# COMPSCI 614: Randomized Algorithms with Applications to Data Science

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University of Massachusetts Amherst. Spring 2024. Lecture 18

- Problem Set 4 is due Monday 4/22 at 11:59pm.
- No quiz this week.

#### Summary

#### Last Class:

- Leverage score intuition.
- Connection to spectral graph sparsification
- Connection to effective resistances in electrical networks. Note: I am not going to finish this full derivation – see Lecture 17 slides if you are interested.

#### Today:

- New unit: Markov Chains.
- Markov chain based algorithms for 2-SAT and 3-SAT.

- A discrete time stochastic process is a collection of random variables  $X_0, X_1, X_2, \ldots,$
- A discrete time stochastic process is a Markov chain if is it memoryless:

$$Pr(X_t = a_t | X_{t-1} = a_{t-1}, \dots, X_0 = a_0) = Pr(X_t = a_t | X_{t-1} = a_{t-1})$$
$$= P_{a_{t-1}, a_t}.$$

**Question:** In a Markov chain, is  $X_t$  independent of  $X_{t-2}, X_{t-3}, \ldots, X_0$ ?

## **Transition Matrix**

A Markov chain  $X_0, X_1, ...$  where each  $X_i$  can take m possible values, is specified by the transition matrix  $P \in [0, 1]^{m \times m}$  with

$$P_{j,k} = \Pr(\mathbf{X}_{i+1} = k | \mathbf{X}_i = j).$$

Let  $q_i \in [0, 1]^{1 \times m}$  be the distribution of  $X_i$ . Then  $q_{i+1} = q_i P$ .



# **Graph View**

Often viewed as an underlying state transition graph. Nodes correspond to possible values that each **X**<sub>i</sub> can take.



The Markov chain is **irreducible** if the underlying graph consists of single strongly connected component.

**Motivating Example:** Find a satisfying assignment for a 2-CNF formula with *n* variables.

 $(x_1 \vee \bar{x}_2) \wedge (\bar{x}_1 \vee \bar{x}_3) \wedge (x_1 \vee x_2) \wedge (x_4 \vee \bar{x}_3) \wedge (x_4 \vee \bar{x}_1)$ 

A simple 'local search' algorithm:

- 1. Start with an arbitrary assignment.
- 2. Repeat 2*mn*<sup>2</sup> times, terminating if a satisfying assignment is found:
  - Chose an arbitrary unsatisfied clause.
  - Pick one of the variables in the clause uniformly at random, and switch the assignment of the variable.
- 3. If a valid assignment is not found, return that the formula is unsatisfiable.

**Claim:** If the formula is satisfiable, the algorithm finds a satisfying assignment with probability  $\ge 1 - 2^{-m}$ .

Fix a satisfying assignment *S*. Let  $X_i \le n$  be the number of variables that are assigned the same values as in *S*, at step *i*.



- $X_{i+1} = X_i \pm 1$  since we flip one variable in an unsatisfied clause.
- ·  $Pr(X_{i+1} = X_i + 1) \ge$
- $\Pr(X_{i+1} = X_i 1) \leq$

 $(x_1 \vee \bar{x}_2) \wedge (\bar{x}_1 \vee \bar{x}_3) \wedge (x_1 \vee x_2) \wedge (x_4 \vee \bar{x}_3) \wedge (x_4 \vee \bar{x}_1)$ 

# Coupling to a Markov Chain

The number of correctly assigned variables at step *i*, X<sub>i</sub>, obeys  $Pr(X_{i+1} = X_i + 1) \ge \frac{1}{2}$  and  $Pr(X_{i+1} = X_i - 1) \le \frac{1}{2}$ . Is X<sub>0</sub>, X<sub>1</sub>, X<sub>2</sub>,... a Markov chain?

