

COMPSCI 614: Randomized Algorithms with Applications to Data Science

Prof. Cameron Musco

University of Massachusetts Amherst. Spring 2024.

Lecture 20

Summary

Last Time: Markov Chain Fundamentals

- The gambler's ruin problem.
- Aperiodicity and stationary distribution of a Markov chain.
- The fundamental theorem of Markov chains.
- Example of a uniform stationary distribution for a symmetric Markov chain (shuffling).

Last Time: Markov Chain Fundamentals

- The gambler's ruin problem.
- Aperiodicity and stationary distribution of a Markov chain.
- The fundamental theorem of Markov chains.
- Example of a uniform stationary distribution for a symmetric Markov chain (shuffling).

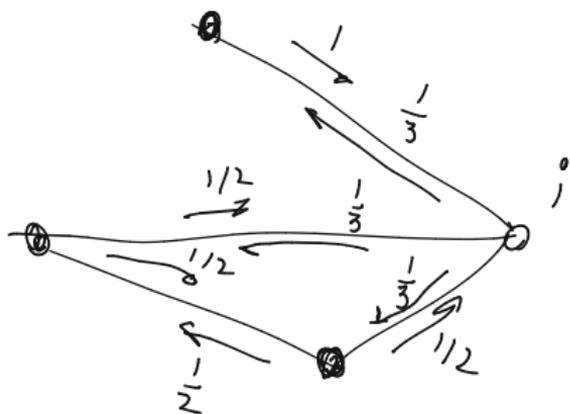
Today: Mixing Time Analysis

- How quickly does a Markov chain actually converge to its stationary distribution?
- Mixing time and its analysis via coupling.

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

- What is the state space of this chain? nodes of graph
- What is the transition probability $P_{i,j}$? $\frac{1}{d_i}$

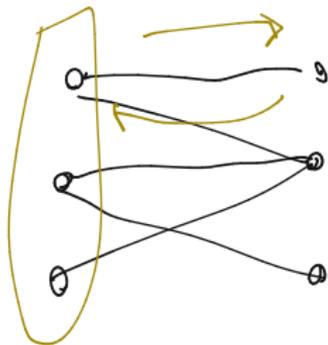


$$P = \begin{bmatrix} 0 & \frac{1}{d_1} & 0 & 0 \\ \frac{1}{d_1} & 0 & 0 & 0 \\ 0 & 0 & \frac{1}{d_1} & 0 \\ 0 & 0 & 0 & \frac{1}{d_1} \end{bmatrix}$$

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

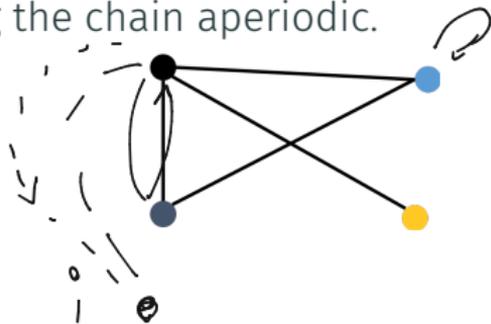
- What is the state space of this chain?
- What is the transition probability $P_{i,j}$?
- Is this chain aperiodic?



Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

- What is the state space of this chain?
- What is the transition probability $P_{i,j}$?
- Is this chain aperiodic?
- If the graph is not bipartite, then there is at least one odd cycle, making the chain aperiodic.



Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi(i) = \pi P_{:,i} = \sum_j \pi(j) P_{j,i}$$

$$\boxed{\pi} \begin{bmatrix} P_{:,i} \end{bmatrix} = \boxed{\pi(i)}$$

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,i} = \sum_{j=1}^n \pi(j) P_{j,i} = \sum_{j \in \mathcal{N}(i)} \frac{d_j}{2|E|} \cdot \frac{1}{d_j}$$

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,i} = \sum_{j \in \mathcal{N}(i)} \pi(j) P_{j,i} = \sum_{j \in \mathcal{N}(i)} \frac{d_j}{2|E|} \cdot \frac{1}{d_j} = \sum_{j \in \mathcal{N}(i)} \frac{1}{2|E|}$$

