Video Stream Quality Impacts Viewer Behavior: Inferring Causality Using Quasi-Experimental Designs

S. Shunmuga Krishnan
Akamai Technologies
sarumuga@akamai.com

Ramesh K. Sitaraman
University of Massachusetts, Amherst
& Akamai Technologies
ramesh@cs.umass.edu

ABSTRACT

The distribution of videos over the Internet is drastically transforming how media is consumed and monetized. Content providers, such as media outlets and video subscription services, would like to ensure that their videos do not fail, startup quickly, and play without interruptions. In return for their investment in video stream quality, content providers expect less viewer abandonment, more viewer engagement, and a greater fraction of repeat viewers, resulting in greater revenues. The key question for a content provider or a CDN is whether and to what extent changes in video quality can cause changes in viewer behavior. Our work is the first to establish a causal relationship between video quality and viewer behavior, taking a step beyond purely correlational studies. To establish causality, we use Quasi-Experimental Designs, a novel technique adapted from the medical and social sciences.

We study the impact of video stream quality on viewer behavior in a scientific data-driven manner by using extensive traces from Akamai’s streaming network that include 23 million views from 6.7 million unique viewers. We show that viewers start to abandon a video if it takes more than 2 seconds to start up, with each incremental delay of 1 second resulting in a 5.8% increase in the abandonment rate. Further, we show that a moderate amount of interruptions can decrease the average play time of a viewer by a significant amount. A viewer who experiences a rebuffer delay equal to 1% of the video duration plays 5% less of the video in comparison to a similar viewer who experienced no rebuffering. Finally, we show that a viewer who experienced failure is 2.32% less likely to revisit the same site within a week than a similar viewer who did not experience a failure.

Categories and Subject Descriptors

C.4 [Performance of Systems]: Measurement techniques, Performance attributes; C.2.4 [Computer-Communication Networks]: Distributed Systems—Client/server

Keywords

Video quality, Internet Content Delivery, User Behavior, Causal Inference, Quasi-Experimental Design, Streaming Video, Multimedia

1. INTRODUCTION

The Internet is radically transforming all aspects of human society by enabling a wide range of applications for business, commerce, entertainment, news and social networking. Perhaps no industry has been transformed more radically than the media and entertainment segment of the economy. As media such as television and movies migrate to the Internet, there are twin challenges that content providers face whose ranks include major media companies (e.g., NBC, CBS), news outlets (e.g., CNN), sports organizations (e.g., NFL, MLB), and video subscription services (e.g., Netflix, Hulu).

The first major challenge for content providers is providing a high-quality streaming experience for their viewers, where videos are available without failure, they startup quickly, and stream without interruptions [24]. A major technological innovation of the past decade that allows content providers to deliver higher-quality video streams to a global audience of viewers is the content delivery network (or, CDN for short) [8, 17]. CDNs are large distributed systems that consist of hundreds of thousands of servers placed in thousands of ISPs close to end users. CDNs employ several techniques for transporting [2, 12] media content from the content provider’s origin to servers at the “edges” of the Internet where they are cached and served with higher quality to the end user. (See [17] for a more detailed description of a typical CDN architecture.)

The second major challenge of a content provider is to actually monetize their video content through ad-based or subscription-based models. Content providers track key metrics of viewer behavior that lead to better monetization. Primary among them relate to viewer abandonment, engagement, and repeat viewership. Content providers know that reducing the abandonment rate, increasing the play time of each video watched, and enhancing the rate at which viewers return to their site increase opportunities for advertising and upselling, leading to greater revenues. The key question is whether and by how much increased stream quality can cause changes in viewer behavior that are conducive to improved monetization. Relatively little is known from a scientific standpoint about the all-important causal link between video stream quality and viewer behavior for online media. Exploring the causal impact of quality on behavior...
and developing tools for such an exploration are the primary focus of our work.

While understanding the link between stream quality and viewer behavior is of paramount importance to the content provider, it also has profound implications for how a CDN must be architected. An architect is often faced with trade-offs on which quality metrics need to be optimized by the CDN. A scientific study of which quality metrics have the most impact on viewer behavior can guide these choices. As an example of viewer behavior impacting CDN architecture, we performed small-scale controlled experiments on viewer behavior a decade ago that established the relative importance of the video to startup quickly and play without interruptions. These behavioral studies motivated an architectural feature called prebursting [12] that was deployed on Akamai’s live streaming network that enabled the CDN to deliver streams to a media player at higher than the encoded rate for short periods of time to fill the media player’s buffer with more data more quickly, resulting in the stream starting up faster and playing with fewer interruptions. It is notable that the folklore on the importance of startup time and rebuffering were confirmed in two recent important large-scale scientific studies [9, 15]. Our current work sheds further light on the important nexus between stream quality and viewer behavior and, importantly, provides the first evidence of a causal impact of quality on behavior.

1.1 Measuring quality and viewer behavior

The advent of customizable media players supporting major formats such as Adobe Flash, Microsoft Silverlight, and Apple HTTP streaming has revolutionized our ability to perform truly large-scale studies of stream quality and viewer behavior as we do in this work, in a way not possible even a few years ago. It has become possible to instrument media players with an analytics plug-in that accurately measures and reports both quality and behavioral metrics from every viewer on a truly planetary scale.

1.2 From correlation to causality

The ability to measure stream quality and viewer behavior on a global scale allows us to correlate the two in a statistically significant way. For each video watched by a viewer, we are able to measure its quality including whether the stream was available, how long the stream took to start up, and how much rebuffering occurred causing interruptions. We are also able to measure the viewer’s behavior including whether he/she abandoned the video and how long he/she watched the video.

As a first step, we begin by simply correlating important quality metrics experienced by the viewers to the behavior that they exhibit. For instance, we discover a strong correlation between an increase in the delay for the video to start up and an increase in rate at which viewers abandon the video. Several of our results are the first quantitative demonstration that certain key streaming quality metrics are correlated with key behavioral metrics of the viewer. However, the deeper question is not just whether quality and behavior are correlated but whether quality can causally impact viewer behavior. While correlation is an important first step, correlation does not necessarily imply causality. The holy grail of a content provider or a CDN architect is to discover causal relationships rather than just correlational ones, since they would like to know with some certainty that the significant effort expended in improving stream quality will in fact result in favorable viewer behavior.

In fact, a purely correlational relationship could even lead one astray, if there is no convincing evidence of causality, leading to poor business decisions. For instance, both video quality (say, video bitrates) and viewer behavior (say, play time) have been steadily improving over the past decade and are hence correlated in a statistical sense. But, that fact alone is not sufficient to conclude that higher bitrates cause viewers to watch longer, unless one can account for other potential “confounding” factors such as the available video content itself becoming more captivating over time.

