COMPSCI 614: Randomized Algorithms with Applications to Data Science

Prof. Cameron Musco

University of Massachusetts Amherst. Spring 2024. Lecture 19 • I will send responses to project progress reports soon.

Last Week: Start on Markov Chains.

- Start on Markov chains and their analysis
- Markov chain based algorithms for satisfiability: $\approx n^2$ time for 2-SAT, and $\approx (4/3)^n$ for 3-SAT.

Today: Markov Chains Continued

- The gambler's ruin problem.
- Aperiodicity and stationary distribution of a Markov chain.
- The fundamental theorem of Markov chains.

Markov Chain Review

 A discrete time stochastic process is a Markov chain if is it memoryless:

$$\Pr(X_t = a_t | X_{t-1} = a_{t-1}, \dots, X_0 = a_0) = \Pr(X_t = a_t | X_{t-1} = a_{t-1})$$

• If each X_t can take *m* possible values, the Markov chain is specified by the transition matrix $P \in [0, 1]^{m \times m}$ with

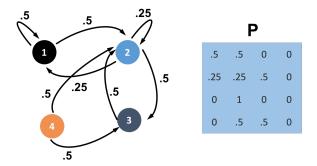
$$P_{i,j} = \Pr(\mathbf{X}_{t+1} = j | \mathbf{X}_t = i).$$

• Let $q_t \in [0,1]^{1 \times m}$ be the distribution of X_i . Then $q_{t+1} = q_t P$.

q ₁				Р				q ₂				
.5	.5	0	0	.5	.5	0	0	=	.375	.375	.25	0
				.25	.25	.5	0	_				
				0	1	0	0					
				0	.5	.5	0					

Markov Chain Review

Often viewed as an underlying state transition graph. Nodes correspond to possible values that each \mathbf{X}_t can take.



The Markov chain is irreducible if the underlying graph consists of single strongly connected component.

Gambler's Ruin

Gambler's Ruin



- You and 'a friend' repeatedly toss a fair coin. If it hits heads, you give your friend \$1. If it hits tails, they give you \$1.
- You start with ℓ_1 and your friend starts with ℓ_2 . When either of you runs out of money the game terminates.
- What is the probability that you win ℓ_2 ?

Gambler's Ruin Markov Chain

Let $X_0, X_1, ...$ be the Markov chain where X_t is your profit at step t. $X_0 = 0$ and: $P_{-\ell_1, -\ell_1} = P_{\ell_2, \ell_2} = 1$ $P_{i,i+1} = P_{i,i-1} = 1/2$ for $-\ell_1 < i < \ell_2$ 1/21/21/21/21/21/21/21/2

• ℓ_1 and ℓ_2 are absorbing states.

• All *i* with $-\ell_1 < i < \ell_2$ are transient states. I.e., $\Pr[\mathbf{X}_{t'} = i \text{ for some } t' > t | \mathbf{X}_t = i] < 1.$

Observe that this Markov chain is also a Martingale since $\mathbb{E}[X_{t+1}|X_t] = X_t.$

Let X_0, X_1, \ldots be the Markov chain where X_t is your profit at step t. $X_0 = 0$ and:

$$\begin{aligned} P_{-\ell_1,-\ell_1} &= P_{\ell_2,\ell_2} = 1 \\ P_{i,i+1} &= P_{i,i-1} = 1/2 \text{ for } -\ell_1 < i < \ell_2 \end{aligned}$$

We want to compute $q = \lim_{t \to \infty} \Pr[X_t = \ell_2]$.

By linearity of expectation, for any *i*, $\mathbb{E}[X_i] = 0$. Further, for $q = \lim_{t \to \infty} \Pr[X_t = \ell_2]$, since $-\ell_1, \ell_2$ are the only non-transient states,

$$\lim_{t\to\infty}\mathbb{E}[X_t]=\ell_2q+-\ell_1(1-q)=0.$$

Solving for q, we have $q = \frac{\ell_1}{\ell_1 + \ell_2}$.

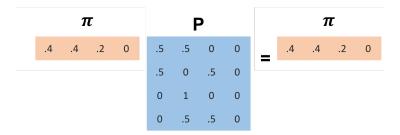
What if you always walk away as soon as you win just \$1. Then what is your probability of winning, and what are your expected winnings?