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,j} = \sum_j \pi(j) P_{j,i} = \sum_j \frac{d_j}{2|E|} \cdot \frac{1}{d_j} = \sum_j \frac{1}{2|E|} = \frac{d_i}{2|E|} = \pi(i).$$

$\left(\sum_{i=1}^n d_i \right)$

Stationary Distribution Example 2

Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

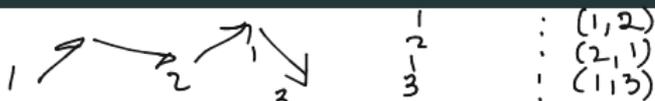
Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,i} = \sum_j \pi(j) P_{j,i} = \sum_j \frac{d_j}{2|E|} \cdot \frac{1}{d_j} = \sum_j \frac{1}{2|E|} = \frac{d_i}{2|E|} = \pi(i).$$

I.e., the probability of being at a given node i is dependent only on the node's degree, not on the structure of the graph in any other way.



Stationary Distribution Example 2



Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node i at step t , then it moves to any of i 's neighbors at step $t + 1$ with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,j} = \sum_j \pi(j) P_{j,i} = \sum_j \frac{d_j}{2|E|} \cdot \frac{1}{d_j} = \sum_j \frac{1}{2|E|} = \frac{d_i}{2|E|} = \pi(i).$$

I.e., the probability of being at a given node i is dependent only on the node's degree, not on the structure of the graph in any other way.

What is the stationary distribution over the edges?



$$\frac{1}{|E|}$$

Mixing Times

Total Variation Distance

Definition (Total Variation (TV) Distance)

For two distributions $p, q \in [0, 1]^m$ over state space $[m]$, the total variation distance is given by:

$$\|p - q\|_{TV} = \frac{1}{2} \sum_{i \in [m]} |p(i) - q(i)| = \max_{A \subseteq [m]} |p(A) - q(A)|.$$
$$\frac{1}{2} \|p - q\|_1$$

$$p = \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \\ 0 \end{bmatrix} \quad q = \begin{bmatrix} \frac{1}{3} \\ \frac{1}{3} \\ \frac{1}{3} \end{bmatrix}$$

$$\|p - q\|_{TV} = \frac{1}{2} \left(\frac{1}{6} + \frac{1}{6} + \frac{1}{3} \right) = \frac{1}{3}$$

A = event of being in state 3

A^c = event ... state 1 or 2

Total Variation Distance

Definition (Total Variation (TV) Distance)

For two distributions $p, q \in [0, 1]^m$ over state space $[m]$, the total variation distance is given by:

$$\|p - q\|_{TV} = \frac{1}{2} \sum_{i \in [m]} |p(i) - q(i)| = \max_{A \subseteq [m]} |p(A) - q(A)|.$$

Kantorovich-Rubinstein duality: Let P, Q be possibly correlated random variables with marginal distributions p, q . Then

$$\begin{aligned} P &= Q \\ P? & Q? \\ P &= Q \end{aligned}$$

$$\|p - q\|_{TV} \leq \Pr[P \neq Q].$$

$$= \min_{P, Q} \Pr(P \neq Q)$$

$$P = \begin{bmatrix} 1/2 \\ 1/2 \end{bmatrix} \quad Q = \begin{bmatrix} 3/4 \\ 1/4 \end{bmatrix}$$

$$P? \quad Q?$$

$$\Pr(P \neq Q) = \|p - q\|_{TV}$$

$$\begin{aligned} \text{If } P = H \text{ then } Q &= H \\ P = T \text{ then } Q &= \begin{matrix} H \text{ w.p. } 1/2 \\ T \text{ w.p. } 1/2 \end{matrix} \end{aligned}$$

Definition (Mixing Time)

Consider a Markov chain X_0, X_1, \dots with unique stationary distribution π . Let $q_{i,t}$ be the distribution over states at time t assuming $X_0 = i$. The mixing time is defined as:

$$\tau(\epsilon) = \min \left\{ t : \max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \epsilon \right\}.$$

I.e., what is the maximum time it takes the Markov chain to converge to within ϵ in TV distance of the stationary distribution?

