \frac{1}{6}\mathbb{E}[\mathbf{X}]\right) \le 2\exp\left(-\frac{\frac{1}{6}^2 \cdot \frac{4}{5}t}{2+1/6}\right) = \underbrace{O\left(e^{-ct}\right)}_{\mathbf{X} = \mathbf{X} = \mathbf$$

• Setting  $t = O(\log(1/\delta))$  gives failure probability  $e^{-\log(1/\delta)} = \delta$ .

10

**Upshot:** The median of  $t = O(\log(1/\delta))$  independent runs of the hashing algorithm for distinct elements returns  $\widehat{\mathbf{d}} \in [(1 - 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least  $1 - \delta$ .

**Upshot:** The median of  $t = O(\log(1/\delta))$  independent runs of the hashing algorithm for distinct elements returns  $\widehat{\mathbf{d}} \in [(1 - 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least  $1 - \delta$ .

Total Space Complexity: *t* trials, each using  $k = \frac{1}{\epsilon^2 \delta'}$  hash functions, for  $\delta' = 1/5$ . Space is  $\frac{5t}{\epsilon^2} = O\left(\frac{\log(1/\delta)}{\epsilon^2}\right)$  real numbers (the minimum value of each hash function).

**Upshot:** The median of  $t = O(\log(1/\delta))$  independent runs of the hashing algorithm for distinct elements returns  $\widehat{\mathbf{d}} \in [(1 - 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least  $1 - \delta$ .

**Total Space Complexity:** *t* trials, each using  $k = \frac{1}{\epsilon^2 \delta'}$  hash functions, for  $\delta' = 1/5$ . Space is  $\frac{5t}{\epsilon^2} = O\left(\frac{\log(1/\delta)}{\epsilon^2}\right)$  real numbers (the minimum value of each hash function).

No dependence on the number of distinct elements *d* or the number of items in the stream *n*! Both of these numbers are typically very large.

**Upshot:** The median of  $t = O(\log(1/\delta))$  independent runs of the hashing algorithm for distinct elements returns  $\widehat{\mathbf{d}} \in [(1 - 4\epsilon)d, (1 + 4\epsilon)d]$  with probability at least  $1 - \delta$ .

Total Space Complexity: *t* trials, each using  $k = \frac{1}{\epsilon^2 \delta'}$  hash functions, for  $\delta' = 1/5$ . Space is  $\frac{5t}{\epsilon^2} = O\left(\frac{\log(1/\delta)}{\epsilon^2}\right)$  real numbers (the minimum value of each hash function).

No dependence on the number of distinct elements *d* or the number of items in the stream *n*! Both of these numbers are typically very large.

A note on the median: The median is often used as a robust alternative to the mean, when there are outliers (e.g., heavy tailed distributions, corrupted data).

## **Distinct Elements in Practice**

Our algorithm uses continuous valued fully random hash functions.

# **Distinct Elements in Practice**

Our algorithm uses continuous valued fully random hash functions. Can't be implemented...

• The idea of using the minimum hash value of  $x_1, \ldots, x_n$  to estimate the number of distinct elements naturally extends to when the hash functions map to discrete values.

- The idea of using the minimum hash value of  $x_1, \ldots, x_n$  to estimate the number of distinct elements naturally extends to when the hash functions map to discrete values.
- Flajolet-Martin (LogLog) algorithm and HyperLogLog.

- The idea of using the minimum hash value of  $x_1, \ldots, x_n$  to estimate the number of distinct elements naturally extends to when the hash functions map to discrete values.
- Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	1
•	
h(x <sub>n</sub> )	1011000

- The idea of using the minimum hash value of  $x_1, \ldots, x_n$  to estimate the number of distinct elements naturally extends to when the hash functions map to discrete values.
- Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	10100 <u>10</u>
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	l I
•	1
h(x <sub>n</sub> )	101 <u>1000</u>

Estimate # distinct elements based on maximum number of trailing zeros **m**.

- The idea of using the minimum hash value of  $x_1, \ldots, x_n$  to estimate the number of distinct elements naturally extends to when the hash functions map to discrete values.
- Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	1
•	
h(x <sub>n</sub> )	1011000

Estimate # distinct elements based on maximum number of trailing zeros **m**. The more distinct hashes we see, the higher we expect this maximum to be.

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
h(x <sub>3</sub> )	1001110
	:
•	•
h(x <sub>n</sub> )	1011 <mark>000</mark>

Estimate # distinct elements based on maximum number of trailing zeros **m**.

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

$h(x_1)$	101001 <mark>0</mark>
<b>h</b> (x <sub>2</sub> )	1001 <u>100</u>
<b>h</b> (x <sub>3</sub> )	1001110
•	'

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

a) O(1)  $b(O(\log d))$ c)  $O(\sqrt{d})$  d) O(d)

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

 $Pr(h(x_i) has x trailing zeros) =$ 

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	:
	•
h(x <sub>n</sub> )	

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$Pr(\mathbf{h}(x_i) \text{ has } x \text{ trailing zeros}) = \frac{1}{2^x}$$

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	:
	:

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

 $\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}}$ 

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010
<b>h</b> (x <sub>2</sub> )	1001100
<b>h</b> (x <sub>3</sub> )	1001110
	:
	:

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}} = \frac{1}{d}.$$

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

$h(x_1)$	101001 <mark>0</mark>
<b>h</b> (x <sub>2</sub> )	10011 <mark>00</mark>
<b>h</b> (x <sub>3</sub> )	1001110
•	•

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}} = \frac{1}{d}.$$

So with d distinct hashes, expect to see 1 with log d trailing zeros. Expect  $\underline{m} \approx \log d$ .