While inferring causality is generally difficult, a key tool widely used in the social and medical sciences to infer causality from observational data is a Quasi Experimental Design (QED) [23]. Intuitively, a QED is constructed to infer if a particular “treatment” (i.e., cause) results in a particular “outcome” (i.e., effect) by pairing each person in the observational data who has had treatment with a random untreated person who is “significantly identical” to the treated person in all other respects. Thus, the pairing eliminates the effect of the hidden confounding variables by ensuring that both members of a pair have sufficiently identical values for those variables. Thus, evaluating the differential outcomes between treated and untreated pairs can either strengthen or weaken a causal conclusion that the treatment causally impacts the outcome. While it is impossible to completely eliminate all hidden factors, our causal analysis using QEDs should be viewed as strengthening our correlational observations between treatments and outcomes by eliminating the common threats to a causal conclusion.

1.3 Our Contributions

Our study is one of the largest of its kind of video stream quality and viewer behavior that collects and analyzes a data set consisting of more than 23 million video playbacks from 6.7 million unique viewers who watched an aggregate of 216 million minutes of 102 thousand videos over 10 days.

To our knowledge, our work is the first to provide evidence that video stream quality causally impacts viewer behavior, a conclusion that is important to both content providers and CDNs. Further, our adaptation of Quasi-Experimental Designs (QEDs) is an unique contribution and is of independent interest. QEDs have been used extensively in medical research and the social sciences in the past decades. We expect that our adaptation of QEDs for measurement research in networked systems could be key in a variety of other domains that have so far been limited to correlational studies.

Our work is also the first to quantitatively explore viewer abandonment rates and repeat viewership in relation to stream quality, last-mile connectivity, and video duration. In addition, we also study viewer engagement (e.g., play time) in relation to stream quality (e.g., rebuffering) that has also been recently studied in [9] in a correlational setting, but we take a step beyond correlational analysis to establish a causal relationship between quality and engagement using QEDs. Our work makes the following specific contributions on the impact stream quality on viewer behavior.

- We show that an increase in the startup delay beyond 2 seconds causes viewers to abandon the video. Using regression, we show that an additional increase of the startup delay by 1 second increases the abandonment rate by 5.8%.
- Viewers are less tolerant to startup delay for a short video such as news clip than a long video such an hour-long TV episode. In a quasi experiment, the likelihood of a viewer of short video abandoning earlier than a similar viewer of a long video exceeded the likelihood that the opposite happens by 11.5%.
- Viewers watching video on a better connected computer or device have less patience for startup delay and abandon sooner. In particular, viewers on mobile devices have the most patience and abandon the least, while those on fiber-based broadband abandon the soonest. In a quasi experiment, the likelihood that a viewer on fiber abandoned earlier than a similar viewer on a mobile device exceeded the likelihood that the opposite happens by 38.25%.
- Viewers who experienced an increase in the normalized rebuffer delay, i.e., they experienced more interruptions in the video, played the video for lesser time. In a quasi experiment, a viewer who experienced a rebuffer delay that equals or exceeds 1% of the video duration played 5.02% less of the video in comparison to a similar viewer who experienced no rebuffering.
- A viewer who experienced a failed visit is less likely to return to the content provider’s site to view more videos within a specified time period than a similar viewer who did not experience the failure. The probability that a viewer who experienced failure returns to the content provider’s site within a week is less than the likelihood of a similar viewer who did not experience failures by 2.32%.

We show that above results are statistically significant using the sign test. Further, these results show a significant level of causal impact of stream quality on viewer behavior. In this regard, it is important to recall that small changes in viewer behavior can lead to large changes in monetization, since the impact of a few percentage points over tens of millions of viewers can accrue to large impact over a period of time. Therefore, our work on deriving a causal relationship by systematically accounting for the confounding variables must not be viewed as a definitive proof of causality, as indeed there can be no definitive proof of causality. But, rather, our work significantly increases the confidence in a causal conclusion by eliminating the effect of major confounding factors that could threaten such a conclusion.

2. BACKGROUND

We describe the process of a user watching a stream, defining terms along the way that we will use in this paper.

**Viewer.** A viewer is a user who watches one or more streams using a specific media player installed on the user’s device. A viewer is uniquely identified and distinguished from other viewers by using a GUID (Globally Unique Identifier) value that is set as a cookie when the media player is accessed. To identify the viewer uniquely, the GUID value is generated to be distinct from other prior values in use.

**Views.** A view represents an attempt by a viewer to watch a specific video stream. A typical view would start with the viewer initiating the video playback, for instance, by clicking the play button of the media player\(^1\) (see Figure 1). During a view, the media player begins in the startup state where it connects to the server and downloads a certain specified amount of data, before transitioning to the play state. In the play state, the player uses the data from its buffer and renders the video on the viewer’s screen. Meanwhile, the player continues to download data from the server and stores it in the buffer. Poor network conditions between the server and the player could lead to a situation where the buffer is drained faster than it is being filled. This could lead to a condition where the buffer is empty causing the player to enter the rebuffer state where the viewer experiences an interruption or “freeze” in the video playback. While in the rebuffer state, the player continues to fill its buffer from the server. When the buffer has a specified amount of data, the player enters the play state and the video starts to play again (see Figure 1). A view can end in three ways: a **successful view** ends normally when the video completes; a **failed view** ends with a failure or error due to a problem with the server, network, or content; and, finally, an **abandoned view** ends with the viewer voluntarily abandoning the stream either before the video starts up or after watching some portion of it. Note that a viewer may abandon the view by closing the browser, stopping the stream, or clicking on a different stream. There are other secondary viewer-initiated events or viewer-initiated events that are part of the viewing process. For instance, a viewer could initiate actions such as pausing, fast-forwarding, or rewinding the video stream. Further, the player may switch the bitrate of the encoded media in response to network conditions, such as reduce the bandwidth if there is packet loss. We do not explicitly analyze behaviors associated with these secondary events in this paper, though these could be part of future work.

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\(^1\)For some content providers, a “pre-roll” advertisement is shown before the actual content video is requested by the media player. In that case, our view starts at the point where the actual video is requested.

![Figure 1: Views and Visits.](image-url)
<table>
<thead>
<tr>
<th>Key Metrics</th>
<th>Definition</th>
</tr>
</thead>
<tbody>
<tr>
<td>Failures</td>
<td>Number (or, percentage) of views that fail due to problems with the network, server, or content.</td>
</tr>
<tr>
<td>Startup Delay</td>
<td>Total time in startup state.</td>
</tr>
<tr>
<td>Average Bitrate</td>
<td>The average bitrate at which the video was watched.</td>
</tr>
<tr>
<td>Normalized Re-buffer Delay</td>
<td>Total time in rebuffer state divided by the total duration of the video.</td>
</tr>
</tbody>
</table>

**Figure 2: View-level stream quality metrics.**

the stream was available or if viewing of the video initiated by the viewer failed due to a problem with the network or the server or the content itself (such as a broken link). A failed view can be frustrating to the viewer as he/she is unable to watch the video. A second key metric is startup delay which is the amount of time the viewer waits for the video to start up. Once the video starts playing, the average bitrate at which the video was rendered on the viewer’s screen is a measure of the richness of the presented content. This metric is somewhat complex since it is a function of how the video was encoded, the network connectivity between the server and the client, and the heuristics for bitrate-switching employed by the player. Finally, a fourth type of metric quantifies the extent to which the viewer experienced rebuffering. Rebuffering is also frustrating for the viewer because the video stops playing and “freezes”. We can quantify the rebuffering by computing the rebuffer delay which is total time spent in a rebuffer state and normalizing it by dividing it by the duration of the video.