Mixing Time

Definition (Mixing Time)

Consider a Markov chain X_0, X_1, \dots with unique stationary distribution π . Let $q_{i,t}$ be the distribution over states at time t assuming $X_0 = i$. The mixing time is defined as:

$$\tau(\epsilon) = \min \left\{ t : \max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \epsilon \right\}.$$

I.e., what is the maximum time it takes the Markov chain to converge to within ϵ in TV distance of the stationary distribution?

Note: If $\|q_{i,t} - \pi\|_{TV} \leq \epsilon$ then for any $t' \geq t$, $\|q_{i,t'} - \pi\|_{TV} \leq \epsilon$.

$$q_{i,t} = \pi + e_t$$
$$\|e_t\| = \|q_{i,t} - \pi\|_1 = 2 \cdot TV$$

$$q_{i,t+1} = q_{i,t}P = \pi P + e_t P$$
$$= \pi + e_t P$$

$$e_{t+1} = e_t P$$
$$\|e_{t+1}\| \leq \|e_t\|$$

(by Dobrushin)

Mixing Time Convergence

Typically, it suffices to focus on the mixing time for $\epsilon = 1/2$. We have:

Claim: If X_0, X_1, \dots is finite, irreducible, and aperiodic, then $\tau(\epsilon) \leq \tau(1/2) \cdot c \log(1/\epsilon)$ for large enough constant c .

$$q_{i,t} = \pi + e_t$$

$$q_{i,2t} = \pi + \underbrace{e_t P^t}_{\rightarrow \pi + e_{t+1}}$$

I'll do this later

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\|q_{i,t} - \pi\|_{TV} = \|q_{i,t} - \pi P^t\|_{TV}$$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\|q_{i,t} - \pi\|_{TV} = \|q_{i,t} - \pi P^t\|_{TV}$$

$$= \|q_{i,t} - \left(\sum_j \pi(j) \underbrace{e_j P^t}_{q_{j,t}} \right)\|_{TV}$$

$$\left[\pi(1) \quad \dots \quad \pi(m) \right]$$

$$= \pi(1) \cdot [1 \ 0 \ 0 \ 0 \ 0] + \pi(2) [0 \ 1 \ 0 \ 0] + \dots$$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\begin{aligned}\|q_{i,t} - \pi\|_{TV} &= \|q_{i,t} - \pi P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) e_j P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) q_{j,t}\|_{TV}\end{aligned}$$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\begin{aligned}\|q_{i,t} - \pi\|_{TV} &= \|q_{i,t} - \pi P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) e_j P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) q_{j,t}\|_{TV} \\ &\leq \sum_j \|\pi(j) q_{i,t} - \pi(j) q_{j,t}\|_{TV}\end{aligned}$$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\begin{aligned}\|q_{i,t} - \pi\|_{TV} &= \|q_{i,t} - \pi P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) e_j P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) q_{j,t}\|_{TV} \\ &\leq \sum_j \|\pi(j) q_{i,t} - \pi(j) q_{j,t}\|_{TV} \\ &\leq \sum_j \pi(j) \cdot \|q_{i,t} - q_{j,t}\|_{TV}\end{aligned}$$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\|q_{i,t} - \pi\|_{TV} = \|q_{i,t} - \pi P^t\|_{TV}$$

$$= \|q_{i,t} - \sum_j \pi(j) e_j P^t\|_{TV}$$

$$= \|q_{i,t} - \sum_j \pi(j) q_{j,t}\|_{TV} = \left\| \sum_j \pi(j) q_{i,t} - \sum_j \pi(j) q_{j,t} \right\|$$

$$\leq \sum_j \|\pi(j) q_{i,t} - \pi(j) q_{j,t}\|_{TV}$$

$$\leq \sum_j \pi(j) \cdot \|q_{i,t} - q_{j,t}\|_{TV}$$

$$\leq \max_{j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$$

$$\pi = \sum \pi(j) e_j$$

$$[1 \quad 2 \quad 4]$$

$$= 1 \cdot [1, 0, 0]$$

$$+ 2 \cdot [0, 1, 0]$$

$$+ 4 \cdot [0, 0, 1]$$

~~$$\|q_{i,t} - \pi\| \leq \|q_{i,t} - \sum_j \pi(j) q_{j,t}\| + \|\sum_j \pi(j) q_{j,t} - \pi\|$$~~

use that
 $\sum \pi(j) = 1$

Coupling Motivation

Claim: $\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$.