Stationary Distributions

Stationary Distribution

A stationary distribution of a Markov chain with transition matrix $P \in [0, 1]^{m \times m}$ is a distribution $\pi \in [0, 1]^m$ such that $\pi = \pi P$.

I.e. if $X_t \sim \pi$, then $X_{t+1} \sim \pi P = \pi$.



Think-pair-share: Do all Markov chains have a stationary distribution?

Claim (Existence of Stationary Distribution)

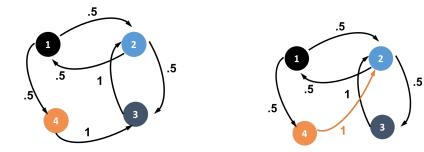
Any Markov chain with a finite state space, and transition matrix $P \in [0, 1]^{m \times m}$ has a stationary distribution $\pi \in [0, 1]^m$ with $\pi = \pi P$.

Follows from the Brouwer fixed point theorem: for any continuous function $f : S \to S$, where S is a compact convex set, there is some x such that f(x) = x.

Periodicity

The periodicity of a state *i* is defined as:

$$T = \gcd\{t > 0 : \Pr(X_t = i \mid X_0 = i) > 0\}.$$



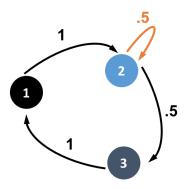
The state is aperiodic if it has periodicity T = 1.

A Markov chain is aperiodic if all states are aperiodic.

Periodicity

Claim

If a Markov chain is irreducible, and has at least one self-loop, then it is aperiodic.



Theorem (The Fundamental Theorem of Markov Chains)

Let X_0, X_1, \ldots be a Markov chain with a finite state space and transition matrix $P \in [0, 1]^{m \times m}$. If the chain is both irreducible and aperiodic,

- 1. There exists a unique stationary distribution $\pi \in [0, 1]^m$ with $\pi = \pi P$.
- 2. For any states *i*, *j*, $\lim_{t\to\infty} \Pr[\mathbf{X}_t = i | X_0 = j] = \pi(i)$. *i.e., for any initial distribution* q_0 , $\lim_{t\to\infty} q_t = \lim_{t\to\infty} q_0 P^t = \pi$.
- 3. $\pi(i) = \frac{1}{\mathbb{E}[\min(t:X_t=i)|X_0=i]}$. *l.e.*, $\pi(i)$ is the inverse of the average expected return time from state i back to i.

In the limit, the probability of being at any state *i* is independent of the starting state.

Stationary Distribution Example 1

Shuffling Markov Chain: Given a pack of *c* cards. At each step draw two random cards, swap them and repeat.

- What is the state space of this chain?
- What is the transition probability $P_{i,j}$? How does it compare to $P_{j,i}$?
- This Markov chain is symmetric and thus its stationary distribution is uniform, $\pi(i) = \frac{1}{c!}$.

Letting m = c! denote the size of the state space,

$$\pi P_{:,i} = \sum_{j} \pi(j) P_{j,i} = \sum_{j} \pi(j) P_{i,j} = \frac{1}{m} \sum_{j} P_{i,j} = \frac{1}{m} = \pi(i).$$

Once we have exhibited a stationary distribution, we know that it is unique and that the chain converges to it in the limit! **Random Walk on an Undirected Graph:** Consider a random walk on an undirected graph. If it is at node *i* at step *t*, then it moves to any of *i*'s neighbors at step t + 1 with probability $\frac{1}{d_i}$.

- What is the state space of this chain?
- What is the transition probability $P_{i,j}$?
- Is this chain aperiodic?
- If the graph is not bipartite, then there is at least one odd cycle, making the chain aperiodic.



Random Walk on an Undirected Graph: Consider a random walk on an undirected graph. If it is at node *i* at step *t*, then it moves to any of *i*'s neighbors at step t + 1 with probability $\frac{1}{d_i}$.

Claim: When the graph is not bipartite, the unique stationary distribution of this Markov chain is given by $\pi(i) = \frac{d_i}{2|E|}$.

$$\pi P_{:,i} = \sum_{j} \pi(j) P_{j,i} = \sum_{j} \frac{d_j}{2|E|} \cdot \frac{1}{d_j} = \sum_{j} \frac{1}{2|E|} = \frac{d_i}{2|E|} = \pi(i).$$

I.e., the probability of being at a given node *i* is dependent only on the node's degree, not on the structure of the graph in any other way.