Many of the above view-level metrics can be easily extended to visit-level or viewer-level metrics. One key visit-level metric that we examine in this paper is a failed visit which is a visit that ends with a failed view. A failed visit could have had successful views prior to the failed view(s). However, a failed visit is important because the viewer tries to play a video one or more times but is unable to do so and leaves the site right after the failure, presumably with a level of frustration.

In our paper, we use many of the key metrics in Figure 2 in our evaluation of the impact of quality on viewer behavior, though these are by no means the only metrics of stream quality. It should be noted that many of the above metrics were incorporated into measurement tools within Akamai and have been in use for more than a decade [24, 1]. The lack of client-side measurements in the early years led to measurements based on automated “agents” deployed around the Internet that simulated synthetic viewers [1, 19] that were then supplemented with server-side logs. In recent years, there is a broad consensus among content providers, CDNs, and analytics providers that these metrics or variations of these metrics matter.

**Metrics for Viewer Behavior.** Our metrics are focused on the key aspects of viewer behavior that are often tracked closely by content providers which we place in three categories (see Figure 3). The first category is abandonment where a viewer voluntarily decides to stop watching the video. Here we are primarily concerned with abandonment where the viewer abandons the video even before it starts playing. A viewer can also abandon a stream after watching a portion of the video which results in a smaller play time, which we account for in the next category of metrics. The second category is viewer engagement that can be measured by play time which is simply the amount of video that the viewer watches. The final category speaks to the behavior of viewers over longer periods of time. A key metric is the return rate of viewers measured as the probability that a viewer returns to the content provider’s site over period of time, say, returning within a day or returning within a week.

### 3. DATA SETS

The data sets that we use for our analysis are collected from a large cross section of actual users around the world who play videos using media players that incorporate the widely-deployed Akamai’s client-side media analytics plug in3. When content providers build their media player, they can choose to incorporate the plugin that provides an accurate means for measuring a variety of stream quality and viewer behavioral metrics. When the viewer uses the media player to play a video, the plugin is loaded at the client-side and it “listens” and records a variety of events that can then be used to stitch together an accurate picture of the playback. For instance, player transitions between the startup, rebuffering, seek, pause, and play states are recorded so that one may compute the relevant metrics. Properties of the playback, such as the current bitrate, bitrate switching, state of the player’s data buffer are also recorded. Further, viewer-initiated action that lead to abandonment such as closing the browser or browser tab, clicking on a different link, etc can also be accurately captured. Once the metrics are captured by the plugin, the information is “beaconed” to an analytics backend that can process huge volumes of data. From every media player at the beginning and end of every view, the relevant measurements are sent to the analytics backend. Further, incremental updates are sent at a configurable periodicity even as the video is playing.

#### 3.1 Data Characteristics

While the Akamai platform serves a significant amount of the world’s enterprise streaming content accounting for several million concurrent views during the day, we choose a smaller but representative slice of the data from 12 content providers that include major enterprises in a variety of verticals including news, entertainment, and movies. We consider only on-demand videos in this study, leaving live videos for future work. We tracked the viewers and views for the chosen content providers for a period of 10 days (see Figure 4). Our data set is extensive and captures 23 million

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3While all our data is from media players that are instrumented with Akamai’s client-side plugin, the actual delivery of the streams could have used any platform and not necessarily just Akamai’s CDN.
<table>
<thead>
<tr>
<th>Total</th>
<th>Avg Per Visit</th>
<th>Avg Per Viewer</th>
</tr>
</thead>
<tbody>
<tr>
<td>Views</td>
<td>23 million</td>
<td>2.39</td>
</tr>
<tr>
<td>Minutes</td>
<td>216 million</td>
<td>22.48</td>
</tr>
<tr>
<td>Videos</td>
<td>102 thousand</td>
<td>1.96</td>
</tr>
<tr>
<td>Bytes</td>
<td>1431 TB</td>
<td>148 MB</td>
</tr>
</tbody>
</table>

Figure 4: Summary of views, minutes watched, distinct videos, and bytes downloaded for our data set.

<table>
<thead>
<tr>
<th>Viewer Geography</th>
<th>Percent Views</th>
</tr>
</thead>
<tbody>
<tr>
<td>North America</td>
<td>78.85%</td>
</tr>
<tr>
<td>Asia</td>
<td>12.80%</td>
</tr>
<tr>
<td>Europe</td>
<td>7.75%</td>
</tr>
<tr>
<td>Other</td>
<td>0.60%</td>
</tr>
</tbody>
</table>

Figure 5: The geography of viewers in our trace at the continent-level.

views from 6.7 million unique viewers, where each viewer on average made 3.42 visits over the period and viewed a total of 32.2 minutes of video. In each visit, there were on average 2.39 views but only 1.96 unique videos viewed, indicating that sometimes the viewer saw the same video twice.

The geography of the viewer was mostly concentrated in North America, Europe and Asia with small contributions from other continents (see Figure 5). More than half the views used cable, though fiber, mobile, and DSL were significant. The fiber category consisted mostly of AT&T Uverse and Verizon FiOS that contributed in roughly equal proportion. The other connection types such as dialup were negligible (see Figure 6). Video duration is the total length (in minutes) of the video (See Figure 7). We divide the videos into short that have a duration of less than 30 minutes and long that have a duration of more than 30 minutes. Examples of short video include news clips, highlight reels for sports, and short television episodes. The median duration was 1.8 minutes, though the mean duration was longer at 5.95 minutes. In contrast, long video consists of long television episodes and movies. The median duration for long videos was 43.2 minutes and the mean was 47.8 minutes.

4. ANALYSIS TECHNIQUES

A key goal is to establish a causal link between a stream quality metric X and viewer behavior metric Y. The first key step is to establish a correlational link between X and Y using the statistical tools for correlation and regression. Next, in accordance with the maxim that “correlation does not imply causation”, we do a more careful analysis to establish causation. We adapt the innovative tool of Quasi Experimental Design (QED) used extensively in the social and medical sciences to problem domains such as ours.

4.1 Correlational Analysis

A key goal is to establish a causal link between a stream quality metric X (say, startup delay) with a viewer behavioral metric Y (say abandonment rate), we start by visually plotting metric X versus metric Y in the observed data. The visual representations are a good initial step to estimating whether or not a correlation exist. As a next step, we also quantify the correlation between X and Y. There are many different ways to calculate the correlation. Primary among them are Pearson’s correlation and Kendall’s correlation that is a type of rank correlation. As observed in [9], Kendall’s correlation is more suitable for a situation such as ours since it does not assume any particular distributional relationship between the two variables. Pearson’s correlation is more appropriate when the correlated variables are approximately linearly related, unlike the relationships that we explore in our work. Kendall’s correlation measures the whether the two variables X and Y are statistically dependent (i.e., correlated) without assuming any specific functional form of their relationship. Kendall’s correlation coefficient τ takes values in the interval $[-1, 1]$ where $\tau = 1$ meaning that X and Y are perfectly concordant, i.e., larger values of X are always associated with larger values for Y, $\tau = -1$ meaning that X and Y are perfectly discordant, i.e., larger values of X are always associated with smaller values of Y, and $\tau$ near 0 implying that X and Y are independent.

