COMPSCI 514: Algorithms for Data Science

Cameron Musco University of Massachusetts Amherst. Fall 2022. Lecture 6

- Problem Set 1 is due tomorrow at 11:59pm in Gradescope.
- Quiz 3 is due Monday at 8pm.

Last Time

Last Class:

- Higher moment bounds and exponential concentration bounds
- Bernstein inequality

This Class:

- Connection between exponential concentration bounds and the central limit theorem.
- The Chernoff bound.
- Bloom filters: random hashing to maintain a large set in small space.

Interpretation as a Central Limit Theorem

Bernstein Inequality (Simplified): Consider independent random variables X_{1}, \ldots, X_{n} falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_{i}]$, $\sigma^{2} = \operatorname{Var}[\sum X_{i}]$, and $s \leq \sigma$. Then:

$$\left[\Pr\left(\left| \sum_{i=1}^{n} X_{i} - \mu \right| \geq s\sigma \right) \leq 2 \exp\left(-\frac{s^{2}}{4}\right). \right]$$

Can plot this bound for different s:



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 has density $\underline{p(s\sigma)} = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{s^2}{2}}$.

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Exercise: Using this can show that for $X \sim \mathcal{N}(0, \sigma^2)$: for any $s \ge 0$,

$$\Pr\left(|\mathbf{X}| \ge s \cdot \sigma\right) \le 2e^{-\frac{s^2}{2}}.$$

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Central Limit Theorem Interpretation: Bernstein's inequality gives a quantitative version of the CLT. The distribution of the sum of *bounded* independent random variables can be upper bounded with a Gaussian (normal) distribution.



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Stronger Central Limit Theorem: The distribution of the sum of *n bounded* independent random variables converges to a Gaussian (normal) distribution as *n* goes to infinity.



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- Why is the Gaussian distribution is so important in statistics, science, ML, etc.?
- Many random variables can be approximated as the sum of a large number of small and roughly independent random effects. Thus, their distribution looks Gaussian by CLT.

A useful variation of the Bernstein inequality for binary (indicator) random variables is:

Chernoff Bound (simplified version): Consider independent random variables X_1, \ldots, X_n taking values in $\{0, 1\}$. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$. For any $\overline{\delta \ge 0}$ $\mathbb{E}[\mathbb{E}[X_i] = \mathbb{E}[X_i] =$ A useful variation of the Bernstein inequality for binary (indicator) random variables is:

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As δ gets larger and larger, the bound falls of exponentially fast.

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What will be the maximum number of items hashed into the same location?

Let $\underline{S_i}$ be the number of items hashed into position *i* and $\underline{S_{i,j}}$ be 1 if x_j is hashed into bucket *i* ($h(x_j) = i$) and 0 otherwise.

m: total number of items hashed and size of hash table. x_1, \ldots, x_m : the items. h: random hash function mapping $x_1, \ldots, x_m \rightarrow [m]$. Let S_i be the number of items hashed into position *i* and $S_{i,j}$ be 1 if x_j is hashed into bucket *i* ($h(x_j) = i$) and 0 otherwise.

$$\underbrace{\mathbb{E}[\mathbf{S}_i]}_{j=1} = \sum_{j=1}^m \mathbb{E}[\mathbf{S}_{i,j}] = \underline{m} \cdot \frac{1}{\underline{m}} = \underline{1}$$

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$$\mathbb{E}[\underline{\mathbf{S}}_{i}] = \sum_{j=1}^{m} \mathbb{E}[\underline{\mathbf{S}}_{i,j}] = m \cdot \frac{1}{m} = 1 = \mu.$$
By the Chernoff Bound: for any $\underline{\delta \ge 0}$,
$$\Pr(\underline{\mathbf{S}}_{i} \underbrace{1} + \delta) \le \Pr\left(\left|\sum_{j=1}^{m} \mathbf{S}_{i,j} - 1\right| \ge \delta \cdot \mu\right) \le 2\exp\left(-\frac{\delta^{2}}{2 + \delta}\right)$$