Define a Markov chain  $Y_0, Y_1, \ldots$  such that  $Y_0 = X_0$  and:

$$Pr(\mathbf{Y}_{i+1} = 1 | \mathbf{Y}_i = 0) = 1$$
  

$$Pr(\mathbf{Y}_{i+1} = j + 1 | \mathbf{Y}_i = j) = 1/2 \text{ for } 1 \le j \le n - 1$$
  

$$Pr(\mathbf{Y}_{i+1} = j - 1 | \mathbf{Y}_i = j) = 1/2 \text{ for } 1 \le j \le n - 1$$
  

$$Pr(\mathbf{Y}_{i+1} = n | \mathbf{Y}_i = n) = 1.$$

- Our algorithm terminates as soon as  $X_i = n$ . We expect to reach this point only more slowly with  $Y_i$ . So it suffices to argue that  $Y_i = n$  with high probability for large enough *i*.
- Formally could use a coupling argument (will see later on).

## Simple Markov Chain Analysis

Want to bound the expected time required to have  $Y_i = n$ .



Let  $h_j$  be the expected number of steps to reach n when starting at node j (i.e., the expected termination time when j variables are assigned correctly.)

$$h_n = 0$$
  

$$h_0 = h_1 + 1$$
  

$$h_j = \frac{h_{j-1}}{2} + \frac{h_{j+1}}{2} + 1 \text{ for } 1 \le j \le n - 2$$

## Simple Markov Chain Analysis

**Claim:**  $h_j = h_{j+1} + 2j + 1$ . Can prove via induction on *j*.

•  $h_0 = h_1 + 1$ , satisfying the claim in the base case.

$$h_{j} = \frac{h_{j-1}}{2} + \frac{h_{j+1}}{2} + 1$$
  

$$h_{j} = \frac{h_{j}}{2} + (j-1) + \frac{1}{2} + \frac{h_{j+1}}{2} + 1$$
  

$$h_{j} = \frac{h_{j}}{2} + \frac{h_{j+1}}{2} + j + \frac{1}{2}.$$

• Rearranging gives:  $h_j = h_{j+1} + 2j + 1$ .

So in total we have:

$$h_0 = h_1 + 1 = h_2 + 3 + 1 = \ldots = \sum_{j=0}^{n-1} (2j+1) = n^2.$$

**Upshot:** Consider the Markov chain  $Y_0, Y_1, ..., and let i^*$  be the minimum *i* such  $Y_{i^*} = n$ . Then  $\mathbb{E}[i^*] \le n^2$ .

- Thus, by Markov's inequality, with probability  $\geq 1/2$ , our 2-SAT algorithms finds a satisfying assignment within  $2n^2$  steps.
- Splitting our  $2mn^2$  total steps into *m* periods of  $2n^2$  steps each, we fail to find a satisfying assignment in all *m* periods with probability at most  $1/2^m$ .

#### 3**-SAT**

**More Challenging Problem:** Find a satisfying assignment for a 3-CNF formula with *n* variables.

 $(x_1 \lor \overline{x}_2 \lor \overline{x}_3) \land (\overline{x}_1 \lor \overline{x}_3 \lor x_4) \land (x_1 \lor x_2 \lor \overline{x}_3).$ 

- 3-SAT is famously NP-hard. What is the naive deterministic runtime required to solve 3-SAT?
- The current best known runtime is *O*(1.307<sup>*n*</sup>) [Hansen, Kaplan, Zamir, Zwick, 2019].
- Will see that our simple Markov chain approach gives an  $O(1.3334^n)$  time algorithm.
- Note that the exponential time hypothesis conjectures that  $O(c^n)$  is needed to solve 3-SAT for some constant c > 1. The strong exponential time hypothesis conjectures that for  $k \to \infty$ , solving *k*-SAT requires  $O(2^n)$  time.

- 1. Start with an arbitrary assignment.
- 2. Repeat *m* times, terminating if a satisfying assignment is found:
  - Chose an arbitrary unsatisfied clause.
  - Pick one of the variables in the clause uniformly at random, and switch the assignment of the variable.
- 3. If a valid assignment is not found, return that the formula is unsatisfiable.