4.2 Causal Analysis

A correlational analysis of stream quality metric X (say,
starup delay) and a viewer behavior metric \( Y \) (say, abandonment rate) could show that \( X \) and \( Y \) are associated with each other. A primary threat to a causal conclusion that an independent variable \( X \) causes the dependent variable \( Y \) is the existence of confounding variables that can impact both \( X \) and \( Y \) (see Figure 8). To take a recent example from the medical literature, a study published in Nature [20] made the causal conclusion that children who sleep with the light on are more likely to develop myopia later in life. But, as it turns out, myopic parents tend to leave the light on more often, as well as pass their genetic predisposition to myopia to their children. Accounting for the confounding variable of parent’s myopia, the causal results were subsequently invalidated or substantially weakened.

More relevant to our own work, lets consider a potential threat to a causal conclusion that a stream quality metric \( X \) (say, startup delay) results in a viewer behavior \( Y \) (say, abandonment). As a hypothetical example, suppose that mobile users tend to have less patience for videos to startup as they tend to be busy and are “on the go”, resulting in greater abandonment. Further assume that mobile users tend to have larger startup delays due to poor wireless connectivity. In this situation, a correlation between startup delay and abandonment may not imply causality unless we can account for the confounding variable of how the viewer is connected to the Internet. In our causal analyses, we systematically identify and account for all or a subset of the following three categories of confounding variables as relevant (see Figure 8).

**Content**. The video\(^4\) being watched could itself influence both quality and viewer behavior. For instance, some videos are more captivating than others leading viewers to watch more of it. Or, some videos may have higher perceived value than others, leading viewers to tolerate more startup delay. The manner in which the video is encoded and the player heuristic used by the media player could also impact stream quality. For instance, the player heuristics that could differ from one content provider to another specifies how much of the video needs to be buffered before the stream can startup or resume play after rebuffering.

**Connection Type**. The manner in which a viewer connects to the Internet, both the device used and typical connectivity characteristics can influence both stream quality and viewer behavior. We use the connection type of the viewer as a confounding variable, where the connection type can take discrete values such as mobile, DSL, cable, and fiber (such as AT&T’s Uverse and Verizon’s FIOS).

**Geography**. Geography of viewer captures several social, economic, religious, and cultural aspects that can influence viewer behavior. For instance, it has been observed by social scientists that the level of patience that consumers exhibit towards a delay in receiving a product varies based on geography of the consumer [5]. Such a phenomena might well be of significance in the extent to which the viewer’s behavior is altered by stream quality. In our work, we analyze viewer’s geography at the granularity of a country.

### 4.2.1 The Quasi Experimental Design (QED) Method

A primary technique for showing that an independent variable \( X \) (called the treatment variable) has a causal impact on a dependent variable \( Y \) (called the outcome variable) is to design a controlled experiment. To design a true experiment in our context, one would have to randomly assign viewers to differing levels of stream quality (i.e., values of \( X \)) and observe the resultant viewer behaviors (values of \( Y \)). The random assignment in such an experiment removes any systematic bias due to the confounding variables that are threats to a causal conclusion. However, the level of control needed to perform such an experiment at scale for our problem is either prohibitively hard, expensive, or even impossible. In fact, there are legal, ethical, and other issues with intentionally degrading the stream quality of a set of viewers to do a controlled experiment. However, there are other domains were a controlled experiment can and are performed, example, A/B testing of web page layouts [11].

Given the inability to perform true experiments, we adapt a technique called QED to discover causal relationships from observational data that already exists. QEDs were developed by social and medical scientists as a similar inability to perform controlled experiments is very common in those domains [23]. In particular, we use a specific type of QED called the matched design [18] where a treated individual (in our case, a view or viewer) is randomly matched with an untreated individual, where both individuals have identical values for the confounding variables. Consequently, any difference in the outcome for this pair can be attributed to the treatment. Our population typically consists of views or viewers and treatment variable is typically is binary. For instance, in Section 7, viewers who experienced “bad” stream quality in the form a failed visit are deemed to be treated and viewers who had normal experience are untreated. We form comparison sets by randomly matching each treated viewer with an untreated viewer such that both viewers are as identical as possible on the confounding variables. Needless to say, the more identical the viewers are in each pair.

\(^4\) Note that our notion of video content is url-based and thus also incorporates the content provider. If the same movie is available from two content providers, they would constitute two different pieces of content for our analysis.
the more effective the matching is in neutralizing the con-
foundling variables. Note that matching ensures that the dis-
tributions of the confounding variables in the treated and
untreated set of viewers are identical, much as if viewers were
randomly assigned to treated and untreated sets in a con-
trolled experiment. Now, by studying the behavioral out-
comes of matched pairs one can deduce whether or not the
treatment variable X has a causal effect on variable Y, with
the influence of the confounding variables neutralized. Note
that treatment variable need not always be stream quality.
Depending on the causal conclusion, we could choose the
treatment variable to content length or connection type, if
we would like to study their impact on viewer behavior.

Statistical Significance of the QED Analysis.
As with any statistical analysis, it is important to evalu-
ate whether the results are statistically significant or if they
could have occurred by random chance. As is customary in
hypothesis testing [14], we state a null hypothesis \( H_0 \) that
contradicts the assertion that we want establish. That is,
\( H_0 \) contradicts the assertion that X impacts Y and states
that the treatment variable X has no impact on the outcome
variable Y. We then compute the “p-value” defined to be the
probability that the null hypothesis \( H_0 \) is consistent with the
observed results. A “low” p-value lets us reject the null hy-
pothesis, bolstering our conclusions from the QED analysis
as being statistically significant. However, a “high” p-value
would not allow us to reject the null hypothesis. That is, the
QED results could have happened through random chance
with a “sufficiently” high probability that we cannot reject
\( H_0 \). In this case, we conclude that the results from the QED
analysis are not statistically significant.

The definition of what constitutes a “low” p-value for a
result to be considered statistically significant is somewhat
arbitrary. It is customary in the medical sciences to con-
clude that a treatment is effective if the p-value is at most
0.05. The choice of 0.05 as the significance level is largely
cultural and can be traced back to the classical work of R.A.
Fisher about 90 years ago. Many have recently argued that
the significance level must be much smaller. We concur and
choose the much more stringent 0.001 as our significance
level, a level achievable in our field given the large amount of
experimental subjects (tens of thousands treated-untreated
pairs) but is rarely achievable in medicine with human sub-
jects (usually in the order of hundreds of treated-untreated
pairs). However, our results are unambiguously significant
and not very sensitive to the choice of significance level. All
our results turned out to be highly significant with p-values
of \( 4 \times 10^{-5} \) or smaller, except for one conclusion with a larger
p-value that we deemed statistically insignificant.