$$\begin{aligned}\|q_{i,t} - \pi\|_{TV} &= \|q_{i,t} - \pi P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) e_j P^t\|_{TV} \\ &= \|q_{i,t} - \sum_j \pi(j) q_{j,t}\|_{TV} \\ &\leq \sum_j \|\pi(j) q_{i,t} - \pi(j) q_{j,t}\|_{TV} \\ &\leq \sum_j \pi(j) \cdot \|q_{i,t} - q_{j,t}\|_{TV} \\ &\leq \max_{j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}.\end{aligned}$$

Coupling: A common technique for bounding the mixing time by showing that $\max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV}$ is small.

Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

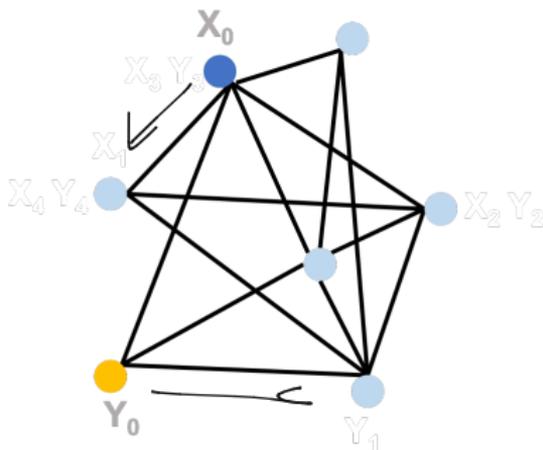
1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.

Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.

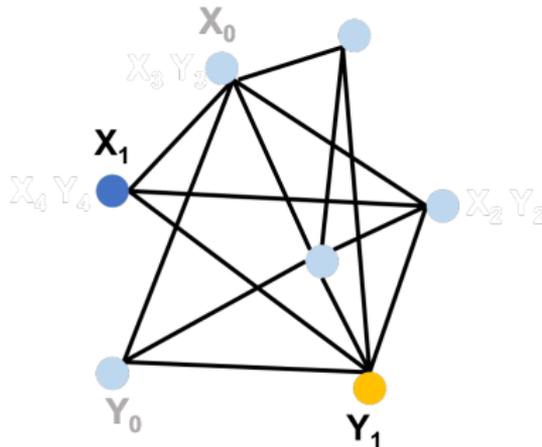


Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.

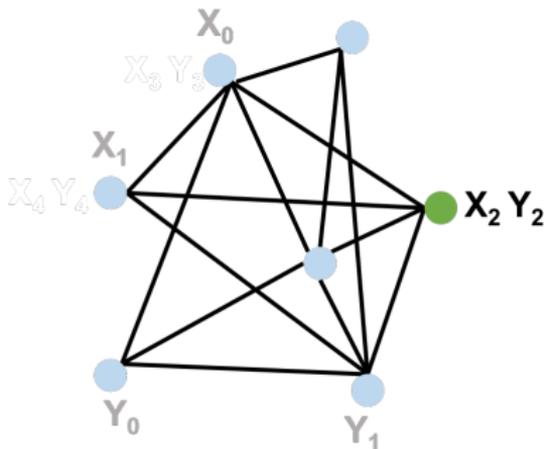


Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.



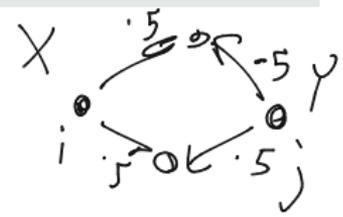
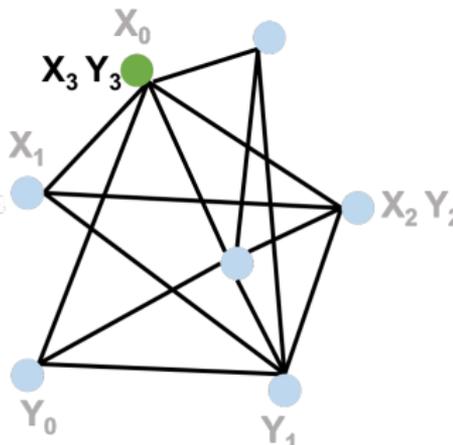
Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.