## Randomized 3-SAT Analysis

As in the 2-SAT setting, let **X**<sub>*i*</sub> be the number of correctly assigned variables at step *i*. We have:

$$Pr(X_i = X_{i-1} + 1) \ge$$
$$Pr(X_i = X_{i-1} - 1) \le$$

Define the coupled Markov chain  $Y_0, Y_1, ...$  as before, but with  $Y_i = Y_{i-1} + 1$  with probability 1/3 and  $Y_i = Y_{i-1} - 1 = 2/3$ .



How many steps do you expect are needed to reach  $Y_i = n$ ?

Letting  $h_j$  be the expected number of steps to reach n when starting at node j,

$$h_n = 0$$
  

$$h_0 = h_1 + 1$$
  

$$h_j = \frac{2h_{j-1}}{3} + \frac{h_{j+1}}{3} + 1 \text{ for } 1 \le j \le n - 1$$

- We can prove via induction that  $h_j = h_{j+1} + 2^{j+2} 3$  and in turn,  $h_0 = 2^{n+2} 4 3n$ .
- Thus, in expectation, our algorithm takes at most  $\approx 2^{n+2}$  steps to find a satisfying assignment if there is one.
- Is this an interesting result?

# Modified 3-SAT Algorithm

**Key Idea:** If we pick our initial assignment uniformly at random, we will have  $\mathbb{E}[X_0] = n/2$ . With very small, but still non-negligible probability,  $X_0$  will be much larger, and our random walk will be more likely to find a satisfying assignment.

#### Modified Randomized 3-SAT Algorithm:

Repeat *m* times, terminating if a satisfying assignment is found:

- 1. Pick a uniform random assignment for the variables.
- 2. Repeat 3*n* times, terminating if a satisfying assignment is found:
  - Chose an arbitrary unsatisfied clause.
  - Pick one of the variables in the clause uniformly at random, and switch the assignment of the variable.

If a valid assignment is not found, return that the formula is unsatisfiable.

## Modified 3-SAT Analysis

Consider a single random assignment with  $X_0 = n - j$ . I.e., we need to correct *j* variables to find a satisfying assignment.

Let  $q_j$  be a lower bound on the success probability in this case. Since  $j \le n$  and since we run the search process for 3n steps,

$$q_j = \Pr[\mathbf{X}_{3n} = n]$$

$$\geq \Pr[X_{3j} = n]$$

 $\geq$  Pr[take exactly 2*j* steps forward and *j* steps back in 3*j* steps]

$$= \binom{3j}{j} \left(\frac{2}{3}\right)^j \cdot \left(\frac{1}{3}\right)^{2j}.$$

Via Stirling's approximation,  $\binom{3j}{j} \ge \frac{1}{\sqrt{j}} \cdot \frac{3^{3j-2}}{2^{2j-2}}$ , giving:

$$q_j \ge \frac{2^2}{3^2\sqrt{j}} \cdot \frac{3^{3j}}{2^{2j}} \cdot \frac{2^j}{3^{3j}} \approx \frac{1}{\sqrt{j} \cdot 2^j} \ge \frac{1}{\sqrt{n} \cdot 2^j}.$$

## Modified 3-SAT Analysis

Our overall probability of success in a single trial is then lower bounded by:

$$q \ge \sum_{j=0}^{n} \Pr[\mathbf{X}_{0} = n - j] \cdot q_{j}$$
$$\ge \sum_{j=0}^{n} {n \choose j} \cdot \frac{1}{2^{n}} \cdot \frac{1}{\sqrt{n} \cdot 2^{j}}$$
$$\ge \frac{1}{\sqrt{n} \cdot 2^{n}} \sum_{j=0}^{n} {n \choose j} \cdot \frac{1}{2^{j}}$$
$$= \frac{1}{\sqrt{n} \cdot 2^{n}} \cdot \left(\frac{3}{2}\right)^{n} = \frac{1}{\sqrt{n}} \cdot \left(\frac{3}{4}\right)^{n}.$$

Thus, if we repeat for  $m = O\left(\sqrt{n} \cdot \left(\frac{4}{3}\right)^n\right) = O(1.33334^n)$  trials, with very high probability, we will find a satisfying assignment if there is one.