The primary technique that we employ for evaluating sta-
tistical significance is the sign test that is a non-parametric
test that makes no distributional assumptions and is par-
ticularly well-suited for evaluating matched pairs in a QED
setting [26]. We sketch the intuition of the technique here,
while deferring the specifics to the technical sections. For
each matched pair \((u, v)\), where \( u \) received treatment and
\( v \) did not receive treatment, we define the differential out-
come denoted by \( outcome(u, v) \) as the numerical difference
in the outcome of \( u \) and the outcome of \( v \). If \( H_0 \) holds,
then the outcomes of the treated and untreated individuals
are identically distributed, since the treatment is assumed
to have no impact on the outcome. Thus, the differential
outcome is equally likely to be a positive number as a neg-
ative number. Thus, for \( n \) independently selected matched
pairs, the number of positive values of the differential out-
come (call it \( X \)) follows the binomial distribution with \( n \)
trials and probability 1/2. In a measured sample consisting of
a total of \( n \) non-zero values of the differential outcome,
suppose that \( x \) have positive values. Given that \( H_0 \) holds,
the probability (i.e., p-value) of such an occurrence is at
most \( \operatorname{Prob}(|X − n/2| ≥ | x − n/2|) \), which is sum of both
tails of the binomial distribution. Evaluating the above tail
probability provides us the required bound on the p-value.

As an aside, note that a different but distribution-specific
significance test called the paired t-test may be applicable in
other QED situations. A paired t-test uses the Student’s t
distribution and requires that the differential outcome has a
normal distribution. Since our differential outcome does not
have a normal distribution, we rely on the distribution-free
non-parametric sign test that is more generally applicable.

Some Caveats.
It is important to understand the limitations of our QED
tools, or for that matter any experimental technique of infer-
ence. Care should be taken in designing the quasi-experiment
to ensure that the major confounding variables are explicitly
or implicitly captured in the analysis. If there exists con-
founding variables that are not easily measurable (example,
the gender of the viewer) and/or are not identified and con-
trolled, these unaccounted dimensions could pose a risk to a
causal conclusion, if indeed they turn out to be significant.

Our work on deriving a causal relationship by systematically
accounting for the confounding variables must not be viewed
as a definitive proof of causality, as indeed there can be no
definitive proof of causality. But, rather, our work increases
the confidence in a causal conclusion by accounting for po-
tential major sources of confounding. This is of course a
general caveat that holds for all domains across the sciences
that attempt to infer causality from observational data.

5. VIEWER ABANDONMENT
We address the question of how long a viewer will wait
for the stream to start up, a question of great importance
that has not been studied systematically to our knowledge.
However, the analogous problem of how long a user will wait
for web content to download has received much attention.
In 2006, Jupiter Research published a study based on inter-
viewing 1,058 online shoppers and postulated what is known
in the industry as the “4-second rule” that states that an av-
gerage online shopper is likely to abandon a web site if a
web page does not download in 4 seconds [21]. But, a recent
study [16] implied that the users have become impatient over
time and that even a 400 ms delay can make users search
less. Our motivation is to derive analogous rules for stream-
ing where startup delay for video is roughly analogous to
download time for web pages.

Assertion 5.1. An increase in startup delay causes more
abandonment of viewers.

To investigate if our assertion holds, we classify each view
into 1-second buckets based on their startup delay. We then
compute for each bucket the percentage of views assigned
to that bucket that were abandoned. From Figure 9, we see
that the percent of abandoned views and startup delay are
positively correlated with a Kendall correlation of 0.72.
Suppose now that we build a media delivery service that provides a startup delay of exactly \( x \) seconds for every view. What percent of views delivered by this system will be abandoned? To estimate this metric, we define a function called \( f_{\text{abandonment}}(x) \) that equals

\[
100 \times \frac{\text{Impatient}(x)}{\text{Impatient}(x) + \text{Patient}(x)},
\]

where \( \text{Impatient}(x) \) is all views that were abandoned after experiencing less than \( x \) seconds of startup delay and \( \text{Patient}(x) \) are views where the viewer waited at least \( x \) time without abandoning. That is, \( \text{Impatient}(x) \) (resp., \( \text{Patient}(x) \)) corresponds to views where the viewer did not (resp., did) demonstrate the patience to hold on for \( x \) seconds without abandoning. Note that a view in \( \text{Patient}(x) \) could still have been abandoned at some time greater than \( x \). Also, note that a view where the video started to play before \( x \) seconds does not provide any information on whether the viewer would have waited until \( x \) seconds or not, and so is considered neither patient or impatient. Figure 10 shows the abandonment rate computed from our data which is near zero for the first \( T \) seconds, but starts to rise rapidly as the startup delay increases. Fitting a simple regression to initial part of the curve shows that abandonment rate increases by \( 0.58 \) for each 1-second increase in startup delay.

**Assertion 5.2.** Viewers are less tolerant of startup delay for short videos in comparison to longer videos.

Researchers who study the psychology of queuing [13] have shown that people have more patience for waiting in longer queues if the perceived value of the service that they are waiting for is greater. Duration of the service often influences its perceived value with longer durations often perceived as having greater value. People often tolerate the 30 minute delay for the checkin process for a 4-hour plane ride but would find the same wait excessive for a 10-minute bus ride. On the same principle, is it true that viewers would be more patient for the video to startup if they expect to be watching the video for longer period of time?

To investigate our assertion, we first classify the views based on whether the content is short with duration smaller than 30 minutes (e.g., news clip), or long with duration longer than 30 minutes (e.g., movies). The Kendall correlation between the two variables, percent of abandoned videos and startup delay, were 0.68 and 0.90 for short and long videos respectively, indicating a strong correlation for each category. Further, Figure 11 shows abandonment rate for each type of content as a function of the startup delay. One can see that viewers typically abandon at a larger rate for short videos than for long videos.

**Assertion 5.3.** Viewers watching videos on a better connected computer or device have less patience for startup delay and so abandon sooner.

The above assertion is plausible because there is some evidence that users who expect faster service are more likely to be disappointed when that service is slow. In fact, this is often touted as a reason for why users are becoming less and less able to tolerate web pages that download slowly. To study whether or not this is true in a scientific manner, we segment our views into four categories based on their connection type that indicates how the corresponding viewer is connected to the Internet. The categories in roughly the increasing order of connectivity are mobile, DSL, cable modem, and fiber (such as Verizon FIOS or AT&T Uverse). In all four categories, we see a strong correlation between the two variables, percent of abandoned views and startup delay. The Kendall correlations for mobile, DSL, cable, and fiber are 0.68, 0.74, 0.71, and 0.75 respectively. Further, in Figure 12, we show the abandonment rate for each connection type. We can see that viewers abandon significantly less on mobile in comparison with the other categories, for
Figure 11: Viewers abandon at a higher rate for short videos than for long videos.

a given startup delay. Some difference in abandonment is discernible between the other categories in the rough order of cable, DSL, and fiber though they are much smaller.