$$X_{t+1} = Y_{t+1}$$
$$\Pr(X_{t+1} = j | X_t = i) = P_{i,j}$$
$$\vdots$$

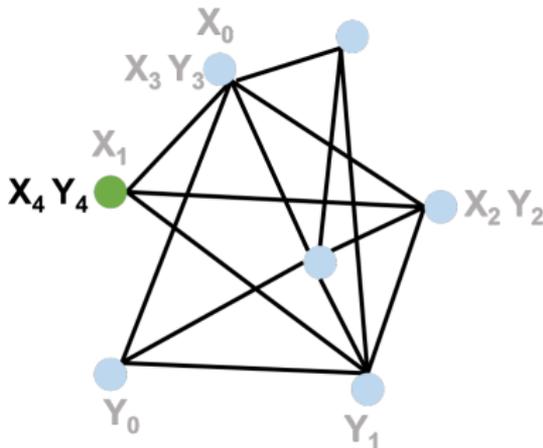


Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.



$$\Pr(T_{ij} = 1) = 1/2$$

$$\Pr(\bar{T}_{ij} = 2) = 1/4$$

⋮

Formal Coupling Definition

Definition (Coupling)

For a finite Markov chain X_0, X_1, \dots with transition matrix $P \in \mathbb{R}^{m \times m}$, a coupling is a joint process $(X_0, Y_0), (X_1, Y_1), \dots$ such that:

1. $X_0 = i$ and $Y_0 = j$ for some $i, j \in [m]$.
2. $\Pr[X_t = j | X_{t-1} = i] = \Pr[Y_t = j | Y_{t-1} = i] = P_{i,j}$
3. If $X_t = Y_t$, then $X_{t+1} = Y_{t+1}$.

Theorem (Mixing Time Bound via Coupling)

For a finite, irreducible, and aperiodic Markov chain X_0, X_1, \dots and any valid coupling $(X_0, Y_0), (X_1, Y_1), \dots$ letting

$$T_{i,j} = \min\{t : X_t = Y_t | X_0 = i, Y_0 = j\},$$

$$\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV} \leq \max_{i,j \in [m]} \Pr[T_{i,j} > t].$$

Coupling Theorem Proof

Theorem (Mixing Time Bound via Coupling)

For a finite, irreducible, and aperiodic Markov chain X_0, X_1, \dots and any valid coupling $(X_0, Y_0), (X_1, Y_1), \dots$ letting

$$T_{i,j} = \min \{ t : X_t = Y_t \mid X_0 = i, Y_0 = j \}$$

$$\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV} \leq \max_{i,j \in [m]} \Pr[T_{i,j} > t].$$

Follows from **Kantorovich-Rubinstein duality**.

Coupling Theorem Proof

Theorem (Mixing Time Bound via Coupling)

For a finite, irreducible, and aperiodic Markov chain X_0, X_1, \dots and any valid coupling $(X_0, Y_0), (X_1, Y_1), \dots$ letting

$$T_{i,j} = \min\{t : X_t = Y_t | X_0 = i, Y_0 = j\},$$

$$\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV} \leq \max_{i,j \in [m]} \Pr[T_{i,j} > t].$$

Follows from [Kantorovich-Rubinstein duality](#).

For X_t, Y_t distributed by evolving the chain for t steps starting from state i or j respectively, we have:

$$\max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV} \leq \max_{i,j \in [m]} \Pr[X_t \neq Y_t] = \max_{i,j \in [m]} \Pr[T_{i,j} > t]$$

