COMPSCI 514: Algorithms for Data Science

Cameron Musco University of Massachusetts Amherst. Fall 2023. Lecture 5

Logistics

$$|n - \hat{n}| \leq \xi n$$

 $n - \hat{n} \leq \epsilon n$
 $n - \hat{n} \leq \epsilon n$
 $n - \hat{n} \geq -\epsilon n$

- Problem Set 1 is due this Friday at 11:59pm.
- A useful technique: to prove that $|a| \le b$, prove both sides: that $a \le b$ and that $a \ge -b$.
- Quiz question on class pacing:
 - Way too fast: 9.
 - A bit too fast: 43.
 - Just right: 64.
 - A bit too slow: 3.
 - Way too slow: 0.

Last Time

$$h \quad h(x) = ax + b \mod n$$

5

Last Class:

- $\cdot\,$ 2-universal and pairwise independent hash functions.
- The union bound.
- Application to hashing for load balancing.

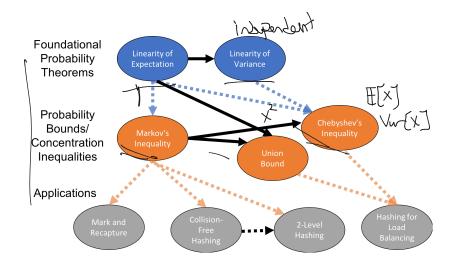
Last Class:

- 2-universal and pairwise independent hash functions.
- The union bound.
- Application to hashing for load balancing.

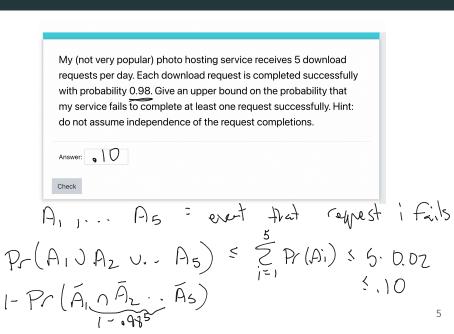
This Time:

• Exponential concentration bounds and the central limit theorem.

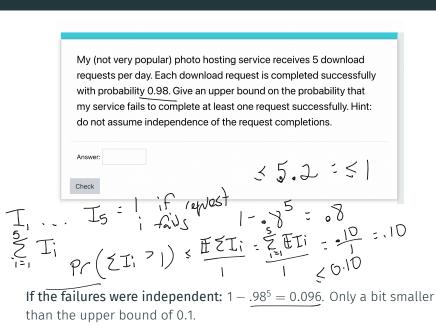
Concept Map



Quiz Questions



Quiz Questions



More Union Bound Intuition

Flipping Coins

We flip n = 100 independent coins, each are heads with probability 1/2 and tails with probability 1/2. Let H be the number of heads.

$$\mathbb{E}[H] = 50$$
 and $Var[H] = \sqrt{\sqrt{\left(\sum_{i=1}^{n} H_{i}\right)}}$

$$= \frac{100}{2} lw(Hi)$$

$$= Hi^{2} - (Hi)^{2}$$

$$= Hi^{2} - (Hi)^{2}$$

$$= 5 - .25 = .25$$

$$= .25$$

$$= .25$$

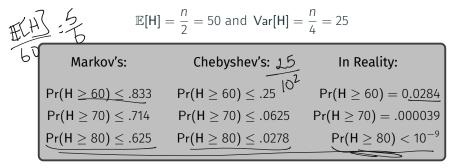
Flipping Coins

We flip n = 100 independent coins, each are heads with probability 1/2 and tails with probability 1/2. Let **H** be the number of heads.

$$\mathbb{E}[\mathbf{H}] = \frac{n}{2} = 50 \text{ and } \mathbf{Var}[\mathbf{H}] = \frac{n}{4} = 25$$

Flipping Coins

We flip n = 100 independent coins, each are heads with probability 1/2 and tails with probability 1/2. Let **H** be the number of heads.



H has a simple Binomial distribution, so can compute these probabilities exactly.

Can we obtain tighter concentration bounds that still apply to very general distributions?

Can we obtain tighter concentration bounds that still apply to very general distributions?

• Markov's: $\Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$. First Moment.

Can we obtain tighter concentration bounds that still apply to very general distributions?