5.1 QED for Assertion 5.2

First, we devise a QED to study the impact of content length on abandonment (Assertion 5.2). Therefore, we make the content length (long or short) the treatment variable and the outcome measures patience of the viewer to startup delay. The viewer’s patience to startup delay can be influenced by both the viewer’s geography and connection type which we use as the confounding variables. Specifically, we form matched pairs \((u, v)\) such that view \(u\) is a short video that was abandoned and view \(v\) is a long video that was abandoned and \(u\) and \(v\) are watched by viewers from the same geography, and the viewers have the same connection type. The matching algorithm is described as follows.

1. **Match step.** Let the treated set \(T\) be all abandoned views for short content and let untreated set \(C\) be all the abandoned views for long content. For each \(u \in T\) we pick uniformly and randomly a \(v \in C\) such that \(u\) and \(v\) belong to viewers in the same geography and have the same connection type. The matched set of pairs \(M \subseteq T \times C\) have the same attributes for the confounding variables and differ only on the treatment.

2. **Score step.** For each pair \((u, v) \in M\), we compute an outcome \(\text{outcome}(u, v)\) to be \(+1\) if \(u\) was abandoned with a smaller startup delay than \(v\). If \(u\) was abandoned with a larger startup delay than \(v\), then \(\text{outcome}(u, v) = -1\). And, \(\text{outcome}(u, v) = 0\), if the startup delays when \(u\) and \(v\) were abandoned are equal. Now,

\[
\text{Net Outcome} = \left( \sum_{(u,v) \in M} \text{outcome}(u,v) \right) / |M| \times 100.
\]

Figure 12: Viewers who are better connected abandon sooner.

Note that a positive value for net outcome provides positive (supporting) evidence for Assertion 5.3, while a negative value provides negative evidence for the assertion. The results of the matching algorithm produced a net outcome of 11.5%. The net outcome shows that the matched pairs that support Assertion 5.2 exceed those that negate the assertion by 11.5%. The positive net outcome provides evidence of causality that was not provided by the prior correlational analysis alone by eliminating the threats posed by the identified confounding variables.

To derive statistical significance of the above QED result, we formulate a null hypothesis \(H_0\) that states that the treatment (long versus short video) has no impact on abandonment. If \(H_0\) holds, the outcome \(\text{outcome}(u, v)\) is equally likely to be positive (+1) as negative (-1). We now use the sign test that we described in Section 4.2 to derive a bound on the p-value. Since we matched \(n = 78,840\) pairs, if \(H_0\) holds, the expected number pairs with a positive outcome is \(n/2 = 78,840/2 = 39,420\). Our observational data however had \(x = 43,954\) pairs with positive scores, i.e., \(x - n/2 = 4534\) pairs in excess of the mean. We bound the p-value by showing that it is extremely unlikely to have had 4534 positive pairs in excess of the mean by computing the the two-sided tail of the binomial distribution with \(n\) trials and probability 1/2:

\[
p\text{-value} \leq \text{Prob}\left(|X - n/2| \geq x - n/2\right) \leq 3.3 \times 10^{-229} \quad (1)
\]

The above bound for the p-value is much smaller than the required significance level of 0.001 and leads us to reject the null hypothesis \(H_0\). Thus, we conclude that our QED analysis is statistically significant.

5.2 QED for Assertion 5.3

To investigate a causal conclusion for Assertion 5.3, we set up a QED where the treatment is the connection type of the user and the outcome measures the relative tolerance of
the viewer to startup delay. For each pair of network types $A$ and $B$, we run a matching algorithm where the treated set $T$ is the set of all abandoned views with connection type $A$ and untreated set is all abandoned views with connection type $B$. The matching algorithm used is identical to the one described earlier except that the match criterion in step 1 is changed to match for identical content and identical geography. That is, for every matched pair $(u, v)$, view $u$ has network type $A$ and view $v$ has network type $B$ but both are views for the same video and belong to viewers in the same geography.

The results of the matching algorithm are shown in Figure 13. For instance, our results show that the likelihood that a mobile viewer exhibited more patience than a fiber viewer is greater than the likelihood that opposite holds by a margin of 38.25\%. Much as in Section 5.1, we use the sign test to compute the p-value for each QED outcome in the table. All QED outcomes in Figure 13 turned out to be statistically significant with exceedingly small p-values, except the dsl-versus-cable comparison that was inconclusive. Specifically, our results show that a that a mobile viewer exhibits more patience than other (non-mobile) viewers, and the result holds with exceedingly small p-values ($< 10^{-17}$). Our results also provide strong evidence for DSL and cable users being more patient than fiber users, though the p-value for the dsl-versus-fiber was somewhat larger ($4.6 \times 10^{-5}$) but still statistically significant. The dsl-versus-cable comparison was however inconclusive and not statistically significant as the p-value of the score was 0.06 that is larger than our required significance level of 0.001.

Figure 13: Net QED outcomes support the causal impact of connection type on viewer patience for startup delay, though the impact is more pronounced between mobile and the rest. The p-value for all entries are very small ($< 10^{-17}$), except dsl-versus-cable (0.06) and dsl-versus-fiber ($4.6 \times 10^{-5}$).

<table>
<thead>
<tr>
<th>Treatment</th>
<th>Untreated</th>
<th>dsl</th>
<th>cable</th>
<th>fiber</th>
</tr>
</thead>
<tbody>
<tr>
<td>mobile</td>
<td>33.81</td>
<td>35.40</td>
<td>38.25</td>
<td></td>
</tr>
<tr>
<td>dsl</td>
<td>-</td>
<td>-0.75</td>
<td>2.67</td>
<td></td>
</tr>
<tr>
<td>cable</td>
<td>-</td>
<td>-</td>
<td>3.65</td>
<td></td>
</tr>
</tbody>
</table>

6. VIEWER ENGAGEMENT

We study the extent to which a viewer is engaged with the video content of the content provider. A simple metric that measures engagement is play time. Here we study play time on a per view basis, though one could study play time aggregated over all views of a visit (called visit play time) or play time aggregated over all visits of a viewer (called viewer play time). Figure 14 shows the CDF of play time over our entire data set. A noticeable fact is that a significant number of views have very small play time with the median play time only 35.4 seconds. This is likely caused by “video surfing” where a viewer quickly views a sequence of videos to see what might of interest to him/her, before settling in on the videos that he/she wants to watch. The fact that a viewer watched on average of 22.48 minutes per visit (cf. Figure 4) is consistent with this observation. Play time is clearly impacted by both the interest level of the viewer in the video and the stream quality. Viewer interest could itself be a function of complex factors, for instance, Italian viewers might be more interested in soccer world cup videos than American viewers, even more so if the video is of a game where Italy is playing. In understanding the impact of stream quality on viewer engagement, the challenge is to neutralize the bias from confounding variables not related to stream quality such as viewer interest, geography, and connection type. Since more rebuffer delay is expected of videos with a longer duration, we use normalized rebuffer delay\(^5\) that equals $100 \times \text{(rebuffer delay/video duration)}$.