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$$\Pr(\mathbf{S}_i \ge 1 + \delta) \le \Pr\left(\left|\sum_{i=1}^n \mathbf{S}_{i,j} - 1\right| \ge \delta\right) \le 2 \exp\left(-\frac{\delta^2}{2 + \delta}\right).$$

$$\frac{2\log_{10} + 1}{\Pr(S_{i} \ge 1 + \delta)} \le \Pr\left(\left|\sum_{i=1}^{n} S_{i,j} - 1\right| \ge \delta\right) \le 2\exp\left(-\frac{\delta^{2}}{2 + \delta}\right).$$

Set $\delta = 20 \log m$. Gives:

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$$\Pr(\underline{\mathbf{S}_i} \ge 20 \log m + 1) \le 2 \exp\left(-\frac{(20 \log m)^2}{2 + 20 \log m}\right)$$

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$$\frac{\operatorname{Set} \delta = 20 \log m. \text{ Gives:}}{\Pr(\mathbf{S}_{i} \ge 20 \log m + 1)} \le 2 \exp\left(-\frac{(20 \log m)^{2}}{2 + 20 \log m}\right) \le \frac{\operatorname{Per}(-18 \log m)}{2 \exp(-18 \log m)} \le \frac{2}{m^{18}}.$$

$$\frac{10 \operatorname{Per}(-20 \log m)}{2 \operatorname{Per}(-20 \log m)} = 2 \exp\left(-\frac{20^{2} \operatorname{Per}(-20 \log m)}{22}\right) = 2 \exp\left(-\frac{20^{2} \operatorname{Per}(-20 \log m)}{22}\right)$$

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Apply Union Bound:

$$\Pr(\max_{i \in [m]} \mathbf{S}_i \ge 20 \log m + 1) = \Pr\left(\bigcup_{i=1}^m (\mathbf{S}_i \ge 20 \log m + 1)\right)$$

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Using Chebyshev's inequality could only show the maximum load is bounded by $O(\sqrt{m})$ with good probability (good exercise).

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- So, even with a simple linked list to store the items in each bucket, worst case query time is $O(\log m)$.
- Using Chebyshev's inequality could only show the maximum load is bounded by $O(\sqrt{m})$ with good probability (good exercise).
- The Chebyshev bound holds even with a pairwise independent hash function. The stronger Chernoff-based bound can be shown to hold with a *k*-wise independent hash function for $k = O(\log m)$.

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Solution: Bloom filters (repeated random hashing). Will use much less space than a hash table.

Chose k independent random hash functions $\mathbf{h}_1, \dots, \mathbf{h}_k$ mapping the universe of elements $U \to [m]$.

 $\sqrt{Maintain an array A containing m bits, all initially 0.}$

- *insert*(*x*): set all bits $A[h_1(x)] = ... = A[h_k(x)] := 1$.
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- When url x comes in, if query(x) = 1, cache the page at x. If not, run *insert*(x) so that if it comes in again, it will be cached.
- False positive: A new url (possible one-hit-wonder) is cached. If the bloom filter has a false positive rate of $\delta = .05$, the number of cached one-hit-wonders will be reduced by at least 95%.

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Movies

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 When a new rating is inserted for (user_x, movie_y), add (user_x, movie_y) to a bloom filter.

• Before reading (*user_x, movie_y*) (possibly via an out of memory access), check the bloom filter, which is stored in memory.

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- Before reading (*user_x*, *movie_y*) (possibly via an out of memory access), check the bloom filter, which is stored in memory.
- False positive: A read is made to a possibly empty cell. A $\delta = .05$ false positive rate gives a <u>95</u>% reduction in these empty reads.