- Markov's: $Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$. First Moment.
- Chebyshev's: $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| \ge t) = \Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]|^2 \ge t^2) \le \frac{\operatorname{Var}[\mathbf{X}]}{t^2}$. Second Moment.

Can we obtain tighter concentration bounds that still apply to very general distributions?

- Markov's: $Pr(X \ge t) \le \frac{\mathbb{E}[X]}{t}$. First Moment.
- Chebyshev's: $\Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]| \ge t) = \Pr(|\mathbf{X} \mathbb{E}[\mathbf{X}]|^2 \ge t^2) \le \frac{\operatorname{Var}[\mathbf{X}]}{t^2}$. Second Moment.
- What if we just apply Markov's inequality to even higher moments?

A Fourth Moment Bound

Consider any random variable X:

$$\Pr(|\mathbf{X} - \mathbb{E}[\mathbf{X}]| \ge t) = \Pr\left((\mathbf{X} - \mathbb{E}[\mathbf{X}])^4 \ge t^4\right)$$

$$\mathbf{X}^{\mathcal{N}} \stackrel{\stackrel{>}{\rightarrow}}{\longrightarrow} \quad \mathbf{|\mathbf{X}|} \stackrel{>}{\rightarrow} \underbrace{|\mathbf{Y}|}{\longrightarrow}$$

A Fourth Moment Bound

Consider any random variable X:

$$\Pr(|X - \mathbb{E}[X]| \ge t) = \Pr((X - \mathbb{E}[X])^4 \ge t^4) \le \frac{\mathbb{E}\left[(X - \mathbb{E}[X])^4\right]}{t^4}.$$

$$\mathsf{Pr}(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \ge t) = \mathsf{Pr}\left(\left(\mathsf{X} - \mathbb{E}[\mathsf{X}]\right)^4 \ge t^4\right) \le \frac{\mathbb{E}\left[\left(\underbrace{\mathsf{X} - \mathbb{E}[\mathsf{X}]}_{t^4}\right)^4\right]}{t^4}.$$

Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\Pr(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \ge t) = \Pr\left((\mathsf{X} - \mathbb{E}[\mathsf{X}])^4 \ge t^4\right) \le \frac{\mathbb{E}\left[(\mathsf{X} - \mathbb{E}[\mathsf{X}])^4\right]}{t^4}$$

Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\mathbb{E}\left[\left(\mathsf{H} - \mathbb{E}[\mathsf{H}]\right)^{4}\right] = \mathbb{E}\left[\left(\sum_{i=1}^{100} \mathsf{H}_{i} - 50\right)^{4}\right]$$

where $H_i = 1$ if coin flip *i* is heads and 0 otherwise.

$$\Pr(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \ge t) = \Pr\left((\mathsf{X} - \mathbb{E}[\mathsf{X}])^4 \ge t^4\right) \le \frac{\mathbb{E}\left[(\mathsf{X} - \mathbb{E}[\mathsf{X}])^4\right]}{t^4}$$

Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\mathbb{E}\left[\left(\mathsf{H} - \mathbb{E}[\mathsf{H}]\right)^{4}\right] = \mathbb{E}\left[\left(\sum_{i=1}^{100}\mathsf{H}_{i} - 50\right)^{4}\right] = \sum_{\substack{i,j,k,\ell \in \mathcal{L} \\ i,j,k,\ell \in \mathcal{L} \\ k \neq \ell}} \underbrace{c_{ijk\ell}}_{lk} \mathbb{E}\left[\mathsf{H}_{i}\mathsf{H}_{j}\mathsf{H}_{k}\mathsf{H}_{\ell}\right]_{lb}$$
where $\mathsf{H}_{i} = 1$ if coin flip *i* is heads and 0 otherwise. Then apply some messy calculations...