**Assertion 6.1.** An increase in (normalized) rebuffer delay can cause a decrease in play time.

To evaluate the above assertion, we first classify views by bucketing their normalized rebuffer delay into 1\% buckets. Then, we compute and plot the average play time for all views within each bucket (see Figure 15). The decreasing trend visualizes the negative correlation that exists between normalized rebuffer delay and play time. The Kendall correlation between the two metrics is $-0.421$, quantifying the negative correlation.

6.1 QED Analysis

To examine the causality of Assertion 6.1, we devise a QED where the treatment set $T$ consists of all views that suffered normalized rebuffer delay more than a certain threshold $\gamma\%$. Given a value of $\gamma$ as input, the treated views are matched with untreated views that did not experience rebuffering as follows.

1. **Match step.** We form a set of matched pairs $M$ as follows. Let $T$ be the set of all views who have a normalized rebuffer delay of at least $\gamma\%$. For each view $u$ in $T$, suppose that $u$ reaches the normalized rebuffer delay threshold $\gamma\%$ when viewing the $t^{th}$ second of the

\(^5\)Note that normalized rebuffer delay can go beyond 100\% if we rebuffer for longer than the total duration of the video.
Figure 15: Correlation of normalized rebuffer delay with play time.

video, i.e., view $u$ receives treatment after watching the first $t$ seconds of the video, though more of the video could have been played after that point. We pick a view $v$ uniformly and randomly from the set of all possible views such that

(a) the viewer of $v$ has the same geography, connection type, and is watching the same video as the viewer of $u$.

(b) View $v$ has played at least $t$ seconds of the video without rebuffering till that point.

2. Score step. For each pair $(u, v) \in M$, we compute

$$outcome(u, v) = \frac{\text{play time of } v - \text{play time of } u}{\text{video duration}}.$$  

$$Net\ Outcome = \left( \frac{\sum_{(u, v) \in M} outcome(u, v)}{|M|} \right) \times 100.$$

Note that closer we can make the matched views $u$ and $v$ in variables other than the treatment, the more accurate our QED results. Though as a practical matter, adding too many matching parameters can highly reduce the availability of matches, eventually impacting the statistical significance of the results. It is worth noting step 1(b) above where we ensure that $v$ watches the video to at least the same point as when $u$ first received treatment. Thus, at the time both $u$ and $v$ play the $t^{th}$ second of the video, they have viewed the same content and the only difference between them is one had rebuffering and the other did not. The net outcome of the matching algorithm can be viewed as the difference in the play time of $u$ and $v$ expressed as a percent of the video duration. Figure 16 shows that on average a view that experienced normalized rebuffer delay of 1% or more played 5.02% less of the video. There is a general upward trend in the net outcome when the treatment gets harsher with increasing values of $\gamma$. Much as in Section 5.1, we use the sign test to compute the $p$-values for each QED outcome. All $p$-values were extremely small as shown in Figure 16, making the results statistically significant.

7. REPEAT VIEWERSHIP

We study the viewers who after watching videos on a content provider’s site return after some period of time to watch more. Repeat viewers are valued highly-valued by media content providers as these viewers are more engaged and more loyal to the content provider’s site. Even a small decrease (or, increase) in the return rate of viewers can have a large impact on the business metrics of the content provider. Clearly, a number of factors, including how captivating the video content is to the viewer, influence whether or not a viewer returns. However, we show that stream quality can also influence whether or not a viewer returns.

The most drastic form of quality degradation is failure when a viewer is unable to play a video successfully. Failures can be caused by a number of issues, including problems with the content (broken links, missing video files, etc), the client software (media player bugs, etc), or the infrastructure (network failure, server overload, etc). More frustrating than a failed view is a failed visit where a viewer tries to play videos from the content providers site but fails and leaves the site immediately after the failure, presumably with some level of frustration. (Note that the definition of a failed visit does not preclude successful views earlier in that visit before the last view(s) that failed.) We focus on the impact of a failed visit experienced by a viewer on his/her likelihood of returning to the content provider’s site.

**Assertion 7.1.** A viewer who experienced a failed visit is less likely to return to the content provider’s site to view more videos within a specified time period than a similar viewer who did not experience a failed visit.

To examine if the above assertion holds, we classify each of our views as either failed or normal (i.e., not failed). For each failed visit (resp., normal visit), we compute the return time which is defined to be the next time the viewer returns to the content provider’s site. (Return time could be infinite if they do not return to the site within our trace window.) Figure 17 shows the CDF of the return time for both failed visits and normal visits. It can be seen that there is significant reduction in the probability of return following a failed visit as opposed to a normal one. For instance, the probability of returning within 1 day after a failed visit is 8.0%.

<table>
<thead>
<tr>
<th>Normalized Rebuffer Delay $\gamma$ (percent)</th>
<th>Net Outcome (percent)</th>
<th>$p$-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>5.02</td>
<td>$&lt; 10^{-14}$</td>
</tr>
<tr>
<td>2</td>
<td>5.54</td>
<td>$&lt; 10^{-14}$</td>
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<tr>
<td>3</td>
<td>5.7</td>
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<tr>
<td>4</td>
<td>6.66</td>
<td>$&lt; 10^{-8}$</td>
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<tr>
<td>5</td>
<td>6.27</td>
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<tr>
<td>6</td>
<td>7.38</td>
<td>$&lt; 10^{-4}$</td>
</tr>
<tr>
<td>7</td>
<td>7.48</td>
<td>$&lt; 10^{-3}$</td>
</tr>
</tbody>
</table>

Figure 16: A viewer who experienced more rebuffer delay on average watched less video than an identical viewer who had no rebuffer.
versus 11% after a normal one. Likewise, the probability of returning within 1 week after a failed visit is 25% versus 27% after a normal one.

Figure 17: Probability of the return after a failed visit and after a normal visit. The probability of returning within a specified return time is distinctly smaller after a failed visit than after a normal one.

7.1 QED Analysis

We perform a QED analysis to strengthen Assertion 7.1 by considering viewers\(^6\) with a failed visit to be the treated set \(T\). For each \(u \in T\) we find a matching viewer \(v\) that is similar to \(u\) in all the confounding variables. As before, we ensure that viewers \(u\) and \(v\) are from the same geography, have the same connection type, and are viewing content from the same content provider. However, there is a subtle characteristic that need to be matched. Specifically, we need also ensure that the propensity of \(u\) to watch videos prior to when \(u\) received treatment is equivalent to the corresponding propensity of \(v\). This ensures that any differential behavior after the treatment can be attributed to the treatment itself.

To reinforce the last point, a viewer who watches more video at a site is more likely to have had a failed view. Therefore, the treated set \(T\) of viewers has a bias towards containing more frequent visitors to site who also watch more video. Figure 18 shows the CDF of the aggregate play time of a viewer across all visits. It can be seen that the treated set \(T\) has viewers who have watched for more time in aggregate.

To neutralize this effect, we match on the number of prior visits and aggregate play time in step 1 (c) below and make them near identical, so that we are comparing two viewers who have exhibited similar propensity to visit the site prior to treatment. The use of a similarity metric of this kind for matching is common in QED analysis and is similar in spirit to propensity score matching of [22].