$\begin{matrix} | & | \\ \times & \gamma \end{matrix}$

Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

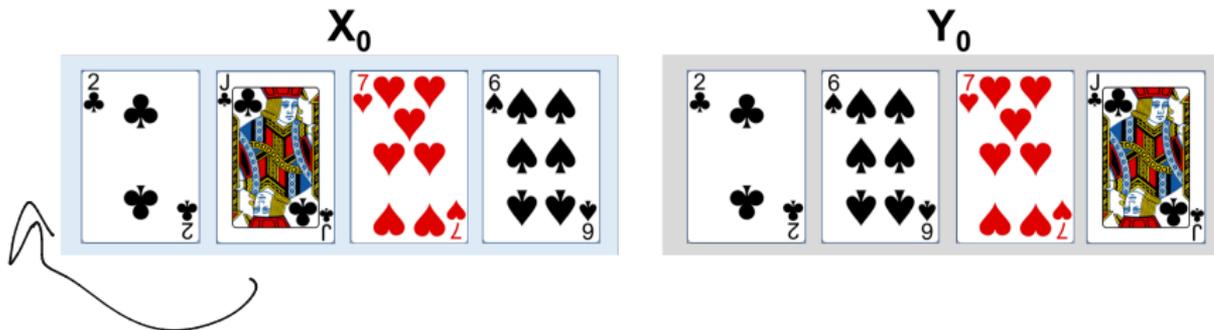
- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.

Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.

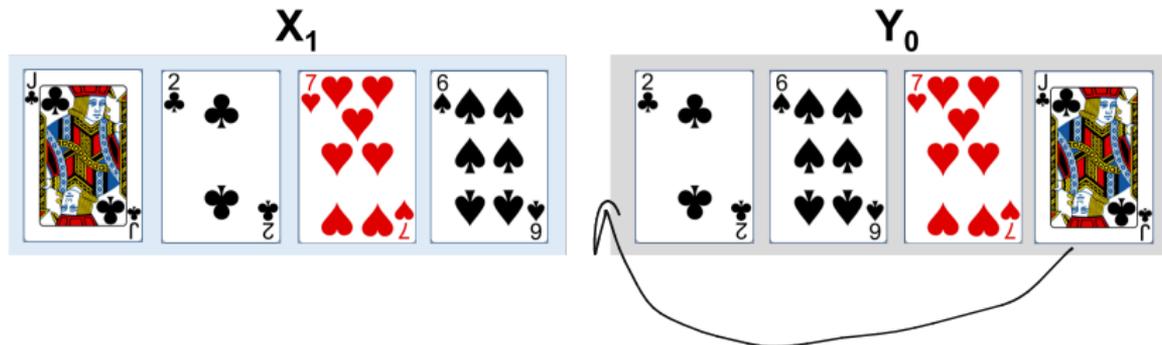


Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.

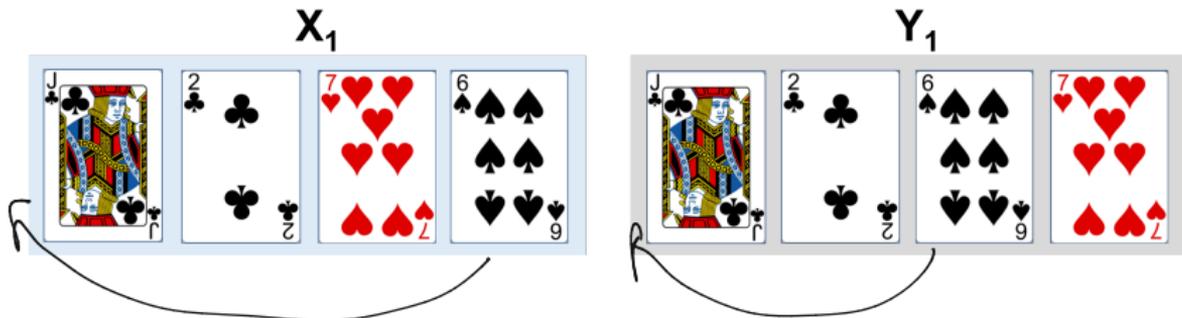


Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.

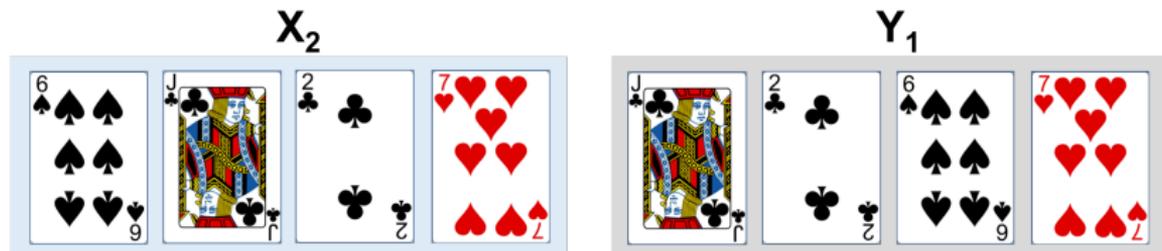


Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.

