### COMPSCI 514: ALGORITHMS FOR DATA SCIENCE

Cameron Musco University of Massachusetts Amherst. Spring 2020. Lecture 3

#### By Thursday:

- Sign up for Piazza.
- Sign up for Gradescope (code on class website) and fill out the Gradescope consent poll on Piazza. Contact me via email if you don't consent to use Gradescope.
- First problem set will be **available in the next day or two**, due 2/14.

#### Last Class We Covered:

- Markov's inequality: the most fundamental concentration bound.
- Random hash functions, collision free hashing, and two-level hashing (analysis with linearity of expectation and Markov's inequality.)
- $\cdot$  2-universal and pairwise independent hash functions.
- Chebyshev's inequality and an elementary proof of the law of large numbers.

**Today:** We'll see even stronger concentration bounds than Chebyshev's inequality – exponential tail bounds.

• Will show a version of the central limit theorem.



**First:** We'll show learn about the union bound and apply it to randomized load balancing.

#### Randomized Load Balancing:



- *n* requests randomly assigned to *k* servers using a random hash function.
- Letting  $\mathbf{R}_i$  be the number of requests assigned to server *i*,  $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$  and we provision each server with the capacity to serve twice its expected load:  $\frac{2n}{k}$  requests.
- What is the probability that a server exceeds its capacity?
- To apply Chebyshev's inequality, need to bound Var[R<sub>i</sub>].

Recall that we can write  $\mathbf{R}_i$  as:

 $\mathbf{R}_{i} = \sum_{j=1}^{n} \mathbf{R}_{i,j} \operatorname{Var}[\mathbf{R}_{i}] = \sum_{j=1}^{n} \operatorname{Var}[\mathbf{R}_{i,j}] \qquad \text{(linearity of variance)}$ 

where  $\mathbf{R}_{i,j}$  is 1 if request *j* is assigned to server *i* and 0 o.w.

$$\begin{aligned} \operatorname{Var}[\mathbf{R}_{i,j}] &= \mathbb{E}\left[\left(\mathbf{R}_{i,j} - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2}\right] \\ &= \operatorname{Pr}(\mathbf{R}_{i,j} = 1) \cdot \left(1 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2} + \operatorname{Pr}(\mathbf{R}_{i,j} = 0) \cdot \left(0 - \mathbb{E}[\mathbf{R}_{i,j}]\right)^{2} \\ &= \frac{1}{k} \cdot \left(1 - \frac{1}{k}\right)^{2} + \left(1 - \frac{1}{k}\right) \cdot \left(0 - \frac{1}{k}\right)^{2} \\ &= \frac{1}{k} - \frac{1}{k^{2}} \leq \frac{1}{k} \implies \operatorname{Var}[\mathbf{R}_{i}] \leq \frac{n}{k}. \end{aligned}$$

n: total number of requests, k: number of servers randomly assigned requests,  $R_i$ : number of requests assigned to server i.

Letting  $\mathbf{R}_i$  be the number of requests sent to server i,  $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$ and  $\operatorname{Var}[\mathbf{R}_i] \leq \frac{n}{k}$ .

Applying Chebyshev's:

$$\Pr\left(\mathbf{R}_{i} \geq \frac{2n}{k}\right) \leq \Pr\left(|\mathbf{R}_{i} - \mathbb{E}[\mathbf{R}_{i}]| \geq \frac{n}{k}\right) \leq \frac{n/k}{n^{2}/k^{2}} = \frac{k}{n}.$$

- Overload probability is extremely small when  $k \ll n!$
- Might seem counterintuitive bound gets worse as k grows.
- When *k* is large, the number of requests each server sees in expectation is very small so the law of large numbers doesn't 'kick in'.

n: total number of requests, k: number of servers randomly assigned requests,  $R_i$ : number of requests assigned to server i.

What is the probability that the maximum server load exceeds  $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{k}$ . I.e., that some server is overloaded if we give each  $\frac{2n}{k}$  capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right) = \Pr\left(\left[\mathbf{R}_{1} \geq \frac{2n}{k}\right] \cup \left[\mathbf{R}_{2} \geq \frac{2n}{k}\right] \cup \ldots \cup \left[\mathbf{R}_{k} \geq \frac{2n}{k}\right]\right) = \Pr\left(\left[\mathbf{R}_{1} \geq \frac{2n}{k}\right]\right)$$

We want to show that  $\Pr\left(\bigcup_{i=1}^{k} \left[\mathbf{R}_{i} \geq \frac{2n}{k}\right]\right)$  is small.