\(^6\)Note that in this matching we are matching viewers and not views as we are evaluating the repeat viewership of a viewer over time.

The matching algorithm follows.

1. **Match step.** We produce a matched set of pairs \(M\) as follows. Let \(T\) be the set of all viewers who have had a failed visit. For each \(u \in T\) we pick the first failed visit of viewer \(u\). We then pair \(u\) with a viewer \(v\) picked uniformly and randomly from the set of all possible viewers such that

   (a) Viewer \(v\) has the same geography, same connection type as \(u\), and is watching the content from the same content provider as \(u\).

   (b) Viewer \(v\) had a normal visit at about the same time (within ±3 hours) as the first failed visit of viewer \(u\). We call the failed visit of \(u\) and the corresponding normal visit of \(v\) that occurred at a similar time as matched visits.

   (c) Viewer \(u\) and \(v\) have the same number of visits and about the same total viewing time (±10 minutes) prior to their matched visits.

2. **Score step.** For each pair \((u, v) \in M\) and each return time \(\delta\). We assign \(outcome(u, v, \delta)\) to \(-1\) if \(u\) returns within the return time and \(v\) does not, \(+1\) if \(v\) returns within the return time and \(u\) does not, and \(0\) otherwise.

   \[
   Net\ Outcome(\delta) = \left( \frac{\sum_{(u, v) \in M} outcome(u, v, \delta)}{|M|} \right) \times 100.
   \]

Figure 19 shows the outcome of the matching algorithm for various values of the return time \(\delta\). The positive values of the outcome provide strong evidence of the causality of Assertion 7.1 since it shows that viewers who experienced a normal visit returned more than their identical pair with a failed visit. To take a numerical example, for \(\delta = 1\) day, 458,621 pairs were created. The pairs where the normal viewer returned but its identical failed pair did not exceed
<table>
<thead>
<tr>
<th>Return Time $\delta$ (in days)</th>
<th>Outcome (percent)</th>
<th>P-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>2.38</td>
<td>$&lt;10^{-6}$</td>
</tr>
<tr>
<td>2</td>
<td>2.51</td>
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<td>6</td>
<td>1.90</td>
<td>$&lt;10^{-14}$</td>
</tr>
<tr>
<td>7</td>
<td>2.32</td>
<td>$&lt;10^{-6}$</td>
</tr>
</tbody>
</table>

Figure 19: A viewer who experienced a failed visit is less likely to return within a time period than a viewer who experienced a normal visit.

the pairs where the opposite happened. The amount of pairs in excess was 10,909 pairs, which is 2.38% of the total pairs. Using the sign test, we show that the p-value is extremely small ($2.2 \times 10^{-58}$), providing strong evidence of statistical significance for the outcome. Note that as $\delta$ increases, the outcome score remained in a similar range. However, one would expect that for very large $\delta$ values the effect of the failed event should wear off, but we did not analyze traces that were long enough to evaluate if such a phenomenon occurs. All p-values remain significantly smaller than our threshold of significance of 0.001, allowing us to conclude that the results are statistically significant.

8. RELATED WORK

The quality metrics considered here have more than a dozen years of history within industry where early measurement systems used synthetic “measurement agents” deployed around the world to measure metrics such as failures, startup delay, rebuffing, and bitrate, example, Akamai’s Stream Analyzer measurement system [1, 24]. There have been early studies at Akamai on streaming quality metrics using these tools [19]. However, truly large-scale studies were made possible only with the recent advent of client-side measurement technology that could measure and report detailed quality and behavioral data from actual viewers. To our knowledge, the first important large-scale study and closest in spirit to our work is the study of viewer engagement published last year [9] that shows several correlational relationships between quality (such as rebuffing), content type (such as live, short/long VoD), and viewer engagement (such as play time). A recent sequel to the above work [15] studies the use of quality metrics to enhance video delivery. A key differentiation of our work from prior work is our focus on establishing causal relationships, going a step beyond just correlation. While our viewer engagement analysis was also correlationally established in [9], our work takes the next step in ascertaining the causal impact of rebuffing on play time. Besides our results on viewer engagement, we also establish key assertions pertaining to viewer abandonment and repeat viewership that are the first quantitative results of its kind. However, it must be noted that [9] studies a larger set of quality metrics, including join time, average bitrate, and rendering quality, and a larger class of videos including live streaming, albeit without establishing causality.

The work on quasi-experimental design in the social and medical sciences has a long and distinguished history stretching several decades that is well documented in [23]. Though its application to data mining is more recent. In [18], the authors use QEDs to answer questions about user behavior in social media such as Stack Overflow and Yahoo Answers.

There are a number of other studies on perceived quality though they tend to be small-scale studies or do not link the quality to user behavior [10, 7]. There has also been prior work for other types of systems. For instance, the relationship between page download times and user satisfaction [3] for the web and quantifying user satisfaction for Skype [6]. There has also been work on correlating QoS with QoE (quality of experience) for multimedia systems using human subjects [27]. These of course have a very different focus from our work and do not show causal impact. There has been significant amount of work in workload characterization of streaming media, P2P, and web workloads [25, 4]. Even though we do characterize the workload to a degree, our focus is quality and viewer behavior.

9. CONCLUSIONS

Our work is the first to demonstrate a causal nexus between stream quality and viewer behavior. The results presented in our work are important because they are the first quantitative demonstration that key quality metrics causally impact viewer behavioral metrics that are key to both content providers and CDN operators. As all forms of media migrate to the Internet, both video monetization and the design of CDNs will increasingly demand a true causal understanding of this nexus. Establishing a causal relationship by systematically eliminating the confounding variables is immensely important, as mere correlational studies have the potential costly risk of making incorrect conclusions.

Our work breaks new ground in understanding viewer abandonment and repeat viewership. Further, it sheds more light on the known correlational impact of quality on viewer engagement by establishing its causal impact. Our work on startup delay show that more delay causes more abandonment, for instance, a 1 second increase in delay increases the abandonment rate by 5.8%. We also showed the strong impact of rebuffing on the video play time. For instance, we showed that a viewer experiencing a rebuffer delay that equals or exceeds 1% of the video duration played 5.02% less of the video in comparison with a similar viewer who experienced no rebuffing. Finally, we examined the impact of failed visits and showed that a viewer who experienced failures is less likely to return to the content provider’s site in comparison to a similar viewer who did not experience failures. In particular, we showed that a failed visit decreased the likelihood of a viewer returning within a week by 2.32%. While reviewing these results, it is important to remember that small changes in viewer behavior can lead to large changes in monetization, since the impact of a few percentage points over tens of millions of viewers can accrue to large impact over a period of time.

As more and more data become available, we expect that our QED tools will play an increasing larger role in establishing key causal relationships that are key drivers of both the content provider’s monetization framework and the CDN’s next-generation delivery architecture. The increasing scale of the measured data greatly enhances the statistical significance of the derived conclusions and the efficacy of our tools. Further, we expect that our work provides an important tool for establishing causal relationships in other areas.
of measurement research in networked systems that have so far been limited to correlational studies.

10. ACKNOWLEDGEMENTS

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11. REFERENCES