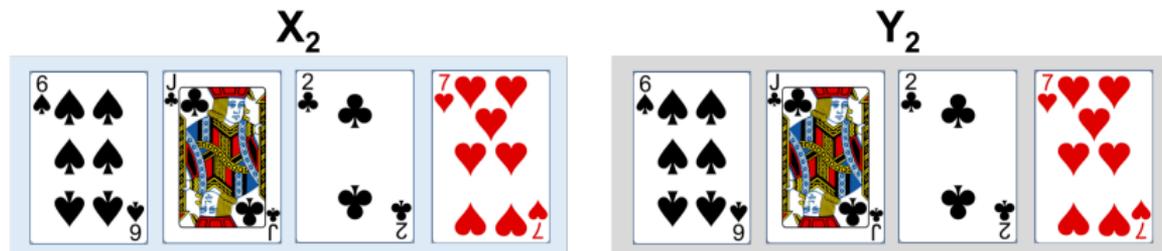


Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.



Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.
- Can check that this is a valid coupling since X_t, Y_t have the correct marginal distributions, and since
$$X_t = Y_t \implies X_{t+1} = Y_{t+1}$$

Coupling Example: Mixing Time of Shuffling

How many times do we need to swap a random card to the top of the deck so that the distribution of orderings on our cards is ϵ -close in TV distance to the uniform distribution over all permutations?

Coupling:

- Let X_0, X_1, \dots be the Markov chain where a random card is moved to the top in each step.
- Let Y_0, Y_1 be a correlated Markov chain. When card S is swapped to the top in the X chain, swap S to the top in the Y chain as well.
- Can check that this is a valid coupling since X_t, Y_t have the correct marginal distributions, and since
$$X_t = Y_t \implies X_{t+1} = Y_{t+1}$$
- Observe that $X_t = Y_t$ as soon as all c unique cards have been swapped at least once. How many swaps does this take?

Coupling Example: Mixing Time of Shuffling

$$\begin{aligned} \max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} &\leq \max_{i,j \in [m]} \Pr[T_{i,j} > t] \\ &\leq \Pr[\text{< } c \text{ unique cards are swapped in } t \text{ swaps}] \end{aligned}$$

Coupling Example: Mixing Time of Shuffling

$$\begin{aligned} \max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} &\leq \max_{i,j \in [m]} \Pr[T_{i,j} > t] \\ &\leq \Pr[\text{< } c \text{ unique cards are swapped in } t \text{ swaps}] \end{aligned}$$

By coupon collector analysis for $t \geq c \ln(c/\epsilon)$, this probability is bounded by ϵ . In particular, by the fact that $(1 - \frac{1}{c})^{c \ln c/\epsilon} \leq \frac{\epsilon}{c}$ plus a union bound over c cards.

Coupling Example: Mixing Time of Shuffling

$$\begin{aligned}\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} &\leq \max_{i,j \in [m]} \Pr[T_{i,j} > t] \\ &\leq \Pr[\text{< } c \text{ unique cards are swapped in } t \text{ swaps}]\end{aligned}$$

By coupon collector analysis for $t \geq c \ln(c/\epsilon)$, this probability is bounded by ϵ . In particular, by the fact that $(1 - \frac{1}{c})^{c \ln c/\epsilon} \leq \frac{\epsilon}{c}$ plus a union bound over c cards.

Thus, for $t \geq c \ln(c/\epsilon)$,

$$\max_{i \in [m]} \|q_{i,t} - \pi\|_{TV} \leq \max_{i,j \in [m]} \|q_{i,t} - q_{j,t}\|_{TV} \leq \epsilon.$$

I.e., $\tau(\epsilon) \leq c \ln(c/\epsilon)$.