A Fourth Moment Bound

Consider any random variable X:

$$\begin{aligned}
\left(X \downarrow Y\right)^{C_{-}} & X^{L} \downarrow Y \downarrow Y^{2} \\
Pr(|X - \mathbb{E}[X]| \ge t) = Pr\left(\left(X - \mathbb{E}[X]\right)^{4} \ge t^{4}\right) \le \frac{\mathbb{E}\left[\left(X - \mathbb{E}[X]\right)^{4}\right]}{t^{4}}.
\end{aligned}$$

Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\mathbb{E}\left[\left(\mathsf{H} - \mathbb{E}[\mathsf{H}]\right)^{4}\right] = \mathbb{E}\left[\left(\sum_{i=1}^{100}\mathsf{H}_{i} - 50\right)^{4}\right] = \sum_{i,j,k,\ell} \mathcal{C}_{ijk\ell} \mathbb{E}[\mathsf{H}_{i}\mathsf{H}_{j}\mathsf{H}_{k}\mathsf{H}_{\ell}] = 1862.5$$

where $H_i = 1$ if coin flip *i* is heads and 0 otherwise. Then apply some messy calculations...

$$\mathsf{Pr}(|\mathsf{X} - \mathbb{E}[\mathsf{X}]| \ge t) = \mathsf{Pr}\left(\left(\mathsf{X} - \mathbb{E}[\mathsf{X}]\right)^4 \ge t^4\right) \le \frac{\mathbb{E}\left[\left(\mathsf{X} - \mathbb{E}[\mathsf{X}]\right)^4\right]}{t^4}$$

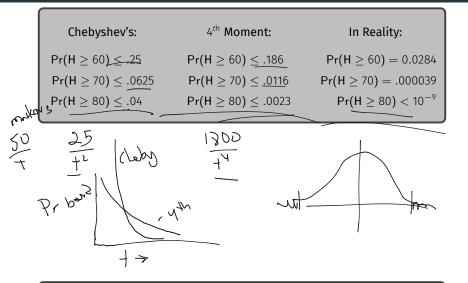
Application to Coin Flips: Recall: n = 100 independent fair coins, H is the number of heads.

• Bound the fourth moment:

$$\mathbb{E}\left[\left(\mathsf{H} - \mathbb{E}[\mathsf{H}]\right)^{4}\right] = \mathbb{E}\left[\left(\sum_{i=1}^{100} \mathsf{H}_{i} - 50\right)^{4}\right] = \sum_{i,j,k,\ell} c_{ijk\ell} \mathbb{E}[\mathsf{H}_{i}\mathsf{H}_{j}\mathsf{H}_{k}\mathsf{H}_{\ell}] = 1862.5$$

where $H_i = 1$ if coin flip *i* is heads and 0 otherwise. Then apply some messy calculations...

• Apply Fourth Moment Bound: $\Pr(|\mathbf{H} - \mathbb{E}[\mathbf{H}]| \ge t) \le \frac{1862.5}{t^4}$. $\frac{25}{t^2}$



Chebyshev's:	4 th Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds?

Chebyshev's:	4 th Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$\Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds? 2G 1900

• Yes! To a point.



Chebyshev's:	4 th Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds?

• Yes! To a point.

• In fact – don't need to just apply Markov's to $|\underline{X} - \mathbb{E}[X]|^k$ for some k. Can apply to any monotonic function $f(|X - \mathbb{E}[X]|)$. $\mathbb{E} f(|X - \mathbb{E}[X]|)$

Chebyshev's:	4 th Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds?

- Yes! To a point.
- In fact don't need to just apply Markov's to |X E[X]|^k for some k. Can apply to any monotonic function f(|X E[X]|).
- Why monotonic?

Chebyshev's:	4 th Moment:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .186$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .0116$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le .0023$	$Pr(H \ge 80) < 10^{-9}$

Can we just keep applying Markov's inequality to higher and higher moments and getting tighter bounds?

F(N) (+14)

In fact – don't need to just apply Markov's to |X – E[X]|^k for some k. Can apply to any monotonic function f(|X – E[X]|).
 To the properties of the

Moment Generating Function: Consider for any *t* > 0:

$$M_{t}(\mathbf{X}) = \underbrace{e^{t \cdot (\mathbf{X} - \mathbb{E}[\mathbf{X}])}}_{\mathbf{X}}$$

Moment Generating Function: Consider for any *t* > 0:

$$M_t(\mathbf{X}) = \underbrace{e^{t \cdot (\mathbf{X} - \mathbb{E}[\mathbf{X}])}}_{k=0} = \sum_{k=0}^{\infty} \underbrace{\frac{t^k (\mathbf{X} - \mathbb{E}[\mathbf{X}])^k}{k!}}_{k!}$$

Moment Generating Function: Consider for any *t* > 0:

$$M_t(\mathsf{X}) = e^{t \cdot (\mathsf{X} - \mathbb{E}[\mathsf{X}])} = \sum_{k=0}^{\infty} \frac{t^k (\mathsf{X} - \mathbb{E}[\mathsf{X}])^k}{k!}$$

• $M_t(\mathbf{X})$ is monotonic for any t > 0.