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

<b>h</b> (x <sub>1</sub> )	1010010	
<b>h</b> (x <sub>2</sub> )	1001100	
<b>h</b> (x <sub>3</sub> )	1001110	
-		
•		
h(x <sub>n</sub> )	1011000	

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}} = \frac{1}{d}.$$

So with *d* distinct hashes, expect to see 1 with log *d* trailing zeros. Expect  $\mathbf{m} \approx \log d$ . **m** takes log log *d* bits to store.

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

$h(x_1)$	101001 <mark>0</mark>
<b>h</b> (x <sub>2</sub> )	10011 <mark>00</mark>
<b>h</b> (x <sub>3</sub> )	1001110
•	

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}} = \frac{1}{d}.$$

So with *d* distinct hashes, expect to see 1 with log *d* trailing zeros. Expect  $\mathbf{m} \approx \log d$ . **m** takes log log *d* bits to store.

**Total Space:**  $O\left(\underbrace{\log \log d}{\epsilon^2}\right)$  for an  $\epsilon$  approximate count.

Flajolet-Martin (LogLog) algorithm and HyperLogLog.

$h(x_1)$	101001 <mark>0</mark>
<b>h</b> (x <sub>2</sub> )	10011 <mark>00</mark>
<b>h</b> (x <sub>3</sub> )	1001110
•	

Estimate # distinct elements based on maximum number of trailing zeros **m**.

With *d* distinct elements, roughly what do we expect **m** to be?

$$\Pr(\mathbf{h}(x_i) \text{ has } \log d \text{ trailing zeros}) = \frac{1}{2^{\log d}} = \frac{1}{d}.$$

So with *d* distinct hashes, expect to see 1 with log *d* trailing zeros. Expect  $\mathbf{m} \approx \log d$ . **m** takes log log *d* bits to store.

**Total Space:**  $O\left(\frac{\log \log d}{\epsilon^2}\right)$  for an  $\epsilon$  approximate count.

Note: Careful averaging of estimates from multiple hash functions.

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used 
$$= O\left(\frac{\log \log d}{\epsilon^2}\right)$$

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>  
=  $\frac{1.04 \cdot 5}{.02^2}$  = 13000 bits  $\approx 1.6 \text{ kB}$ 

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>  
=  $\frac{1.04 \cdot 5}{.02^2}$  = 13000 bits  $\approx 1.6 \ kB$ 

**Mergeable Sketch:** Consider the case (essentially always in practice) that the items are processed on different machines.

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>  
=  $\frac{1.04 \cdot 5}{.02^2}$  = 13000 bits  $\approx 1.6 \ kB$ 

**Mergeable Sketch:** Consider the case (essentially always in practice) that the items are processed on different machines.

• Given data structures (sketches)  $HLL(x_1, \ldots, x_n)$ ,  $HLL(y_1, \ldots, y_n)$  is is easy to merge them to give  $HLL(x_1, \ldots, x_n, y_1, \ldots, y_n)$ .

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>  
=  $\frac{1.04 \cdot 5}{.02^2}$  = 13000 bits  $\approx 1.6 \ kB$ 

**Mergeable Sketch:** Consider the case (essentially always in practice) that the items are processed on different machines.

• Given data structures (sketches)  $HLL(x_1, ..., x_n)$ ,  $HLL(y_1, ..., y_n)$  is is easy to merge them to give  $HLL(x_1, ..., x_n, y_1, ..., y_n)$ . How?

Using HyperLogLog to count 1 billion distinct items with 2% accuracy:

space used = 
$$O\left(\frac{\log \log d}{\epsilon^2}\right)$$
  
=  $\frac{1.04 \cdot \lceil \log_2 \log_2 d \rceil}{\epsilon^2}$  bits<sup>1</sup>  
=  $\frac{1.04 \cdot 5}{.02^2}$  = 13000 bits  $\approx 1.6 \ kB$ 

**Mergeable Sketch:** Consider the case (essentially always in practice) that the items are processed on different machines.

- Given data structures (sketches)  $HLL(x_1, ..., x_n)$ ,  $HLL(y_1, ..., y_n)$ is is easy to merge them to give  $HLL(x_1, ..., x_n, y_1, ..., y_n)$ . How?
- Set the maximum *#* of trailing zeros to the maximum in the two sketches.
- 1. 1.04 is the constant in the HyperLogLog analysis. Not important!

Use Case: Exploratory SQL-like queries on tables with 100s billions of rows.  $\sim 5$  million count distinct queries per day.

Use Case: Exploratory SQL-like queries on tables with 100s billions of rows.  $\sim 5$  million count distinct queries per day. E.g.,

- Count number if distinct users in Germany that made at least one search containing the word 'auto' in the last month.
- Count number of distinct subject lines in emails sent by users that have registered in the last week, in comparison to number of emails sent overall (to estimate rates of spam accounts).

Use Case: Exploratory SQL-like queries on tables with 100s billions of rows.  $\sim 5$  million count distinct queries per day. E.g.,

- Count number if distinct users in Germany that made at least one search containing the word 'auto' in the last month.
- Count number of distinct subject lines in emails sent by users that have registered in the last week, in comparison to number of emails sent overall (to estimate rates of spam accounts).

Traditional *COUNT*, *DISTINCT* SQL calls are far too slow, especially when the data is distributed across many servers.

Questions on distinct elements counting?