How do we do this? Note that  $\mathbf{R}_1, \ldots, \mathbf{R}_k$  are correlated in a somewhat complex way.

*n*: total number of requests, *k*: number of servers randomly assigned requests,  $R_i$ : number of requests assigned to server *i*.  $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$ .  $Var[\mathbf{R}_i] = \frac{n}{k}$ .

**Union Bound:** For any random events  $A_1, A_2, ..., A_k$ ,

 $\Pr(A_1 \cup A_2 \cup \ldots \cup A_k) \leq \Pr(A_1) + \Pr(A_2) + \ldots + \Pr(A_k).$ 



When is the union bound tight? When  $A_1, ..., A_k$  are all disjoint.

On the first problem set, you will prove the union bound, as a consequence of Markov's inquality.

What is the probability that the maximum server load exceeds  $2 \cdot \mathbb{E}[\mathbf{R}_i] = \frac{2n}{k}$ . I.e., that some server is overloaded if we give each  $\frac{2n}{k}$  capacity?

$$\Pr\left(\max_{i}(\mathbf{R}_{i}) \geq \frac{2n}{k}\right) = \Pr\left(\bigcup_{i=1}^{k} \left[\mathbf{R}_{i} \geq \frac{2n}{k}\right]\right)$$
$$\leq \sum_{i=1}^{k} \Pr\left(\left[\mathbf{R}_{i} \geq \frac{2n}{k}\right]\right) \qquad \text{(Union Bound)}$$
$$\leq \sum_{i=1}^{k} \frac{k}{n} = \frac{k^{2}}{n} \qquad \text{(Bound from Chebyshev's)}$$

As long as  $k \leq O(\sqrt{n})$ , with good probability, the maximum server load will be small (compared to the expected load).

*n*: total number of requests, *k*: number of servers randomly assigned requests, **R**<sub>i</sub>: number of requests assigned to server *i*.  $\mathbb{E}[\mathbf{R}_i] = \frac{n}{k}$ .  $Var[\mathbf{R}_i] = \frac{n}{k}$ . The number of servers must be small compared to the number of requests ( $k = O(\sqrt{n})$ ) for the maximum load to be bounded in comparison to the expected load with good probability.

- There are many requests routed to a relatively small number of servers so the load seen on each server is close to what is expected via law of large numbers.
- A Useful Exercise: Given n requests, and assuming all servers have fixed capacity C, how many servers should you provision so that with probability ≥ 99/100 no server is assigned more than C requests?

*n*: total number of requests, *k*: number of servers randomly assigned requests.

# Questions on union bound, Chebyshev's inequality, random hashing?

We flip n = 100 independent coins, each are heads with probability 1/2 and tails with probability 1/2. Let **H** be the number of heads.

$$\mathbb{E}[H] = \frac{n}{2} = 50 \text{ and } Var[H] = \frac{n}{4} = 25 \rightarrow s.d. = 5$$
Markov's:Chebyshev's:In Reality: $Pr(H \ge 60) \le .833$  $Pr(H \ge 60) \le .25$  $Pr(H \ge 60) = 0.0284$  $Pr(H \ge 70) \le .714$  $Pr(H \ge 70) \le .0625$  $Pr(H \ge 70) = .000039$  $Pr(H \ge 80) \le .625$  $Pr(H \ge 80) \le .0278$  $Pr(H \ge 80) < 10^{-9}$ 

**H** has a simple Binomial distribution, so can compute these probabilities exactly.

**To be fair....** Markov and Chebyshev's inequalities apply much more generally than to Binomial random variables like coin flips.

Can we obtain tighter concentration bounds that still apply to very general distributions?

- Markov's:  $Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$ . First Moment.
- Chebyshev's:  $Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| \ge t) = Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]|^2 \ge t^2) \le \frac{Var[\mathbf{X}]}{t^2}$ . Second Moment.
- What if we just apply Markov's inequality to even higher moments?

Consider any random variable X:

$$\Pr(|\mathbf{X} - \mathbb{E}[\mathbf{X}]| \ge t) = \Pr\left((\mathbf{X} - \mathbb{E}[\mathbf{X}])^4 \ge t^4\right) \le \frac{\mathbb{E}\left[(\mathbf{X} - \mathbb{E}[\mathbf{X}])^4\right]}{t^4}$$

Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\mathbb{E}\left[\left(\mathsf{H} - \mathbb{E}[\mathsf{H}]\right)^{4}\right] = \mathbb{E}\left[\left(\sum_{i=1}^{100} \mathsf{H}_{i} - 50\right)^{4}\right] = \sum_{i,j,k,\ell} c_{ijk\ell} \mathbb{E}[\mathsf{H}_{i}\mathsf{H}_{j}\mathsf{H}_{k}\mathsf{H}_{\ell}] = 1862.5$$

where  $H_i = 1$  if coin flip *i* is heads and 0 otherwise. Then apply some messy calculations...