Moment Generating Function: Consider for any t > 0: $M_t(\mathbf{X}) = e^{t \cdot (\mathbf{X} - \mathbb{E}[\mathbf{X}])} = \sum_{k=0}^{\infty} \frac{t^k (\mathbf{X} - \mathbb{E}[\mathbf{X}])^k}{k!}$

- $M_t(\mathbf{X})$ is monotonic for any t > 0.
- Weighted sum of all moments, with <u>t</u> controlling how slowly the weights fall off (larger t = slower falloff).

12K 12K

Moment Generating Function: Consider for any *t* > 0:

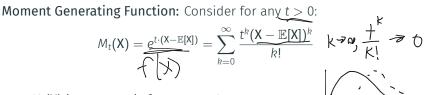
$$M_t(\mathbf{X}) = e^{t \cdot (\mathbf{X} - \mathbb{E}[\mathbf{X}])} = \sum_{k=0}^{\infty} \frac{t^k (\mathbf{X} - \mathbb{E}[\mathbf{X}])^k}{k!}$$

- $M_t(\mathbf{X})$ is monotonic for any t > 0.
- Weighted sum of all moments, with *t* controlling how slowly the weights fall off (larger *t* = slower falloff).
- Choosing <u>t</u> appropriately lets one prove a number of very powerful exponential concentration bounds (exponential tail bounds).

Moment Generating Function: Consider for any t > 0:

$$M_t(\mathbf{X}) = e^{t \cdot (\mathbf{X} - \mathbb{E}[\mathbf{X}])} = \sum_{k=0}^{\infty} \frac{t^k (\mathbf{X} - \mathbb{E}[\mathbf{X}])^k}{k!}$$

- $M_t(\mathbf{X})$ is monotonic for any t > 0.
- Weighted sum of all moments, with *t* controlling how slowly the weights fall off (larger *t* = slower falloff).
- Choosing *t* appropriately lets one prove a number of very powerful exponential concentration bounds (exponential tail bounds).
- Chernoff bound, Bernstein inequalities Hoeffding's inequality, Azuma's inequality, Berry-Esseen theorem, etc.



- $M_t(\mathbf{X})$ is monotonic for any t > 0.
- Weighted sum of all moments, with t controlling how slowly the weights fall off (larger t = slower falloff).
- Choosing *t* appropriately lets one prove a number of very powerful exponential concentration bounds (exponential tail bounds).
- Chernoff bound, Bernstein inequalities, Hoeffding's inequality, Azuma's inequality, Berry-Esseen theorem, etc.
- We will not cover the proofs in this class (although we may in the next problem set).

Bernstein Inequality

$$\begin{array}{c} \{0, 1\} \notin [-1, 1] \\ X = \{0, 1\} \notin [-1, 1] \\ \\ \hline X = \{0, 1\} \\$$

Bernstein Inequality

Bernstein Inequality: Consider independent random variables X_1, \ldots, X_n all falling in [-M, M]. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$ and $\sigma^2 = Var[\sum_{i=1}^n X_i] = \sum_{i=1}^n Var[X_i]$. For any $t \ge 0$: $Pr\left(\left|\sum_{i=1}^n X_i - \mu\right| \ge t\right) \le 2 \exp\left(-\frac{t^2}{2\sigma^2 + \frac{4}{3}Mt}\right)$.

Assume that M = 1 and plug in $t = s \cdot \sigma$ for $s \leq \sigma$.

Bernstein Inequality: Consider independent random variables $X_1, ..., X_n$ all falling in [-1,1]. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$ and $\sigma^2 = Var[\sum_{i=1}^n X_i] = \sum_{i=1}^n Var[X_i]$. For any $s \ge 0$: $Pr\left(\left|\sum_{i=1}^n X_i - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^2}{4}\right)$.

Assume that M = 1 and plug in $t = s \cdot \sigma$ for $s \leq \sigma$.

Bernstein Inequality: Consider independent random variables X_1, \ldots, X_n all falling in [-1,1]. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$ and $\sigma^2 = Var[\sum_{i=1}^n X_i] = \sum_{i=1}^n Var[X_i]$. For any $s \ge 0$: $Pr\left(\left|\sum_{i=1}^n X_i - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^2}{4}\right)$.

