COMPSCI 650 Applied Information Theory	Mar 8, 2016
Lecture 1	
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Today we will spend sometime discussing about the upcoming midterm. Hence the length of the lecture is smaller than usual.

## 1 Secret Sharing Schemes

This is an example that we discuss in the class.

Suppose a secret S has to be divided between two users. The share of user 1 is  $X_1$ , and share of user 2 is  $X_2$ . We must have the following two properties:

- 1.  $H(S|X_1) = H(S|X_2) = H(S)$ .
- 2.  $H(S|X_1, X_2) = 0.$

Let  $S \in \mathcal{X} \equiv \{0, \dots, q-1\}$ . Then choose  $X_1$  randomly and uniformly from  $\mathcal{X}$ . Let  $X_2 = (X_1 + S)$  mod q. This scheme clearly satisfies the above two properties.

Question: Can you think of a generalization?

# 2 Parameter Estimation

Consider an indexed family of distributions  $\{f(x; \theta)\}$ . X is the underlying random variable with sample space  $\mathcal{X}$ . That is we have,

- $f(x;\theta) \ge 0.$
- $\int f(x;\theta)dx = 1.$

Here  $\theta \in \Theta$ , the parameter set.

An estimator is a mapping

 $T: \mathcal{X}^n \to \Theta.$ 

The estimation error is

$$E_{\theta}(T(X_1,\ldots,X_n)-\theta),$$

where the  $\theta$  in the subscript denote that the expect ion is taken with respect to  $f(x; \theta)$ . The estimator is called *unbiased* if the estimation error is 0 for all  $\theta \in \Theta$ . Note that we can ask for more out of an estimator; such as  $P_{\theta}(|T(X_1, \ldots, X_n) - \theta| > \epsilon) \to 0$  for any  $\epsilon > 0$ . An estimator  $T_1$  dominates another estimator  $T_2$  if for all  $\theta$ ,

$$E_{\theta}(T_1(X_1,\ldots,X_n)-\theta)^2 \le E_{\theta}(T_2(X_1,\ldots,X_n)-\theta)^2.$$

This begs the question: what is the best estimator?

#### 2.1 Score function

The score function is:

$$V(X) = \frac{\partial}{\partial \theta} \ln f(X;\theta) = \frac{\frac{\partial}{\partial \theta} f(X;\theta)}{f(X;\theta)}$$

If  $X_1, \ldots, X_n$  are i.i.d.  $f(x; \theta)$ , then:

$$V(X_1, \dots, X_n) = \frac{\partial}{\partial \theta} \ln f(X_1, \dots, X_n; \theta) = \frac{\partial}{\partial \theta} \sum_i \ln f(X_i; \theta) = \sum_i V(X_i).$$

Also, we must have,

$$E_{\theta}V = \int f(x;\theta) \frac{\frac{\partial}{\partial \theta} f(x;\theta)}{f(x;\theta)} dx = \frac{\partial}{\partial \theta} \int f(x;\theta) dx = \frac{\partial}{\partial \theta} 1 = 0.$$

Hence,

 $\mathbf{If}$ 

$$\operatorname{Var}_{\theta}(V) = E_{\theta}V^2.$$

### 2.2 Fisher Information

How to quantify the amount of information about  $\theta$  that is present in the data? Define,  $J(\theta) = \operatorname{Var}_{\theta}(V) = E_{\theta}V^{2}.$ 

$$X_1, \ldots, X_n$$
 are i.i.d.  $f(x; \theta)$ , then  $\operatorname{Var}_{\theta}(\sum_i V(X_i)) = \sum_i \operatorname{Var}_{\theta}(V(X_i))$ . Hence,  
 $J_n(\theta) = nJ(\theta)$ .

#### 2.3 Cramer-Rao bound

**Theorem 1** The mean-square error of any unbiased estimator T is

$$\operatorname{Var}_{\theta}(T) \ge \frac{1}{J(\theta)}$$

**Proof** Just an application of the Cauchy-Schwartz inequality. Note,

 $E_{\theta}T = \theta$ 

and

$$E_{\theta}V = 0.$$

Hence,

$$E_{\theta}\Big((V - E_{\theta}V)(T - E_{\theta}T)\Big) = E_{\theta}\Big(VT - V\theta\Big)$$
$$= E_{\theta}\Big(VT\Big)$$
$$= \int f(x;\theta)\frac{\frac{\partial}{\partial\theta}f(x;\theta)}{f(x;\theta)}T(x)dx$$
$$= \frac{\partial}{\partial\theta}\int f(x;\theta)T(x)dx$$
$$= \frac{\partial}{\partial\theta}\theta$$
$$= 1.$$

But from Cauchy-Schwartz,

$$1 = \left( E_{\theta} \Big( (V - E_{\theta} V) (T - E_{\theta} T) \Big) \right)^2 \le \operatorname{Var}_{\theta} V \cdot \operatorname{Var}_{\theta} (T) = J(\theta) \cdot \operatorname{Var}_{\theta} (T).$$

Example: Consider  $X_1, \ldots, X_n$  all i.i.d. Gaussian  $\mathcal{N}(\theta, \sigma^2)$ .

$$V(X) = \frac{\partial}{\partial \theta} \ln \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(X-\theta)^2}{2\sigma^2}} = \frac{X-\theta}{\sigma^2}.$$

Therefore,

$$J(\theta) = \operatorname{Var}_{\theta} V = \frac{\sigma^2}{\sigma^4} = \frac{1}{\sigma^2}.$$

SInce,  $J_n(\theta) = nJ(\theta)$ , from Cramer-Rao,

$$\operatorname{Var}_{\theta}(T) \ge \frac{\sigma^2}{n}.$$

Now consider the estimator

$$T(X_1,\ldots,X_n) = \frac{1}{n}\sum_i X_i.$$

We have,

$$E_{\theta}(T-\theta)^2 = \operatorname{Var}_{\theta}\left(\frac{1}{n}\sum_{i}X_i\right) = \frac{1}{n}\sum_{i}\operatorname{Var}_{\theta}(X_i) = \frac{\sigma^2}{n}.$$

So we can achieve the Cramer-Rao bound!