More Applications

- **Database Joins:** Quickly eliminate most keys in one column that don't correspond to keys in another.
- **Recommendation systems:** Bloom filters are used to prevent showing users the same recommendations twice.
- Spam/Fraud Detection:
 - Bit.ly and Google Chrome use bloom filters to quickly check if a url maps to a flagged site and prevent a user from following it.
 - Can be used to detect repeat clicks on the same ad from a single IP-address, which may be the result of fraud.
- **Digital Currency:** Some Bitcoin clients use bloom filters to quickly pare down the full transaction log to transactions involving bitcoin addresses that are relevant to them (SPV: simplified payment verification).

For a bloom filter with m bits and k hash functions, the insertion and query time is O(k).

For a bloom filter with *m* bits and *k* hash functions, the insertion and query time is O(k). How does the false positive rate δ depend on *m*, *k*, and the number of items inserted?

Step 1: What is the probability that after inserting *n* elements, the *i*th bit of the array A is still 0? $n \times k$ total hashes must not hit bit *i*.

$$\Pr(\underline{A[i]=0)} = \Pr\left(\underbrace{\mathbf{h}_{1}(x_{1}) \neq i \cap \ldots \cap \mathbf{h}_{k}(x_{k}) \neq i}_{\cap \underline{\mathbf{h}_{1}(x_{2}) \neq i \ldots \cap \mathbf{h}_{k}(x_{2}) \neq i \cap \ldots}\right)$$

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$$\cap h_1(x_2) \neq i \ldots \cap h_k(x_2) \neq i \cap \ldots)$$
$$= \underbrace{Pr(h_1(x_1) \neq i) \times \ldots \times Pr(h_k(x_1) \neq i) \times Pr(h_1(x_2) \neq i) \ldots}_{i \in \mathbb{N}}$$

 $k \cdot n$ events each occuring with probability 1 - 1/m

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$$Pr(A[h_1(w)] = \ldots = A[h_k(w)] = 1)$$
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=
$$\left(1 - e^{-\frac{kn}{m}}\right)^k$$
Analysis

How does the false positive rate δ depend on *m*, *k*, and the number of items inserted?

Step 1: What is the probability that after inserting *n* elements, the *i*th bit of the array A is still 0?

$$\Pr(A[i] = 0) = \left(1 - \frac{1}{m}\right)^{kn} \approx e^{-\frac{kn}{m}}$$

Step 2: What is the probability that querying a new item *w* gives a false positive?

$$Pr(A[\mathbf{h}_{1}(w)] = \dots = A[\mathbf{h}_{k}(w)] = 1)$$

=
$$Pr(A[\mathbf{h}_{1}(w)] = 1) \times \dots \times Pr(A[\mathbf{h}_{k}(w)] = 1)$$

=
$$\left(1 - e^{-\frac{kn}{m}}\right)^{k}$$
 Actually Incorrect!

n: total number items in filter, *m*: number of bits in filter, *k*: number of random hash functions, h_1, \ldots, h_k : hash functions, *A*: bit array, δ : false positive rate.

Analysis

How does the false positive rate δ depend on *m*, *k*, and the number of items inserted?

Step 1: What is the probability that after inserting *n* elements, the *i*th bit of the array A is still 0?

$$\Pr(A[i] = 0) = \left(1 - \frac{1}{m}\right)^{kn} \approx e^{-\frac{kn}{m}}$$

Step 2: What is the probability that querying a new item *w* gives a false positive?

$$\begin{aligned} \Pr\left(A[\mathbf{h}_1(w)] &= \dots = A[\mathbf{h}_k(w)] = 1\right) \\ &= \Pr(A[\mathbf{h}_1(w)] = 1) \times \dots \times \Pr(A[\mathbf{h}_k(w)] = 1) \\ &= \left(1 - e^{-\frac{kn}{m}}\right)^k \quad \text{Actually Incorrect! Dependent events.} \end{aligned}$$

n: total number items in filter, *m*: number of bits in filter, *k*: number of random hash functions, h_1, \ldots, h_k : hash functions, *A*: bit array, δ : false positive rate.