• Apply Fourth Moment Bound:  $\Pr(|\mathbf{H} - \mathbb{E}[\mathbf{H}]| \ge t) \le \frac{1862.5}{t^4}$ .

Chebyshev's:	4 <sup>th</sup> Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$\Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$\Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds?

- Yes! To a point.
- In fact don't need to just apply Markov's to  $|\mathbf{X} \mathbb{E}[\mathbf{X}]|^k$  for some *k*. Can apply to any monotonic function  $f(|\mathbf{X} \mathbb{E}[\mathbf{X}]|)$ .
- Why monotonic?  $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| > t) = \Pr(f(|\mathbf{X} \mathbb{E}[\mathbf{X}]|) > f(t)).$

H: total number heads in 100 random coin flips.  $\mathbb{E}[H] = 50$ .

**Moment Generating Function:** Consider for any t > 0:

$$M_t(\mathsf{X}) = e^{t \cdot (\mathsf{X} - \mathbb{E}[\mathsf{X}])} = \sum_{k=0}^{\infty} \frac{t^k (\mathsf{X} - \mathbb{E}[\mathsf{X}])^k}{k!}$$

- $M_t(\mathbf{X})$  is monotonic for any t > 0.
- Weighted sum of all moments, with *t* controlling how slowly the weights fall off (larger *t* = slower falloff).
- Choosing *t* appropriately lets one prove a number of very powerful exponential concentration bounds (exponential tail bounds).
- Chernoff bound, Bernstein inequalities, Hoeffding's inequality, Azuma's inequality, Berry-Esseen theorem, etc.
- $\cdot$  We will not cover the proofs in the this class.

Bernstein Inequality: Consider independent random variables  $X_1, \ldots, X_n$  all falling in [-M, M] [-1,1]. Let  $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$  and  $\sigma^2 =$  $\operatorname{Var}[\sum_{i=1}^{n} \mathbf{X}_{i}] = \sum_{i=1}^{n} \operatorname{Var}[\mathbf{X}_{i}].$  For any  $t \ge 0$ s  $\ge 0$ :  $\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \geq t\right) \leq 2 \exp\left(-\frac{t^{2}}{2\sigma^{2} + \frac{4}{3}Mt}\right).$  $\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^{2}}{4}\right).$ 

Assume that M = 1 and plug in  $t = s \cdot \sigma$  for  $s \leq \sigma$ .

Compare to Chebyshev's:  $Pr\left(\left|\sum_{i=1}^{n} X_{i} - \mu\right| \ge s\sigma\right) \le \frac{1}{s^{2}}$ .

• An exponentially stronger dependence on s!

Consider again bounding the number of heads H in n = 100 independent coin flips.

Chebyshev's:	Bernstein:	In Reality:
$\Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .15$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .00086$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le 3^{-7}$	$Pr(H \ge 80) < 10^{-9}$

Getting much closer to the true probability.

**H**: total number heads in 100 random coin flips.  $\mathbb{E}[\mathbf{H}] = 50$ .

**Bernstein Inequality:** Consider independent random variables  $X_1, \ldots, X_n$  falling in [-1,1]. Let  $\mu = \mathbb{E}[\sum X_i]$  and  $\sigma^2 = \text{Var}[\sum X_i]$ .  $\Pr\left(\left|\sum_{i=1}^n X_i - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^2}{4}\right)$ .

Can plot this bound for different s:



Looks a lot like a Gaussian (normal) distribution.

$$\mathcal{N}(0,\sigma^2)$$
 has density  $p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{x^2}{2\sigma^2}}$ 

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 has density  $p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{x^2}{2\sigma^2}}$ .

**Exercise:** Using this can show that for  $X \sim \mathcal{N}(0, \sigma^2)$ : for any  $s \ge 0$ ,

$$\Pr(|\mathbf{X}| \ge s \cdot \sigma) \le O(1) \cdot e^{-\frac{s^2}{2}}.$$

Essentially the same bound that Bernstein's inequality gives!

