

EM in two pages

Brendan O'Connor

January 31, 2017

We would like to maximize *incomplete data loglikelihood* for seen data x , latent variables z , and model parameters θ . (Latent variables are local to an instance, but parameters cut across the dataset.) Unfortunately the sum inside the log is hard to deal with.

$$\ell(\theta) = \log p(x|\theta) = \log \sum_z p(x, z|\theta) = \sum_i \log \sum_{z_i} p(x_i, z_i|\theta)$$

Assume we have a model where, if only we knew the z 's, it would be easy. That is, assume we have a good algorithm to maximize the *complete* likelihood $\max_{\theta} \log p(x, z|\theta)$, when z is known and fixed. This motivates why we want to derive EM in the first place.

EM derivation: add in a $q_i(z)/q_i(z)$ term (weird special local probabilities for the latent variables, for every instance) and apply Jensen's inequality to get the EM bound.

$$\ell(\theta) = \sum_i \log \sum_{z_i} \frac{q_i(z_i)}{q_i(z_i)} p(x_i, z_i|\theta) \tag{1}$$

$$\geq \sum_i \sum_{z_i} q_i(z_i) [\log p(x_i, z_i|\theta) - \log q_i(z_i)] \tag{2}$$

$$\equiv J(q, \theta) \quad (\equiv \sum_i J_i) \tag{3}$$

J consists of weighted complete-loglikelihood, plus the entropy of q . EM is coordinate ascent on J . Maximizing for θ (the M-step) is simply maximizing weighted log-likelihood, since the q entropy term drops out. (For counting-based estimation, this is simply weighted counting. For gradient-based estimation, this is simply weighted gradient calculation.) Maximizing for q (the E-step) leads to setting q to local posteriors. This is because, since $p(x, z) = p(z|x)p(x)$, rewrite as

$$J_i = \sum_z q(z) \log[p(z|x)p(x)/q(z)] = -D(q(z) || p(z|x)) + \log p(x)$$

where the first term is the negative KL divergence between $q(z)$ and $p(z|x)$; to maximize J_i , the $p(x)$ term is irrelevant to q , so set $q(z) := p(z|x)$ to minimize the KL divergence to zero. Thus the EM steps are

- E-step: set all $q_i(z_i) := p(z_i|x_i, \theta)$
- M-step: learn new $\theta := \arg \max_{\theta} \sum_i \sum_{z_i} q_i(z_i) \log p(x_i, z_i|\theta)$

Or as one big equation,

$$\theta^{(new)} := \arg \max_{\theta} \sum_i \sum_{z_i} p(z_i|x_i, \theta^{(old)}) \log p(x_i, z_i|\theta)$$

In mixture models, $p(x, z|\theta) = p(z|\theta)p(x|z, \theta)$. The $q_i(z)$ local posteriors therefore are different than the mixture priors $p(z|\theta)$. When you update the prior parameters (mixture proportions), that's the average across all instances—at any local instance your belief is much different. Also in the E-step you'll want to use Bayes rule, $p(z|x) \propto p(z)p(x|z)$.

There are two ways to break down the J objective:

$$\begin{aligned}
 \ell(\theta) &\equiv \log p_\theta(X) = \sum_i \log p_\theta(x_i) = \sum_i \log \sum_{z_i} p_\theta(x_i, z_i) \\
 &= \sum_i \log \sum_{z_i} \frac{q_i(z_i)}{q_i(z_i)} p_\theta(x_i, z_i) \\
 \text{Jensen's Ineq} \quad \ell(\theta) &\geq \sum_i \sum_{z_i} q_i(z_i) \log \left[\frac{1}{q_i(z_i)} p_\theta(z_i|x_i) p_\theta(x_i) \right] \\
 J(Q, \theta) &\equiv \underbrace{-E_Q \log Q(Z)}_{H(Q)} + \underbrace{E_Q \log p_\theta(Z|X)}_{E_Q \log p_\theta(X, Z)} + \log p_\theta(X) \\
 &\qquad \qquad \qquad \text{Entropy} \qquad \qquad \qquad \text{Weighted LL} \qquad \qquad \text{M-step} \\
 &\qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \max_\theta J(Q, \theta) \\
 &\qquad \qquad \qquad \underbrace{-KL(Q \parallel p_\theta(Z|X))}_{\text{Nonpositive}} \qquad \qquad \underbrace{\ell(\theta)}_{\text{E-step}} \qquad \qquad \max_Q J(Q, \theta) \\
 &\qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \qquad \Rightarrow Q := P_\theta(Z|X)
 \end{aligned}$$

The E-step makes the bound tight for the current θ since it achieves $KL = 0$. Thus after the E-step, a Q has been chosen such that

$$J(Q, \theta^{(cur)}) = \ell(\theta^{(cur)})$$

Thus every iteration of EM results in a θ with a higher log-likelihood.

See also Neal and Hinton 1998, Murphy 2012 chs 11 and 21, MacKay 2003, and <http://cs229.stanford.edu/notes/cs229-notes8.pdf>.