Assume that M = 1 and plug in $t = s \cdot \sigma$ for $s \leq \sigma$.

Compare to Chebyshev's: $Pr\left(\left|\sum_{i=1}^{n} X_{i} - \mu\right| \ge s\sigma\right) \le \frac{1}{s^{2}}$.

Bernstein Inequality: Consider independent random variables X_1, \ldots, X_n all falling in [-1,1]. Let $\mu = \mathbb{E}[\sum_{i=1}^n X_i]$ and $\sigma^2 = Var[\sum_{i=1}^n X_i] = \sum_{i=1}^n Var[X_i]$. For any $s \ge 0$: $\Pr\left(\left|\sum_{i=1}^n X_i - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^2}{4}\right).$

Assume that M = 1 and plug in $t = s \cdot \sigma$ for $s \leq \sigma$.

Compare to Chebyshev's: $Pr\left(\left|\sum_{i=1}^{n} X_{i} - \mu\right| \ge s\sigma\right) \le \frac{1}{s^{2}}$.

• An exponentially stronger dependence on s!

Consider again bounding the number of heads H in n = 100 independent coin flips.

Chebyshev's:	Bernstein:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(\underline{H \ge 60}) \le .21$	$Pr(H \ge 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$\Pr(H \ge 70) \le .005$	$Pr(H \ge 70) = .000039$
Pr <u>(H ≥ 80)</u> ≤ .04	$\Pr(H \ge 80) \le 4^{-5}$	$Pr(H \ge 80) < 10^{-9}$

Consider again bounding the number of heads H in n = 100 independent coin flips.

Chebyshev's:	Bernstein:	In Reality:
$Pr(H \ge 60) \le .25$	$Pr(H \ge 60) \le .21$	$Pr(H \geq 60) = 0.0284$
$Pr(H \ge 70) \le .0625$	$Pr(H \ge 70) \le .005$	$Pr(H \ge 70) = .000039$
$Pr(H \ge 80) \le .04$	$Pr(H \ge 80) \le 4^{-5}$	$Pr(H \ge 80) < 10^{-9}$

Getting much closer to the true probability.

Bernstein Inequality (Simplified): Consider independent random variables X_1, \ldots, X_n falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_i]$, $\sigma^2 = \text{Var}[\sum X_i]$, and $s \leq \sigma$. Then:

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^{2}}{4}\right)$$

Bernstein Inequality (Simplified): Consider independent random variables X_1, \ldots, X_n falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_i]$, $\sigma^2 = \text{Var}[\sum X_i]$, and $s \leq \sigma$. Then:

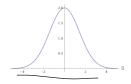
$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2 \exp\left(-\frac{s^{2}}{4}\right).$$

Can plot this bound for different s:

Bernstein Inequality (Simplified): Consider independent random variables X_1, \ldots, X_n falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_i]$, $\sigma^2 = \text{Var}[\sum X_i]$, and $s \leq \sigma$. Then:

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2 \exp\left(-\frac{s^{2}}{4}\right).$$

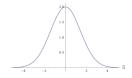
Can plot this bound for different s:



Bernstein Inequality (Simplified): Consider independent random variables X_1, \ldots, X_n falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_i]$, $\sigma^2 = \text{Var}[\sum X_i]$, and $s \leq \sigma$. Then:

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^{2}}{4}\right).$$

Can plot this bound for different s:

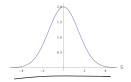


Looks a lot like a Gaussian (normal) distribution.

Bernstein Inequality (Simplified): Consider independent random variables X_1, \ldots, X_n falling in [-1,1]. Let $\mu = \mathbb{E}[\sum X_i]$, $\sigma^2 = \text{Var}[\sum X_i]$, and $s \leq \sigma$. Then:

$$\Pr\left(\left|\sum_{i=1}^{n} \mathbf{X}_{i} - \mu\right| \ge s\sigma\right) \le 2\exp\left(-\frac{s^{2}}{4}\right)$$

Can plot this bound for different s:



Looks a lot like a Gaussian (normal) distribution.

$$\mathcal{N}(0,\sigma^2)$$
 has density $p(s\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{s^2}{2}}$