**Central Limit Theorem Interpretation:** Bernstein's inequality gives a quantitative version of the CLT. The distribution of the sum of *bounded* independent random variables can be upper bounded with a Gaussian (normal) distribution.



**Stronger Central Limit Theorem:** The distribution of the sum of *n bounded* independent random variables converges to a Gaussian (normal) distribution as *n* goes to infinity.



- Why is the Gaussian distribution is so important in statistics, science, ML, etc.?
- Many random variables can be approximated as the sum of a large number of small and roughly independent random effects. Thus, their distribution looks Gaussian by CLT.

A useful variation of the Bernstein inequality for binary (indicator) random variables is:

**Chernoff Bound (simplified version):** Consider independent random variables  $X_1, \ldots, X_n$  taking values in  $\{0, 1\}$ . Let  $\mu = \mathbb{E}[\sum_{i=1}^{n} X_i]$ . For any  $\delta \ge 0$ 

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \geq \delta \mu\right) \leq 2 \exp\left(-\frac{\delta^{2} \mu}{2 + \delta}\right).$$

As  $\delta$  gets larger and larger, the bound falls of exponentially fast.

#### **RETURN TO RANDOM HASHING**



We hash *m* values  $x_1, \ldots, x_m$  using a random hash function into a table with n = m entries.

• I.e., for all  $j \in [m]$  and  $i \in [n]$ ,  $Pr(h(x) = i) = \frac{1}{m}$  and hash values are chosen independently.

What will be the maximum number of items hashed into the same location?  $O(n) \quad O(\log n) \quad O(\sqrt{n}) \quad O(1/n)$  What will be the maximum number of items hashed into the same location?  $O(\log m)$ 

Let  $S_i$  be the number of items hashed into position *i* and  $S_{i,j}$  be 1 if  $x_j$  is hashed into bucket *i* ( $h(x_i) = i$ ) and 0 otherwise.

$$\mathbb{E}[\mathbf{S}_i] = \sum_{j=1}^m \mathbb{E}[\mathbf{S}_{i,j}] = m \cdot \frac{1}{m} = 1 = \mu.$$

By the Chernoff Bound: for any  $\delta \ge 0$ ,

$$\Pr(\mathbf{S}_i \ge 1 + \delta) \le \Pr\left(\left|\sum_{i=1}^n \mathbf{S}_{i,j} - 1\right| \ge \delta\right) \le 2\exp\left(-\frac{\delta^2}{2 + \delta}\right)$$

*m*: total number of items hashed and size of hash table.  $x_1, \ldots, x_m$ : the items. h: random hash function mapping  $x_1, \ldots, x_m \rightarrow [m]$ .

$$\Pr(\mathbf{S}_i \ge 1 + \delta) \le \Pr\left(\left|\sum_{i=1}^n \mathbf{S}_{i,j} - 1\right| \ge \delta\right) \le 2 \exp\left(-\frac{\delta^2}{2 + \delta}\right).$$

Set  $\delta = 20 \log m$ . Gives:

$$\Pr(\mathbf{S}_i \ge 20 \log m + 1) \le 2 \exp\left(-\frac{(20 \log m)^2}{2 + 20 \log m}\right) \le \exp(-18 \log m) \le \frac{2}{m^{18}}.$$

Apply Union Bound:

$$\Pr(\max_{i \in [m]} \mathbf{S}_i \ge 20 \log m + 1) = \Pr\left(\bigcup_{i=1}^m (\mathbf{S}_i \ge 20 \log n + 1)\right)$$
$$\leq \sum_{i=1}^m \Pr(\mathbf{S}_i \ge 20 \log m + 1) \le m \cdot \frac{2}{m^{18}} = \frac{2}{m^{17}}.$$

*m*: total number of items hashed and size of hash table.  $S_i$ : number of items hashed to bucket *i*.  $S_{i,j}$ : indicator if  $x_i$  is hashed to bucket *i*.  $\delta$ : any value  $\geq 0$ .

**Upshot:** If we randomly hash m items into a hash table with m entries the maximum load per bucket is  $O(\log m)$  with very high probability.

- So, even with a simple linked list to store the items in each bucket, worst case query time is  $O(\log m)$ .
- Using Chebyshev's inequality could only show the maximum load is bounded by  $O(\sqrt{m})$  with good probability.
- The Chebyshev bound holds even with a pairwise independent hash function. The stronger Chernoff-based bound can be shown to hold with a *k*-wise independent hash function for  $k = O(\log m)$ .

## Questions?

This concludes probability review/concentration bounds.